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Education and Health Impacts of an Affirmative Action Policy on Minorities in India

by

Robin Dhakal

A dissertation submitted in partial fulfillment Of the requirements for the degree of Doctor of Philosophy Department of Economics College of Arts and Sciences University of South Florida

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Keywords: quota, caste, gross enrollment ratio, dropout rate, infant mortality, under-five mortality

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ABSTRACT

Article 332 of the Constitution of India (1950) stipulates that certain electoral districts in each state should be reserved for minority groups, namely the "Scheduled Caste" (SC) and the Scheduled Tribe" (ST), through the reservation of seats in the states' legislative assemblies. Even though the original article stated that the reservation policy would be in place for just twenty years, it has been amended several times and is still in effect. This dissertation examines the impact of the policy on the education and health outcomes of the SC population. Variations in seat quotas are generated by the timing of elections in different states and the states' fluctuating SC populations. The first paper on education uses data from 25 Indian States and 3 Union Territories for the years 1990-2011 to form a panel dataset to estimate the impact of the quota system on both enrollment and dropout rates among SC students in all levels of schooling. I use the fixed effect regression to test the mechanisms through which an elected SC legislator could have an influence on the education outcomes for the SC population in the represented state. I then use the resulting variables as my controls to identify the causal relationship using the dynamic panel data model. I find that a SC legislator has the potential to influence the number of schools built, as well as the amount of education and welfare expenditure allocated to the SC population. Moreover, I find that the SC political reservation has a positive and statistically significant impact on the SC enrollment rates and a negative and significant impact on the dropout rates, in all levels of schooling. Likewise, I use the NFHS-3 dataset and the Cox

Proportional Hazard Model to estimate the hazard rates (risks of dying) of children under the age of 12 months (IMR) and under the age of 60 months (U5MR) as influenced by different SC quota share quintiles. I find that the 50-60% quota-share quintile has the biggest impact in reducing the IMR and U5MR among the SC children.

CHAPTER ONE: INTRODUCTION

1.1 BACKGROUND

Throughout the world, various forms of social segregation create different sects of people who live as minority groups within national boundaries. A minority population is an often socially and economically disadvantaged and discriminated group of people. Discrimination is not only a social and moral issue, but also an economic issue. According to a report¹ by the Center for American Progress in 2012, workplace discrimination causes a loss of \$64 billion every year as a result of reshuffling 2 million American workers due to some form of discrimination. Similarly, a report² published by The Atlantic in 2013 concludes that gender discrimination may have reduced India's annual growth rate by almost 4% over the past 10 years. Likewise, according to a report³ published by the World Bank in 2014, homophobia and discrimination cost the Indian economy \$30.8 billion every year. The resulting concerns of economists and others have led to the worldwide development and widespread use of "Affirmative Action" (AA) policies designed to prevent societal discrimination of historically disadvantaged groups. Economists have suggested various ways to ensure that those

¹ Burns, C. (2012, March). The Costly Business of Discrimination. Center for American Progress. Retrieved from https://www.americanprogress.org/wp-content/uploads/issues/2012/03/pdf/lgbt_biz_discrimination.pdf ² Jaishankar, D. (2013, March). The Huge Cost of India's Discrimination Against Women. The Atlantic. Retrieved from http://www.theatlantic.com/international/archive/2013/03/the-huge-cost-of-indias-discrimination-against-women/274115/

³ Badgett, M.V. L. (2014, February). The Economic Cost of Homophobia & the Exclusion of LGBT People: A Case Study of India. The World Bank. Retrieved from

https://www.worldbank.org/content/dam/Worldbank/document/SAR/economic-costs homophobia-lgbt-exlusion-india.pdf.

minorities' rights are protected, not only so that they can become equal and productive members of societies, but also to ensure that resource allocation in the society is more efficient. Affirmative action (AA) policies are one of those policies which have been widely used around the world. Such policies are designed to ensure that the historically disadvantaged groups of people do not suffer any discrimination in their societies by enacting policies which favor those who tend to suffer from discrimination. Such policies are often referred to as "positive discrimination" policies and have been enacted in India to protect some of the social, ethnic and religious minorities. In this dissertation, I analyze the impact of a particular AA policy, political reservation, on the education, and health outcomes of a targeted minority group, the Scheduled Caste (SC), in India.

We can trace the idea of AA in the United States to as early as 1865. However, the term "affirmative action" was used for the first time in 1961 when Executive Order 10925 was signed by President John F. Kennedy. The executive order was issued to promote actions that discourage discrimination by ensuring that government contractors "take affirmative action to ensure that applicants are employed, and employees are treated fairly during employment, without regard to their race, creed, color, or national origin." Further executive actions were taken in 1965, by President Lyndon B. Johnson to discourage discrimination in the hiring process. Executive Order 11246 required federal and state employers to "take affirmative action to hire without regard to race, religion and national origin" (gender was added to the list in 1967).

Affirmative action policies in the U.S. have been arguably successful. AA in the U.S. is mostly applicable to employment and in college admission process. The Supreme Court has been involved a few times in rulings pertaining to affirmative action policies. One of the first cases was Regents v Bakke (1978) when the Court upheld the AA policy of using race as one of the factors for college admission. However, the Court also ruled that defining specific quotas is illegal. Similarly, Grutter v Bollinger (2003) Supreme Court ruling also favored the use of race in college admission process. Despite some of the successful court cases, affirmative action in the U.S. remains a contested issue with the Fisher v. Texas (2013) case being the latest related case discussed in the United States Supreme Court. The general successes of AA in the Court have resulted in better outcomes for minorities. According to a report published in 2011 by the Americans for a Fair Chance, there has been an increase in college enrollments of people of color by 57.2%. Similarly, the proportion of women earning bachelor's degree has also been steadily increasing. Likewise, according to statistics from the National Center on Education Statistics, 65% of African American high school graduates immediately enrolled in college in 2011 compared to just 56% in 2007- that number went from 61% to 63% for Hispanic graduates. These improvements in enrollments are attributed to the AA policies.

A different form of AA policy had been enacted in India in 1950, 11 years before the U.S implemented AA to address the rights of African American communities. Even before the 1955 civil rights movement in the United States, India had already implemented a reservation system (a form of quota-based affirmative action) in its 1950 Constitution, to address the issue of discrimination against its minority populations. Minorities in India mostly originate from the social stratification produced by the caste system, which is a key component of Hinduism. Since over 80% of the Indian population identify as Hindu (Census of India, 2011), the caste system has resulted in a large minority group. One of these minorities is identified as the "Scheduled

Caste" (SC)⁴ by the Government of India and they form about 17% of the Indian population (Census of India, 2011). Several articles of the Indian Constitution introduce the nondiscriminatory laws. For example: Article 16(2) of the constitution states that "No citizen shall, on grounds only of religion, race, caste, sex, descent, place of birth, residence or any of them, be ineligible for, or discriminated against in respect or, any employment or office under the State." Another subsection of Article 16(4) states that "Nothing in this article shall prevent the State from making any provision for the reservation of appointments or posts in favor of any backward class of citizens which, in the opinion of the State, is not adequately represented in the services under the State." Similarly, Article 46 states that "The State shall promote with special care the educational and economic interests of the weaker sections of the people, and, in particular, of the Scheduled Castes (SC) and the Scheduled Tribes (ST), and shall protect them from social injustice and all forms of exploitation." The provision for the political reservation/quota comes from Article 334(a), which states that "Reservation of seats and special representation to cease after ten years notwithstanding anything in the foregoing provisions of this Part, the provisions of Constitution relating to the reservation of seats for the Scheduled Castes and the Scheduled Tribes in the House of the People and in the Legislative Assemblies of the States." The reservation of political seats for different minority groups under Article 334(a) was supposed to expire in 1960. However, the 8th amendment of the Constitution extended the provision for an additional 10 years. Since then, it has been extended several times by the 23rd,

⁴ Scheduled Caste does not include all of the "lower caste" population. Brahmnin and Kshatriya are typically considered to be "higher caste"; Vaisya and Shudra are considered to be "lower caste". SC designation only applies to those who are Shudras and who fall outside the caste system, namely Dalits. The caste system is explained in detail in section 2.2.

45th, 62nd, 79th and 95th amendments. The 95th amendment extends the political reservation for the minority groups until January 26th, 2020.

Even though the quota system in India has been in place for over sixty years, the social and economic standings of the SC and ST population have not dramatically improved. Over the past few decades when India saw rising economic growth, the poverty levels among the SC population continued to grow despite various measures taken by the State and Federal governments to help the poor population.⁵ Many studies have pointed out that, over the past two decades, the SC population has seen a growing incidence of poverty, poor education outcomes, higher levels of unemployment, higher childhood mortality, and declining rates of consumption shares (Thorat, 2007; Teltumbde, 1997 & 2001). There have been supporters and critics of the quota system but the policy has seldom become a national political conversation. Politicians are unwilling to debate the issue since it would be a controversial political move and the media and the public are cautious about discussing it, likely because arguing to replace the system with something else might be seen as being "anti-minority".

However, the quota system has generated interest among social scientists and economists over the last decade. The primary question of interest has been "how successful has this program been?" "Success," however, has been variably defined by different economists. Some economists have studied the system's impact on poverty (Chin & Prakash, 2011; Bardhan, Mookherjee & Torrado, 2010) and others have looked at the overall transfer of wealth from higher castes to lower castes in society (Mitra, 2015). Economists have also studied the political reservation's impact on policy making (Pande, 2003 and Madhok, 2013). Additionally, some

⁵ National Council of Educational Research and Training, 2015

economists have focused on the system's impact on crimes and atrocities against minorities (Prakash, Rockmore & Uppal, 2015). All of these papers conclude that there has been a positive impact of political reservation on several groups of minorities in India. However, when it comes to the impact on the Scheduled Caste, there has been varying conclusions. This dissertation tries to provide more conclusive evidence about the impact of such laws on the Scheduled Caste.

The primary focus of this dissertation is the education, and health outcomes of the SC group- one of the minority groups that the law targets (ST is the other group). Since the 1950 Constitution of India did not specify a measure of the law's 'success', the true measure of 'success' comes from the intent of the law in the first place. The political reservation is considered to be a constitutional support to those who have suffered discrimination in society and who lack opportunities in different aspects of the society. In other words, it was put in place so that the socio-economic status of the lower caste (including the SCs) population in the society can be alleviated. Various studies6 show that some of the most important issues to improve the socio-economic status of minorities are education and health, among other areas. In fact, in his book Development as Freedom, published in 1999, Dr. Amartya Sen highlights the importance of education and health as indicators and instruments of human development. Sen argues that development stems from freedom and freedom stems from social opportunities such as better education access and healthcare. Hence, this dissertation studies the impact of the reservation laws on two of those important opportunity indicators- education and health.⁷

⁶ Bloom (2007); Sen (1999)

⁷ The Human Development Index (HDI) also adopts education and health as two of the important measures of "human development" in calculating its annual HDI index for various countries.

The contribution that my research will have relative to existing literature stems from the specificity of the dependent variables I use. While existing literature has examined the impact of political reservation on poverty, policy implications, crime, and wealth transfers, it does not address the impact on the SC population's education, and health outcomes directly. For example, Chin & Prakash found that the political reservation for SC did not have any impact on the SC's poverty outcomes. However, there is evidence that the SC quotas have increased transfers from the well-off high caste population to the lower caste population (Pande, 2003). My dissertation makes a contribution by going a step further in the discussion. I empirically test if the political quotas for SC have contributed to better education, and health outcomes for this minority group in India. By providing a definitive assessment of the impact of reservation policies on the specific socio-economic indicators of the minority, this analysis will fill an existing gap and be a significant contribution to development economics literature. As the majority of the SC population falls below the UN's poverty line, this analysis also has implications for the effectiveness of poverty reduction policies in India.

The rest of this dissertation is organized as follows: **Section 2** of this chapter provides an institutional background on the caste system and the political reservation system in India. **Chapter 2** discusses the education impacts of the reservation system, outlining the literature review, data, empirical strategy, and the results. **Chapter 3** discusses the health impacts of the reservation system, outlining the literature review, data, empirical strategy, and the results. **Chapter 3** discusses the health impacts of the reservation system, outlining the literature review, data, empirical strategy, and the results. **Chapter 4** summarizes the dissertation. The **Appendix** presents the summary statistic table, the results table, and an alternative approach taken in this paper

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1.2 INSTITUTIONAL SETTING

1.2.1 Caste System In India:

One of the large minority groups in India originates from a deep-rooted belief in a social-stratification system — the caste system. A caste system is a form of social stratification unique to Hinduism, which groups people hereditarily into distinct occupational divisions. Most Indian languages use the term "jati" to describe the hereditary social structures known to us as the caste system. The term "caste" comes from the Portuguese term "casta" — which means "race". It is commonly believed that Portuguese travelers came to India in the 16th century and they were exposed to the race-based social stratification system that was already prevalent in India. Hence, the word "casta" was used — to describe what they saw. The term has since evolved into "caste" and is widely used to describe social classifications that are passed along through generations. ⁸

The origin of the caste system in the region is still debated among scholars. A Hindu person would explain the origin of the caste system by referring to the Lord Brahma, a four-handed and four-headed Hindu god, whose hands and heads represent the four castes within which people are expected to arrange themselves. Indeed, the earliest written evidence of the caste system comes from an ancient Hindu scripture, *Vedas*,⁹ with references to Lord Brahma. Even though one of the four *Vedas*, the *Rigveda* (1500 – 1200 BC), lacks direct reference to the

⁸ Ushistory.org. (2016). The Caste System. Ancient Civilizations Online Textbook. Retrieved from http://www.ushistory.org/civ/8b.asp.

⁹ *The Vedas* are a collection of hymns and poems written between 1500 and 1000 BC in India. This collection is considered sacred in the Hindu religion. *The Vedas* are composed of four different Vedic texts- *the Rigveda, the Samaveda, the Yajurveda and the Atharvaveda. The Rigveda* is the largest of these texts, and is considered to be the most important.

caste system and indicates the importance of social mobility, *The Bhagvad Gita*¹⁰ (200 BC – 200 AD) highlights the importance of the caste system in the general society. In addition, the *Manusmriti*, a part of *The Bhagvad Gita*, specifically outlines the rights and duties of the four castes.

A second set of theory on the origin of the caste system refers to a biological adaptation. According to this theory, people in India simply specialized to their God-given skills and attributes- which, over time, transformed the society into a caste system¹¹. However, a more socio-historical theory links the origin of the castes to the Aryan civilization. The Aryans are thought to have come to India in about 1500 BC, mostly from south Europe and north Asia. The theory claims that The Aryans must have introduced the caste system as a means to influence and control the local population. By the time they arrived, there were other non-Indian groups already in India, including the Negrito, Mongoloid, Austroloid and Dravidian. In order to secure their political and economic influence in the society, and to separate themselves from the "inferior" local population, the Aryans introduced some social and religious rules wherein specific groups of people would do specific types of work and only the Aryans could be the priests, warriors and businessmen in society. They were able to maintain and enforce this structure because of their strong army. This structure paved the way to the more established caste system that we know today.

¹⁰ *The Bhagvad Gita,* which is written in the ancient language of Sanskrit, is the most revered Hindu scripture. Even though the origin date of *The Bhagvad Gita* is still contested, most scholars believe that it was written between the fifth century and the second century BC. The Bhagvad Gita is not a part of the four Vedas.

¹¹ Ushistory.org. (2016). The Caste System. Ancient Civilizations Online Textbook. Retrieved from http://www.ushistory.org/civ/8b.asp.

A commonly recurring theme in theories about the origin of the caste system, however, involves the idea of division of labor. The caste system simplifies division of labor by dictating that people who belong to different castes are responsible for specific and separate jobs. People would voluntarily assign themselves in doing the jobs that they are good at doing. Over time, this led to a more formal social stratification system. According to the established system, people belong to one of four castes: Brahmin; Kshatriya; Vaisya; Shudra. People may also be born outside of the caste system, in which case they are called Dalits ("untouchables") and not permitted to join any caste. As with the castes, the Dalits are not able to climb up the social ladder. However, they are part of the Scheduled Caste designation. It is important to note that the caste system only applies to people choosing to follow Hinduism. Brahmins are designated to serve as priests; Kshatriyas as warriors and nobilities; Vaisyas as farmers, traders and artisans; and Shudras as tenant farmers and servants. The hierarchy in the social setting is such that the Brahmins are considered the highest caste and Shudras the lowest caste. Untouchables are assigned no occupational role in the society.

Although a caste designation originally depended upon the kind of job a person did, with time, it became a strictly rigid system in which the caste was assigned at birth and passed down through generations. Each person's birth assigned caste is unalterable during his/her life and is also unalterable for his/her children, grandchildren, and so on. Even though the beginnings of this caste rigidity are hard to trace, it is believed that it is a product of self-interest on the part of ruling emperors throughout India's history. However, historical evidence does confirm that India's caste system was not originally as absolute and rigid as it has become in modern times. In fact the Gupta Dynasty, which ruled India from 320 AD to 550 AD, in addition

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to the Madhuri Nayaks dynasty, who ruled India from 1559 AD-1739 AD, originated from lower castes. It is important to note, however, that from the 12th century onward, much of India was ruled by Muslims. During the initial Muslim rule, the Hindu population was marginalized up to the point where there were hardly any reported populations of Brahmins and Kshatriyas in northern and central India. The Muslim rule created an anti-Muslim sentiment among the Indian Hindu population which further strengthened the caste system by inspiring Hindus to strictly adhere to the elements of their faith, including the caste system. The British took over India from Muslim rule in 1757 AD, thereby lessening the animosity between the Muslim Indians and the Hindu Indians. Some of the documented anecdotal evidence reveals that the British were able to exploit the caste system as a means of social and political control over India. However, they also identified the underlying discrimination brought upon the lower caste population (mostly Shudras and Vaisyas), and, responsively created some of the initial antidiscriminatory laws for the "Scheduled Caste" population of India during 1930-1940.

Typically, the Brahmins and the Kshatriyas are considered "higher caste", and Shudras are considered the lowest caste in the hierarchy. However, the official "Scheduled Caste" designation only includes and protects the Dalits ("untouchables") who fall outside the caste system¹² which represent about 17% of the national population¹³. Some of the Shudra population, who were historically disadvantaged, are officially classified as "Other Backward Class (OBC)" and also have similar protections in place by the Government of India. This dissertation strictly focuses on the SC population.

¹² Scheduled Caste designation is based on a household's last name. All of the Dalits fall under this designation. ¹³ The Scheduled Tribe (ST) does not fall under the "Dalit" designation. The STs are protected because of their indigenous heritage.

Historically, the population of "lower caste" always outnumbered the population of "higher caste"¹⁴. According to the last census taken during the British Empire (1931), the lower caste represented 54% of the total Indian population. India has never taken a caste-based census since its independence from Britain. However, according to four different surveys conducted by the Center for the Study of Developing Societies, between 2004 and 2007, the Brahmins represented about 5% of the national population and Kshatriyas represented 3.46% of the national population. Even though the populations of Brahmins and Kshatriyas do not represent the majority, they constitute a vast number of public officials and politicians in India. According to the same surveys by the CSDS, Brahmins alone constituted 47% of the chief justices and 40% of the associate justices in India. Even though the representation of higher caste in the upper and lower house of the legislative branch of India has been falling — thanks in part to the quota system – the representation is still disproportionate. Brahmins alone represent about 10% of the Members of Parliament (MPs) and hold a significant proportion of public offices. The higher caste population was able to stay in control of governance and form a "majority" in Indian society because of the better education and increased opportunities that they received by virtue of their birth-assigned social rank. The extent of the opportunities they received was also expanded during the British rule, which is why the Brahmins were able to grab a majority of government positions after India's independence from the British. Vaisyas, Shudras and Dalits have comprised a significant portion of the Indian minority in governance for decades. Specific tribes and some religious groups (Muslim and Christian, in particular) form the other part of

¹⁴ The Brahmins and Kshatriyas are generally considered "higher caste". However, in most of the surveys conducted only Shudras and Dalits (untouchables) are considered "lower caste".

the Indian population and governance minority. Even though discriminating against anyone based on caste is illegal in India, peoples' belief in the ranking inherent to the system is strong.

The fact that people must remain in their birth-assigned castes for life means that members of the SC population face a significant obstacle in climbing up the economic ladder. Opportunities for obtaining a higher social status, better education, well-paying jobs, and political involvement, do not exist for these groups of people. It is for this reason that the Indian government decided to implement a quota system wherein some seats (and/or positions) are set aside for people of certain minority classes in employment¹⁵, education and in different states' legislatures. These minority groups are referred to as the "Scheduled Caste" (SC) and the designation is based on the last names they carry. The Indian Constitution of 1950 lists 1108 different last names under the "Scheduled Caste," making these people eligible for participation in the quota system. The tribal minorities are also protected under the quota system and are referred to as the "Scheduled Tribe" (ST). While the SC classification targets the caste minority and "untouchables", the ST classification targets the population who live in tribal areas (including forest dwellers). There are 744 different tribes listed in this category under the Indian Constitution of 1950.

The SC population is one of India's most vulnerable groups despite its growing size. According to the 2011 national census, the SC population was a little over 201 million, which is about 17% of India's total population. Long before the 1950 Constitution of India, this group of people was considered to be the class of people who did not deserve the rights and protections granted to the rest of the general population. Because of their position within the caste system,

¹⁵ In federal employment, certain percentage of the hiring is assigned for SC, ST or the OBC.

the SC population has experienced persistent discrimination and a lack of opportunities in many socio-economic dimensions. The SC group has been historically deprived of access and entitlement not only to economic and political rights but also to social needs such as basic education, health care, and housing. Historical discrimination and exclusion in access to land, capital, education, and health care have led to high levels of economic deprivation and poverty among the SCs. The lack of job opportunities has resulted in an SC population with an exceptionally high dependence on manual wage labor.

Data¹⁶ shows that SCs have lower levels of job market skills, higher underemployment, and lower wage rates compared to rest of the general population because of the limited, if not excluded, access to capital, land, and education they receive (Saravanakumar & Palanisamy, 2013). According to the National Sample Survey 2011, the average monthly per capita expenditures (MPCE)¹⁷ for SC in rural areas was Rs. 1252 (\$18) and in urban areas was Rs. 2028 (\$29). However, the remainder of India's population had MPCE of Rs. 1719 (\$25) in rural areas and Rs. 3242 (\$47) in urban areas- both values significantly higher than those for the SC's. Similarly, the Census of India 2011 shows that only 50.9% of total SC households used some form of banking services compared to 59.9% of the non-SC households. The census data also shows that while only 59% of the SC households had access to electricity, 67% of the non-SC households had access. Likewise, the National Sample Survey 59th round (2003) shows that only 43.3% of SC households owned some land compared to 62.2% of the non-SC households.

¹⁶ Data and statistics referred to here come from the following sources: Census of India 2011, The National Commission for Scheduled Castes (2010), Ministry of Human Resource Development (2011), and the Annual Report of University Grant Commission 2010.

¹⁷ MPCE is the per capita final consumption expenditure of good and services by households.

It is important to note, however, that the *Scheduled Caste and Scheduled Tribe (Prevention of* Atrocities) Act was enacted in 1989. The Act replaced the caste-based customs relating to property rights, employment, wage, and education with a more egalitarian legal framework, under which the SC and the untouchables are granted equal access and rights compared to the general Indian population. However, despite the legal change, SC's actual access to land and other income-earning capital has barely improved. The caste system has continued to function in the private domain of the economy in modified and changing forms. Empirical studies on labor, education, housing, and health services have revealed the persistence of market discrimination towards lower castes, particularly the Dalits (Thorat & Deshpande, 1999; Shah, Deshpande, Mander, Baviskar & Thorat, 2006). The studies conclude that such discrimination, despite laws banning it, results in lack of access to capital, more unemployment, and poor human development, which all together culminate in high poverty and deprivation among the lower caste population (including the SCs). Some studies also highlight the exclusionary and discriminatory working of private industrial labor markets (Papola, 2012). It is also important to note, however, that enforcement of anti-discriminatory laws is often neglected, in part, because of the deep-rooted belief in the caste system within the law enforcement and the private domain of the society.

1.2.2 Reservation in State Legislative Assemblies for the Scheduled Caste:

Since this dissertation focuses on the role of political reservation on the well-being of the SC population, it is important to understand the way seats are reserved and changed for the SC. It is also critical to understand the role of elected state legislatures in forming policies that affect

education, among other sectors. The political reservation in India applies to all states because of a provision in Article 332 of the Constitution of India (effective January 26, 1950). Specifically, the article mandates representation of the SC (and the ST) in the State Legislative Assemblies and the lower house of the Parliament (*Lok Sabha*.) The numbers of seats reserved for the SC are to be proportionate to their share of the total population in a state. More specifically, the proportion of the SC seats in the state assembly is equal to its share of the population as measured by the latest national census. A "Delimitation Commission" is to be formed after every national census is conducted. The role of the Commission is to demarcate the constituency boundaries based on the new population data from the census. Moreover, the commission revises the number of seats reserved for the SC in each state based on the revised constituencies and the SC's new population share. This policy rule allows for some variation in the SC political reservation. More specifically, the variations come from the following elements:

i) With every new census, the number of reserved seats for the SC changes. The census is taken every 10 years. After the census is conducted, a Delimitation Commission is formed. Members of the committee include two Supreme Court and/or high court judges, the chief election commissioner, and a nominee by the central government of India. One of the tasks assigned to the committee is to update the seats quota in the *Lok Sabha* and the State Legislative assemblies based on the newly drawn constituency boundaries. In this duty, the committee is simply matching the reserved seat proportion to the SC population share based on the newest census. Since the total available seats do not change, the committee has to

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change the SC quota seats to match the proportion to population proportion. Thus, the actual number of reserved seats changes every 10 years.¹⁸

- *ii)* Since the number of reserved seats can only be a whole number, there is always a difference between the actual SC population share and the reserved seat share. The small variation in the quota comes from the discrepancy between the required number of reserved seats and the number of seats that are actually reserved given that the number must be an integer.¹⁹
- When the definition of the SC changes, the share of reserved seats changes. The defining
 Constitution of India (1950) listed 1,108 different last names as belonging to the
 Scheduled Caste category. However, by 2008, 100 additional last names had been
 added to the list. The list has been growing with the last addition approved in
 December, 2015²⁰. With each added name, the total SC population and hence, its
 population share, grows. This, in turn, increases the required seat quota for the
 SC.
- iv) Year-to-year variation exists because while the SC population size changes every year,
 the actual reserved seats do not change as frequently. Since the Delimitation
 Commission assigns the reserved seats based on the most recent decennial

¹⁸ For example: The 2001 census shows that the SC population in Bihar was 13,048,608 while its total population was 82,998,509. Hence the SC formed 15.72% of Bihar's population in 2001. Once the Delimitation Commission is formed, its job is to match the reserved seats proportion to something close to 15.72% of the total available seats. The Commission finished its job in 2005 when there were 243 seats available in Bihar. Thus, 15.72% of 243 yields 38.20. The committee established 39 SC-reserved seats 2005 (the policy is to round up to the next higher integer).

¹⁹ For example: The ideal number of seats to be reserved in Bihar after the 2001 census was 38.20. However, 39 seats were chosen to be reserved. Thus, there is a variation of 0.80 in the quota.

²⁰ The addition is mostly due to the demands from several minority groups to be included under SC, so that they can have political representation and other benefits that the SC gets.

census population, but the actual SC population changes yearly (continuously over any time period), there is year-to-year variation in the SC's quota share.

v) The independent variable of interest, QS (Quota Share), varies from year-to-year because the ratio of reserved seat share to population share changes every year. QS is defined mathematically as:

$$QS = \frac{\frac{Quota\ assigned\ for\ SC}{Total\ seats}}{\frac{SC\ Population}{Total\ population}}$$

Given the above definition, most of the variation in the variable comes from the fact that the SC's population share (the denominator) changes every year while the quota proportion (the numerator) changes less frequently. QS is the variable of interest because the constitution requires the reserved seats to be proportional to the population share of the SC (numerator is equal to the denominator). However, the quota proportion (the numerator) and the SC population proportion (the denominator) are only equal every 10 years when the Delimitation Commission meets to determine the quota proportion that matches the SC population proportion. However, because of the point (ii) mentioned above, the QS variable will rarely equal exactly 1.

CHAPTER TWO: EDUCATION

2.1 INTRODUCTION

The education system in India consists of nine levels: preschool; play school; kindergarten; primary school; upper primary (middle) school; secondary school; higher secondary (pre-university) school; undergraduate; and postgraduate. This paper is concerned with education in primary, upper primary, secondary, and higher secondary schools. Primary school (for children aged 6-10 years) comprises grades 1-5; upper primary school (for children aged 11-14 years) comprises grades 6-8; secondary school (for children aged 15-16) comprises grades 9-10; and higher secondary school (for children aged 17-18) comprises grades 11-12. In general, India has seen a rising trend in the literacy rate from just 24% in 1961 to 74% in 2011. A similar trend is seen for the SC population. In 1961, the literacy rate among the SCs was just 10.3%. The rate has increased to 56.5% according to the 2011 census, thanks to several initiatives taken by the central and state governments over the past few decades. However, the literacy rate of the SC population is still significantly lower than that of the remaining Indian population [Refer to table 1.01].

The increasing trend in literacy rates is primarily driven by the increasing trends in enrollment rates in various levels of schooling. The increase in enrollment rates, however, exhibits a similar gap between the general population and the SC population. According to the 2001 census, there was a 10% difference between the primary school enrollment rates of the general population and those of the SC population, which were lower [Refer to table 1.02].

Several factors may explain the low enrollment and literacy rates among the SC population. One of the biggest factors may be the lack of educational resources and infrastructure in the rural parts of India. Several papers cite the lack of well-trained teachers, lack of textbooks, and lack of schools in rural parts of India as the biggest hindrance to the education outcomes of children in these parts.²¹ According to the Census of India 2011, 76.5% of the SC population lives in rural areas, compared to 67% of the non-SC population. Thus, the SCs are in a position to be most adversely affected by the lack of educational resources in these areas. In West Bengal, one of the highest SC-populated states in India, 73.3% of the population lives in rural areas, according to the same 2011 census. Similarly, an increase in the privatization of schools in several states of India may have created a disincentive for members of the SC caste to send their children to schools. This is because private schools are generally more expensive than public schools and the SC population has a lower income level compared to the general population. The National Sample Survey Round 50th to 68th (years 1993 to 2012) has consistently shown that an SC person earns about 38.6 % - 40.8% less in a year than a non-SC person in India. Hence, privatization of schools might lead to fewer enrollments of SC students in costly private schools since the benefits of education to the SC do not increase at the same rate as the cost.

Perhaps a more subtle but compelling reason why the SC school enrollment lags behind general enrollment is the influence of culture and social status on decision making. For instance,

²¹ Education for rural development: towards new policy responses, 2003; The World Bank; Illiteracy Looms Over India, The CS Monitor, 1990.

the dowry system is still prevalent in many parts of India, especially among the lower caste population, including the SCs. According to the dowry system, a bride's family is expected to give money and furniture to the groom's family at the time of the wedding. The requested dowry is always more for an educated girl compared to an uneducated girl. This might provide a sufficient incentive for SC parents to refrain from sending their daughters to school. In fact, according to a report published in 2011 by the Ministry of Human Resource Development (MHRD) of the Government of India, there are 94 female SC students for every 100 male SC students enrolled in primary and upper-primary school levels. Similarly, the dropout rates for female students are much higher than those for male students within the SC population and the dropout rates for SCs overall are higher than those of non-SC caste members²². Additionally, there have been multiple reports of SC children being targeted for bullying and public humiliation by higher caste students as well as teachers (Saravanakumar & Palanisamy, 2013). An expectation of their children being bullied in schools might prevent SC parents from sending their children to schools; resulting in both lower SC enrollment and literacy rates.

Using state-level data, in this paper I examine the impact of political reservations on the education outcomes of the SC population. I first use the fixed effect model to establish some of the mediator variables that the quota system influences. I then make use of these mediator variables and the variations in the quotas created by the timing of the elections and the proportion of SC population, to estimate the effect on education outcomes using the dynamic panel data model (DPD). The DPD model is the best fit for the data because of the

²² According to the Educational Statistics at a Glance report published by the Ministry of Human Resource Development Government of India in 2011, the primary school dropout rate among SC males was 22.3% and among SC females were 24.7%. Likewise, the primary school dropout rate for all of India was 22.3% and for the SC population as a whole was 23.5%.

autoregressive nature of the dependent variable, which results in contemporaneous endogeneity. The DPD controls for this endogeneity problem, which is not corrected by the first difference or the fixed effect models. I find that India's political reservation policy has a statistically significant and positive impact on state-level gross enrollments of SC students in all levels of schooling. Similarly, I find that the political reservation has a statistically significant and negative impact on the state-level dropout rates of SC students.

2.2 LITERATURE REVIEW

There are many political economy papers that focus on the impact of affirmative action policies. Galanter (1984) provides a very thorough qualitative overview of affirmative action policies affecting SCs, in which he provides a description of different AA policies in the sectors of employment, education, and politics. Another often-cited work in this field is An Economic Theory of Democracy (1957), in which Anthony Downs theorizes that candidates who contest for election for public office are likely to commit to their party's agendas. Down's work is helpful in understanding how a public office candidate from a minority group might have a bias towards the agenda and policy preferences of the minority group which elects him or her. The theoretical base for my research is also built upon the work of Osborne and Slivinski (1996). Their citizen-candidate model concludes that the identity of the candidate running for public office has an impact on policy determination. This contributes to my understanding that there is a mechanism through which elected officials are able to influence policy measures, in a way which may affect the education outcomes of the minorities. For instance, the elected minority candidate might have an influence on the amount and distribution of a state's annual education

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and welfare expenditure, which might have an impact on the well-being of the minorities he or she represents. According to a report, The Legislature and the budget, published by The World Bank in 2004, elected legislators in countries including India, Uganda and Zambia, have significant influence on state's annual budget allocation. When discussing India, the report concludes that legislators have direct influence because state budget allocations are decided by committees of elected legislators. The research endeavor described in this paper attempts to link existing literature to empirically study the impact of the unique policy of political quotas on the education outcomes of the SC group.

Several attempts have been made to empirically test the policy outcomes of political reservation at the small scale level of village council (*the Gram Panchayat*). Most states have some seats reserved for the SC, ST, or women for the position of chief of the *Gram Panchayats, Pradhan*. Duflo and Chattopadhyay (2004) used the village-level data for the states of Rajasthan and West Bengal to empirically test the effect of women *Pradhans* on policy determination. They found that the Gram Panchayats that elect a female as their *Pradhan* tend to redirect more public goods towards sectors that benefit the village, such as access to clean drinking water and paved roads. Besley, Pande, Rahman, and Rao (2004) conducted similar research on three other states²³ to understand the impact of identity of a politician on the provision of public good. They found that the Gram Panchayats that elect SC (or ST) as a *Pradhan*, have a significantly higher probability of assigning public resources towards toilet, electricity connection and private water lines. Duflo, Chattopadhyay (2003) also found that a Gram Panchayat with an elected SC Pradhan directs public resources significantly more to the places where the SC population is

²³ The studied states were Karnataka, Tamil Nadu, and Andhra Pradesh.

high. More importantly, Bardhan, Mookherjee, and Torrado (2010) found evidence that an SC Pradhan in a Gram Panchayat in West Bengal secured more benefits from the State and local governments for public infrastructure improvements such as housing and toilet construction than did non-SC Pradhan in other Gram Panchayats. These papers provide further empirical evidence that an elected SC member does have a significant impact on policy determination. I thus have a solid framework for making an investigation into the direct and indirect influence of an SC legislator on the socio-economic issues most pressing for the SC, namely education. I expect to find that the established quota allocation has a positive and significant impact on SC enrollment rates in all levels of schooling.

In the past decade, various researchers have focused attention on the economic outcomes produced by political reservation for women. Duflo and Chattopadhyay (2004) found that, in at least one location in India, the quota reserved for women leaders in public office had a significant impact on the provision of public goods in the location where the women were representing. Specifically, the researchers found that the public goods provided were in accordance to the necessities and priorities of the society. Similarly, Lott and Kenny (1999) have shown that the gender of the legislators have different priorities in their policy implementation, while Edlund & Pande (2002) show that male and female tend to have different political stances on major social issues.

Pande (2003) used the variation in the quota allocation to examine the effects of the political quota provision in state legislatures on various policy outcomes. Unlike previously mentioned papers, she used state-level data for 16 of the largest states of India to analyze the impact of such policies on the transfers to minorities using state and time fixed effects. Pande

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found that political reservation for SC (and ST) in state legislatures have increased transfers from higher caste to lower caste population as the law was intended to do. More specifically, she determined that the SC political reservation has increased the proportion of public jobs reserved for SC members. Her results confirm the citizen-candidate model previously discussed, in that they provide specific evidence that a legislator's identity affects policy outcomes.

Chin and Prakash (2011) expand on Pande's (2003) research by examining poverty as an outcome. If SC legislators have an influence on policy outcomes (as suggested by Pande, 2003), then, presumably, they have an incentive to prescribe policies that improve the economic outcomes of the minorities they represent. Working on this theory, Chin and Prakash (2011) used state level data to study the impact of reservation policies on the overall poverty of a state. Interestingly, they found that increasing the number of seats reserved for SC has no impact on the overall poverty of the states. However, they did find that political reservation for ST has statistically significant impact on poverty reduction wherein a one percentage increase in the ST seat in the state assemblies would decrease the aggregate poverty level in that state by 1.1% . Thus, in agreement with Pande (2003), the authors are able to conclude that the identity of minority legislators can affect key outcomes for represented minorities.

All of the above mentioned studies together show that India's political reservation schemes have effects on the policy outcomes and the provision of public goods. Most of the studies also show that there are some redistribution effects of the reservation in favor of the targeted minority groups. However, there is no evidence that the reservation is helpful for the SCs in every sector of their livelihood. In other words, we do not yet know whether the political quota

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provision will always increase the flow of public resources to minorities in every sector. Hence, it is imperative to empirically examine whether the reservation system has had an impact on the basic human development outcomes of the minorities. This paper approaches the question by estimating the reduced-form effects of the SC political quotas on their education outcomes. According to the twelfth "five year plan" of the Indian government which spans 2012-2017, improvement of education, health, and employment outcomes of citizens are among India's top socio-economic priorities. Therefore, it is of interest to know whether the constitutionally assigned quota system for the SC has been able to improve those outcomes for the SC. While most of the papers in the literature have focused on understanding the impact of political quota provision on inputs to the welfare of the SC population such as funding and infrastructure, this is the first paper to estimate the impact on the outcomes by understanding the causal effects of India's affirmative action on the SC's education outcomes- namely, enrollment and dropout rates. Previous research, as mentioned above, focuses instead on AA's impact on policies and transfers of wealth, which serves as inputs that might have an effect on the outcomes. Although these investigations brought about a preliminary understanding of the net effect of AA on minorities, they did not quantify the impact on the SC's overall well-being. This paper, thus, fills this crucial void in the existing literature.

2.3 DATA

This paper uses annual data from various sources for 25 Indian States and 3 Union Territories²⁴ for the years 1990-2011. The states of Chhattisgarh, Jharkhand, and Uttarakhand (previously known as Uttaranchal) were formed in November of 2000 and hence, only enter my dataset after 2000. The state of Telangana was formed in June of 2014 after it was split from the state of Andhra Pradesh. For this reason, Telangana is not included in my dataset. The Union Territories of Dadra & Nagar Haveli, Andaman & Nicobar Islands, Daman & Diu, and Lakshadweep were not included in the dataset because of a lack of availability of relevant data and their smaller population size compared to the rest of the states/territories in the dataset. The state of Punjab has the highest SC population proportion (28%) and the state of Gujrat has the lowest (7.41%), according to the 2011 census. Both of these states are included in my dataset. These data form an unbalanced panel because of missing observations for some variables for some years. For example, my dataset does not have SC enrollment for primary school for the years 2009 and 2010 due to unavailability of reliable data. Likewise, the SC enrollment data for secondary and higher secondary level of schooling only spans 2006-2011 in the dataset. While the quota variable has observations from 1990 to 2013 (with few missing observations for some states), the GER has large number of missing observations for Arunanchal Pradesh, Assam, Jammu & Kashmir, Meghalaya, Sikkim, and Tamil Nadu. The state of Jammu & Kashmir has not kept reliable records of the SC school enrollment because of the armed conflict with Pakistan since 1947 except for the years 2001-2008. The summary statistics of the variables are presented in Table 1.04 in the Appendix. Because of the unbalanced panel, the number of

²⁴ Union Territories are the administrative divisions in the Republic of India. These territories are not an established State of India.

observations on table 1.04 is different than the number of observations in the individual regression tables (tables 1.06- 1.10.) Methods used to deal with unbalanced panel are discussed in the empirical framework section. All of the included variables are in log form; to avoid the loss of zeros in some of the variables, I have added one before taking the logs. A description of the variables used and their sources is provided below.

2.3.1 Education Data

The education outcomes of the scheduled caste group can be measured using various indicators of education. In this paper, I am particularly interested in the enrollment and dropout rates of SC students in primary, upper primary, secondary, and higher secondary levels of schooling. I use the Gross Enrollment Ratio (GER) for these levels as my dependent variables. GER is defined as the ratio of the number of students who are enrolled in a school to the number of children in the population who are of the corresponding school enrollment age. The United Nations Educational, Scientific and Cultural Organization (UNESCO), officially describes 'Gross Enrollment Ratio' as the total enrollment within a country "in a specific level of education, regardless of age, expressed as a percentage of the population in the official age group corresponding to this level of education." The variables will be referred to as GERP²⁵, GERUP²⁶, GERS²⁷ and GERHS²⁸ for primary, upper primary, secondary and higher secondary gross enrollment ratios, respectively.

²⁵ The GERP (Gross Enrollment Ratio- Primary) variable is calculated as SC primary enrollment divided by the primary-school-going-age population (aged 6-10 in the SC population).

²⁶ The GERUP (Gross Enrollment Ratio- upper primary) variable is calculated as SC upper primary enrollment divided by the upper-primary school going age population (aged 11-13 in the SC population).

Another measure of education outcome which is commonly used in research similar to my own is net enrollment ratio (NER). In contrast to GER, NER is defined as the ratio of number of children of official school-going age who are enrolled in school to the total number of children in the population who are of that age group, expressed as a percentage. Although NER is an appropriate and respected measure of outcome for education research, I have chosen, in this paper, to use GER instead of NER for one specific reason. That is, GER includes all enrollments for each level of schooling regardless of the age of the student, whereas NER includes enrollments only if the enrolled student corresponds to the defined age group for that particular level of schooling. As an example of this distinction, consider a case of a 12 year old student enrolled in primary school, which is technically defined as being for children aged 6-10. The enrollment of this student is counted in the measure of GER, but not in the measure of NER. In this paper, I wish to investigate the education outcomes of all SC students, regardless of whether or not their enrollment corresponds to officially designated age groups. In other words, for the purposes of this analysis, it does not matter whether an SC student enrolls in primary school at age 6 or at age 14—what matters is that the SC student is attending school. Therefore, I have determined that GER is the more appropriate measure of SC education outcome. Additionally, because the SC are a disadvantaged population, it is reasonable to suspect that SC children are commonly unable to start school at the officially designated age, and/or that they experience gaps in education that place them outside of the official age

²⁷ The GERS (Gross Enrollment Ratio- secondary) variable is calculated as SC secondary enrollment divided by the secondary-school-going-age population (aged 14-15 in the SC population).

²⁸ The GERHS (Gross Enrollment Ratio- higher secondary) variable is calculated as SC higher secondary enrollment divided by the higher secondary-school-going-age population (aged 16-17 in the SC population).

brackets for a grade level. Using GER allows me to measure the education outcomes of these students rather than pretending they do not exist, as would be the case with NER.

In India, primary education is defined as grades one through five; upper primary education as grades six through eight; secondary as grades nine and ten, and higher secondary as grades eleven and twelve. The data for SC enrollment and dropouts were obtained from Indiastat (indiastat.com), a database that compiles socio-economic data released by various ministries of India including the Ministry of Human Resource Development of the Government of India. Some of the enrollment and dropout data that were missing in the indiastat database were collected using the District Information System for Education (DISE), which is an education information system operated by the National University of Educational Planning and Administration (NUEPA). The data cover yearly enrollment and dropout rates for each level from the year 1990 to 2011.

As noted by UNESCO²⁹ and the World Bank³⁰, one way educational output can be measured is by assessing access to education. One of the important measures of "access", as outlined by UNESCO, is the rate of enrollment in various levels of school. Therefore, GER is one of the dependent variables I use to study the impact of political reservation on SC's education. Likewise, SC children have, historically, been excluded from the formal education system, which is why various studies have shown that SC children have the highest dropout rates among any social or minorities groups in India, thereby affecting the gross enrollment ratios (Saravanakumar & Palanisamy, 2013). Because of this, I have also included dropout rates as

²⁹ Atchoarena, D. & Gasperini, L. (2003). Education for rural development: towards new policy responses. United Nations Educational, Scientific and Cultural Organization. Retrieved from

http://www.unesco.org/education/efa/know_sharing/flagship_initiatives/towards_new_policy.pdf.

³⁰ Measuring Outputs and Outcomes in IDA Countries (2002), The World Bank

my second dependent variable because whether or not the SC children are enrolling in school and staying in school, offers an indication of whether or not the elected SC legislators are providing improved access to education for the population they represent.

The education related control variables used in this paper are obtained from Indiastat, the Census of India, the Ministry of Human Resource Development of the Government of India, and the Directorate of Economics and Statistics for state governments. Control variables include: the education budget per SC pupil (EBPP)³¹, the student-teacher ratio³² (STRP, STRUP, STRS, and STRHS), the number of schools³³ (NoPS, NoUPS, NoSS, and NoHSS), and the adult literacy rate. Adult literacy rate is included as one of the controls because an educated SC population should be reflected by an increased rate of enrollment in various levels of schools.

2.3.2 Political Data

Since I am specifically interested in the impact of the political quotas on the well-being of the SC population, I have used number of SC seats assigned under each state's legislative assembly as the variable of interest. State's Legislative Assembly constitutes of representatives (Member of Legislative Assembly (MLA)) elected by the voters of electoral districts to the legislature of a state. Within each state, the seats are assigned at the district-level depending on the population of SC in each district. The total number of seats assigned for the SC in each state is determined by the combination of census data and election results. With a few exceptions, the state elections in India are held every five years. However, not all states have elections during

³¹ Education budget per pupil (EBPP) is calculated as a state's annual education budget (in Indian Rupees) divided by the SC primary to higher secondary-school-going population. The education budget is the amount of a state's annual budget that is spent on education.

³² The student-teacher ratio is the number of SC students for every teacher in each State.

³³ Number of school variable represents the number of schools which teach various levels of education in each states. For example NoPS is the number of schools which teach primary level of education.

the same year. The variable used in this paper is the SC seat percentage (QS), which has been previously described in this paper. The SC seats are assigned as a percentage of total legislative seats available, wherein the proportion is matched to the state's proportion of the SC population. These data were extracted from the Election Commission of India's³⁴ online database. Detailed election results are recorded by the Commission both for the general elections and the state elections. The data for SC seat range from 1990 to 2014. It should be noted that the states formed in 2000 (previously mentioned), only enter the political dataset after their first elections held in 2002. The ElectionYear variable represents the dummy variable with a value of 1 if an election was held that year and with 0 otherwise.

2.3.3 Other Variables

Primary and upper primary education were made free in India's public institutions with the passage of the *Right of Children to Free and Compulsory Education Act of 2009,* which went into effect on April, 2010. In this analysis, I have included the annual per capita income of the SC as a control variable given that the GER includes enrollment in both public and private schools in India. As private schools are costly compared to public schools (they were even before the 2009 law), we should expect an SC family's annual income to have an influence on enrollment rates in all levels of schooling. The SC annual income per capita³⁵ is calculated based on the Gross State Domestic Product (GSDP) of each state. Various surveys and studies have shown that over the past two decades, the average SC annual income is around 40% less than the non-SC annual

³⁴ The Election Commission of India is an independent agency set up and authorized by the Constitution of India to conduct fair elections. It is also the agency authorized to collect and store data related to federal and state elections. ³⁵ The data on annual per capita income of SC in different states in India is not available in the public domain. Even the Census of India does not collect this information. Census of India does, however, have data on the different types of assets that SC households possess at the time of census.

income³⁶ (Bhandari, Dutta 2007). Hence, for every dollar that a non-SC person earns, the SC person earns only sixty cents. The SC annual income variable, is thus calculated as 3*GSDP / (5*total population – 2*SC population)³⁷. The data for the GSDP of each state is collected from Indiastat and the Central Statistical Office. Indiastat compiles its GSDP dataset based on information released by the Ministry of Finance of the Government of India. SC income per capita is an important control variable because household income is frequently identified as one of the primary reasons for school dropouts and lack of enrollments (Chauhan, 2016; Saravanakumar, Palanisamy, 2013). The population data is collected from the Censuses of India 1991, 2001, and 2011. The intercensal population estimates are calculated using the exponential projection of the population based on the census population data³⁸. The same approach was taken to estimate the age-group populations of the SC. Within existing literature, exponential and logistic projections are two of the most-widely used methods for obtaining intercensal estimates. Additionally, SC population density was calculated by dividing the total SC population of each state by the total area of the state, as this variable was not otherwise available. This variable is included to control for the effects of migration of SC population from one state to the other.

³⁶ The National Sample Survey Round 50th to 68th (1993 – 2012) consistently shows 38.6 %- 40.8% less income for the SC population. (Source: Ministry of Statistics and Program Implementation, Government of India. http://mail.mospi.gov.in/index.php/catalog)

³⁷ If A is the per capita income of the non-SC, then the SC's per capita income is 0.6A. Hence, A*(total population-SC population) gives the total income share of the non-SC in the GSDP; and 0.6A*SC population gives the total income share of the SC in the GSDP. These two numbers have to add up to the total GSDP of the states. After some algebra, we find that the per capita non-SC income is 5*GSDP/(5*total population – 2*SC population), and, 60% of that number is 3*GSDP/(5*total population), yielding the per capita SC income.

³⁸ The log-linear function of the exponential population change is given as log(P(t)) = log(P(0)) + t*log(1+r). For example, the population in 1992 is given by exp(Ln(Pop1991) + (ln(Pop2001/Pop1991)/)2001-1991)* (1992-1991))

2.4 EMPIRICAL FRAMEWORK

To estimate the effect of political reservation on the education outcomes of the SC population, I take advantage of the variations in the explanatory variable (QS) over time and across states as caused by the above-mentioned factors. Since the quota is assigned according to the proportion of SC population after every census, the variation caused in QS is exogenous to the education outcomes of the SC. I assume linear relationships between the SC primary, upper primary, secondary, and higher secondary gross enrollment ratio (GERP, GERUP, GERS, GERHS), the explanatory variable of interest, quota share (QS), and the vector of control variables. The empirical work for this paper is focused on 26 states and 3 union territories of India for the years 1990-2012. I first test the mechanisms through which the quotas could have an influence on the GERs using the fixed effect regression. Secondly, I use the dynamic panel data (DPD) approach to get my estimation because of the fact that there might be some lag effects of enrollment from one year to the other. As an alternate method of analysis, I also use the Structural equation model (SEM) to estimate the total effect of the quotas on the gross enrollment ratios via several mediator variables (see Appendix C.)

Since the panel data is unbalanced, it is important to test for sample selection bias. If the reason of unbalanced panel is some systemic pattern within states, then the selected sample of data is not stochastic. Therefore, a bias could be present by the selection of states which is not randomized which ensures that the selected data sample is not representative of the population. I have conducted the Heckman Selection Model to test if the dataset has the selection bias. In the model, rho is the correlation between the errors in the two equations- outcome and selection equations. Therefore, the null hypothesis rho = 0 means that there is no selectivity bias. For each

of the dependent variables, we fail to reject the null hypothesis that the rho is zero. The p-value for GERP, GERUP, GERS, and GERHS is 0.47, 0.64, 0.48, and 0.97 respectively. This means that the dataset does not have the sample selection bias. This result is consistent with the expectation since sample selection bias exists only if the dependent variable is observed exclusively for a restricted and nonrandom sample from the population. However, we know that both the enrollment and dropout rates are uniformly observed throughout the sample. Hence, the sample selection can be assumed exogenous since the unbalanced panel is largely caused by a lack of data for entire years at a time. More importantly, some of the observations are missing because of pure lack of availability of reliable data instead of some exogenous unobserved event. Therefore, even though the unbalanced panel is not ideal, there is no reason to believe that there is a big loss of efficiency or a significant bias. More importantly, the DPD method used in this paper can handle unbalanced panel data better than the fixed effects of random effect methods since the DPD approach does not rely on taking the first difference thereby not maximizing the gap within the data. Instead, the DPD estimator uses the orthogonal deviations which preserve the degree of freedom while avoiding the selection bias.

It is also important for me to recognize a mechanism through which the political quotas might have an impact on the education outcomes of the SC population. It is clear that the intent of such quota policies is to have a positive effect on the well-being of the minority groups but a direct link is not as clear. Various studies have theorized an idea of a "role-model" effect, wherein a representative from a minority group serves as a role model to that group if that individual is recognized and accomplished in their fields (Fang, Moro 2011; Robb 1999). Duflo and Chattopadhyay (2004) investigated the importance of the "role-model" effect on the socio-

economic status of women. They discovered that in areas with long-serving female Pradhans in Gram Panchayat, female students tend to set higher goals in their education and hence the gender gap in various education outcomes has shrunk. Likewise, they also found that SC parents were 25% more likely to have more ambitious education goals for their daughters, which in turn helps narrow the gender gap. The authors believe these results stemmed from the "role-mode effect: Seeing women in charge persuaded parents and teens that women can run things, and increased their ambitions. Changing perceptions and giving hope can have an impact on reality." In the same way, an SC member who becomes an elected state legislator could be viewed as a role model for other SC members in the state to achieve more and hence, could drive up enrollment in various levels of schooling. This effect might in fact be larger among SCs compared to the non-SC population because of the overall fewer opportunities SCs have to further their socio-economic standing in society. This suggests that the marginal benefit of education among SCs is most likely higher than that among the non-SCs. Given this reality, seeing an SC member in a state legislator might persuade SC parents to send their children to school with a hope that their kids might be able to reach the same level of influence.

Moreover, previous literature (mentioned above) has shown that state legislators have a significant influence on basic issues such as education in rural India. For instance, elected legislators have the influential duty of determining state's annual budgeted expenditures on education. Education in India is funded by local, state, and central governments. Expenditures on education from the central government steadily rose from 0.64% of GDP in 1951 to 2.31% in 1970 and peaked at 4.26% in 2000. However, by 2004, education expenditures had dropped to 3.49% of GDP. Similarly, there have been variations among the state's education expenditures

as well. The state's education budget in West Bengal, for example, was 16.3% of its annual budget in 2015, which is up by 0.2% from the previous year. It is not unreasonable to expect that the SC legislators are more likely to fight for more education funding for the SCs than the non-SC legislators. This hypothesis has been backed by Down's theory (discussed above), where he argues that an elected official is more likely to cater to the agenda of the group he/she represents. Given that elected legislators determine budget, I hypothesize that education expenditures act as a mediator between the assigned SC quota and the SC enrollment numbers. Therefore, we expect that education budget will have a positive sign in our analysis.

In addition, each state in India also sets aside a certain proportion of its annual budget towards "SC/ST/OBC Welfare." Although the use of these resources depends on the states, historically, the welfare resources have been allocated towards the betterment of education and employment of the SC, ST and OBC groups- in addition to urban development and socio economic justice issues. For example, the Government of West Bengal describes the aim of its Backward Classes Welfare Department as "to improve the quality of life of these socially and economically backward classes and enhancement of capability of the people belonging to these communities, so that they can be very much part & parcel of the mainstream of the Society." Several states have created scholarship programs for higher education of SC students using the welfare funds. Part of the funds is also allocated to fight crimes and atrocities directed towards those minority groups. This allocation of state resource is in the discretion of the elected state legislators. . Therefore, I expect that a SC legislator can influence the amount of funding towards the "SC/ST/OBC Welfare" to a larger extent than a non-SC legislator would. I hypothesize that the welfare budget acts as another mediator between the assigned SC quota

and the SC enrollment rates. Hence, one of the control variables used in my paper is the per capita welfare budget for the SC (WBPC). WBPC is expected to have a positive impact on the enrollment of the SC because the welfare budget is intended to expand resources for the SC population. Similarly, other mechanisms through which an SC legislator can influence SC enrollment might include the number of schools for primary, upper primary, secondary, and higher secondary levels and the student-teacher ratios in those levels of schooling.

To test whether a state's education expenditures, number of schools and welfare expenditures are influenced by the SC political quotas, I use the following model:

$$Y_{st} = \alpha_s + \beta_t + \gamma Quota_{st} + e_{st}$$

wherein s=state, t=year, Y_{st} is the state's education expenditure, number of schools, and welfare expenditure; Quota is the QS variable described above; α_s and β_t are state and time fixed effects, while e_{st} is the state level error term. The primary coefficient of interest is γ , which will be the estimated impact of the quota on the dependent variables mentioned above. The state fixed effects control for any unobserved time-invariant state characteristics on the dependent variables while the time fixed effects control for any policies that affect all states uniformly. Since my omitted variables could be time-invariant or state-invariant, I use the state and time fixed effect regression to estimate my coefficient of interest. A Haussmann specification test concludes that the fixed effect method is more appropriate for this dataset than the random effect model. The results of the model indicate that a state's education expenditures and the total number of schools are influenced by the SC quotas. The results of the regression are discussed on the results section and the tables are in Appendix A. Likewise, to further test those mediator variables, and the causal effects, I have used the Structural Equation Model (SEM), which was first outlined by Sewall Wright (Wright, 1921). The details of the method and results are outlined in Appendix C.

The confirmatory evidence about the mediator variable provided can be used to form a dynamic linear panel model which will estimate a causal effect of the quotas on the enrollment rates. The functional form can be represented by:

$$Y_{st} = \alpha + c_s + \beta t + \gamma Quota_{st} + \theta X_{st} + \varepsilon_{st}$$
(1)

where, X_{st} is a vector of variables which includes a state's expenditures on education per SC pupil, number of schools for various levels of schooling, election year dummy, overall literacy rate, SC population density, student-teacher ratio in all levels of schools, the welfare budget allocated for SC, and the SC income per capita. The education expenditure, welfare expenditure, and the number of schools are included since these variables represent the infrastructure that is essential in improving the enrollment in schools, according to a report from Innovations for Poverty Action (IPA) published in 2012³⁹. The election year dummy is included to make sure that the quota variable does not pick up any election year effects. This allows me to study the impact of the quota share on the enrollment rates not only for the current year, but also for the next four years — until the next election is held. The SC population density variable is included to control for any effect of migration of SC from one state to another between years. Population density is used as a control instead of the total SC population because the area of some states has changed over the years within my dataset. The Y_{st} is my education outcome variable in state

³⁹ The IPA report of 2012 also indicates that other factors that improves school enrollment are: improvements in the labor market, reduction in child labor, reduction in poverty, and reduction in school fees.

s and time t. The state fixed effects (c_s) control for any unobserved time-invariant state characteristics on education outcomes. The Quota variable is the QS as defined earlier. The primary coefficient of interest is γ , which will be the estimated impact of the quota on the education outcomes of the minority group. Also, β estimates the time trend and c denotes the individual state fixed effects.

As with any panel data, there may exist omitted variables that could be fixed for states over time (time invariant) and that could affect the dependent variable (GERs). This creates the linear least squares estimates to be bias. If $cov(Quota_{st}, \alpha_s) \neq 0$, and suggests that the unobserved variables might be correlated with the independent variables, it would then create an upward or downward bias depending on the direction of the correlation. It can be expressed as follows.

$$plim \,\hat{\gamma}_{OLS} = \gamma + \frac{cov(Quota_{st}, \alpha_s)}{\sigma_{Quota}^2}$$

We can deal with this unobserved heterogeneity by taking the first differences or by applying the demeaning transformation (fixed effects model). Equation 1 can also be modeled by the linear least square estimates of the fixed effect model, which are least squares using deviations from the individual means. Likewise, the first difference estimation can be used (equation 2) which removes the constant term and the state effects. This equation can be modelled by the pooled OLS estimator.

$$y_{st} - y_{st-1} = \beta + \gamma \Delta Quota_{st} + \theta \Delta X_{st} + \Delta \varepsilon_{st}$$
(2)

Both the FE (equation 1) and FD (equation 2) methods explicitly model the impact of Quota on GER where γ represents the coefficient of the impact. However, for us to estimate the model using the first difference or the fixed effect approach, strict exogeneity must hold. In other

words, the independent variables in each year in the model have to be uncorrelated with the idiosyncratic errors in each year. This assumption fails here because the Quota variable might not be strictly exogenous because of the way the QS variable is constructed⁴⁰. While we expect higher QS to result in a higher GER, the GER variable might also impact the QS. Similarly, the per capita SC income, Election Year dummy and SC population density could all affect the QS. The pairwise correlation among those variables indicates a fairly strong correlation. For example, the correlation between QS and the welfare expenditure is -0.410, between QS and education expenditure is -0.240, and between QS and number of higher secondary schools is 0.241. [Refer to table 1.05 on page 99 for the complete table]

More importantly, the enrollment of SC students in each year could be correlated with the enrollment in the previous year. This is because children who are enrolled in a grade in a given year are most likely going to be enrolled in the next grade the next year. For example, if most students enrolled in grade four advance to grade five in the following year, then grade five enrollment will be correlated with grade four enrollment for the preceding year. Since the primary level includes grades 1-5, there will be a strong correlation between enrollments in subsequent years. To account for this correlation, I have included the lagged dependent variable as one of my controls in equation 3, as follows. This equation is the level equation.

$$Y_{st} = \alpha + c_s + \beta t + \rho y_{st-1} + \gamma Quota_{st} + \theta X_{st} + \varepsilon_{st}$$
(3)

However, the fixed effect model containing a lagged dependent variable creates an endogeneity problem aside from the one mentioned above. First, the demeaning process which subtracts a state's mean value of the dependent variable and the mean of each independent

⁴⁰ QS is equal to the ratio of SC seat proportion to the SC population proportion.

variable from the respective variables creates a correlation between the regressor and the error term. This would result in estimates of the lagged dependent variable and the variable of interest, QS, to be biased. Therefore, using a fixed effect would still be inconsistent. Likewise, we can construct the first difference model as follows:

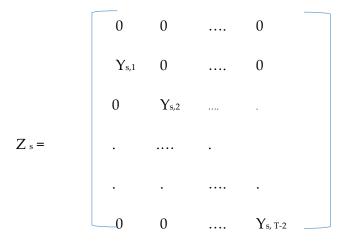
$$\Delta y_{st} = \beta + \rho \Delta y_{st-1} + \gamma \Delta Quota_{st} + \theta \Delta X_{st} + \Delta \varepsilon_{st}$$
⁽⁴⁾

The first difference model still suffers from an endogeneity problem because of the dynamic structure of equation (4) where Δy_{st-1} is correlated with $\Delta \varepsilon_{st}$. Moreover, $\Delta Quota_{st}$ might still be endogenous. Therefore, the estimates that we get for γ using the first difference method of equation 4 will be inconsistent and biased. The solution involves getting rid of the endogeneity generated from both Δy_{st-1} and $\Delta Quota_{st}$. One of the methods to achieve this involves using Δy_{st-2} or y_{st-2} as instruments for Δy_{st-1} . This method was proposed by Anderson-Hsiao (1981). By construction, these lags of the dependent variable will be highly correlated with lags of that variable but uncorrelated with the error term ($\Delta \varepsilon_{st}$) and hence can be used as instruments. Because of this, three or more time period lagged levels of y, can be used as instruments (Holtz-Eakin et al., 1988). Moreover, since the panel used in the paper includes more than three time periods, I will have more instruments than the parameters. This approach, however, has a weakness because of the fairly unbalanced panel that is used in this paper. The first-differencing creates larger gap in the transformed panel which causes a further loss of observation. For example, if one observation y_{st} is missing for the dataset, then both Δy_{st} and $\Delta y_{s,t+1}$ will be missing from the transformed data- thus, magnifying the data loss. One of the solutions to this problem is to use an interpolation to fill the missing data points. However, since the missing observations are stochastic and some observations are missing across state

and year, the ideal interpolation method is hard to estimate without losing efficiency. More importantly, since the dataset possess no sample selection bias, interpolation might not be necessary as long as I have addressed the issue of data loss from the first-difference transformation. As proposed by Arellano and Bover in 1995, using "forward orthogonal deviation" instead of the first-differencing of the observations would prevent loss of data from the differencing transformation. This approach solves the problem of data loss because the approach subtracts the average of all future observations of a variable from the previous observations instead of subtracting the contemporaneous value. In each of the DPD regressions, I have used the orthogonal transformations instead of the first-differences.

However, the instrumental variables approach noted above will likely create autocorrelation. Moreover, it does not use all of the information available in the data sample. Arellano and Bond (AB) (1991) proposed an estimator which uses the lags of all predetermined and strictly exogenous Xs as instruments in addition to the sets of instruments suggested by Anderson-Hsiao. AB suggests using Y_{st-3} as an instrument for Y_{st-1} . In this paper, I have treated Quota as one of the predetermined variables because of its construction and used the lags of it to generate instruments. Arellano and Bond (AB) (1991) used a generalized method of moment's approach to construct a valid set of instruments. AB suggests first taking the firstdifference of the data which removes the time-invariant fixed effect (equation 4). Second, use the lagged dependent variables' values (levels) as instruments for the first-differenced variable, while making sure that the residuals do not have second-order serial correlation. An issue, however, is that if the lagged levels are serially correlated, it provides very little information about the first-differenced variable (Arellano and Bover (1995); Blundell and Bond (BB) (1998)). Therefore, Blundell and Bond (1998) suggested using "system" GMM estimator. In addition to the first-differencing used by AB, Blundell and Bond use the lagged first differences as instruments for equations in levels (non-transformed equations.). Hence, the estimation includes lagged levels as well as lagged difference equations. The two advantages of the system GMM as opposed to the difference GMM are: 1) it preserves the time-invariant regressors, and 2) it provides a test for autocorrelation in differenced residual terms. The former is particularly important because states that tend to elect more minorities in state legislatures, likely differ from other states in ways that affect enrollment and dropout rates as well. For example, if the SC population is higher in one state, the corresponding QS variable is higher, which might translate to a higher enrollment of SC students in that state's schools. Also, the states might be more progressive which, typically, means more minorities are elected to government, and resultantly, there may be higher SC enrollment rates in those states. The progressiveness of states is, in this way, an important factor because SC representatives are capable of running for and being elected to office even when districts are not SC reserved.

With the Anderson-Hsiao approach, the first row of the instrument matrix corresponds to t = 2 because the first observation is lost while applying the first difference transformation. If we include the third lag (y_{t-3}) as a second instrument, we would lose another observation per panel. To avoid the loss of observation, we can construct the instrument matrix from the second lag of **y** as follows where there is one instrument associated with each time period. The following instrument matrix prevents the loss of degree of freedom because of the inclusion of zeros and also because it allows all observations t=2 onward to be included in the regression.



However, all of the available lags of the original untransformed variables can still be used as instruments. If we do so, numbers of instruments available for each time period are different. Hence, as we construct the instruments for higher time periods for each state's time series, additional orthogonality conditions are available. Using these additional orthogonality conditions improves the efficiency of the Arellano-Bond estimator. Similarly, for endogenous variables, lag 2 or higher can be used. For predetermined variables which might not be strictly exogenous, lag 1 could be used, as its value is only correlated with errors dated t – 2 or earlier. The resulting instruments matrix is illustrated below:

However, if all of the available instruments are used, the number of instruments produced will be large because the number of instruments is proportional to T. Hence, when constructing the instrument matrix, I have restricted the number of instruments used in the regression to be closer to the number of groups (n). In all of the regressions, I have selected lags between five and ten of the dependent variable such that I do not have more instruments than the number of states. In other words, I use Y_{st-5} , Y_{st-6} ,..... Y_{st-10} , of the dependent variable such that the number of instruments used is linear in T. Choosing a closer lag would likely create an autocorrelation and hence might have significant AR2 process. If the selected lags are further out $(Y_{st-11}, Y_{st-12} \dots)$, they are less likely to be correlated with the current error term. However, the correlation between the current and the previous observations might be weaker. Hence, my selection of the lag depth is justified. I test the validity of the instruments using the Sargan test and Hansen J test of overidentifying restrictions. As the DPD estimators are the instrumental variable method, it is important to evaluate the Sargan-Hansen diagnostic tests of over identifying restrictions. As the system GMM estimator involves a set of additional restrictions, the Hansens J test for over identifying restrictions is conducted. The regression statistics and the p-values are included in the results table. The results indicate that the Hansen J has a p-value of 0.954 which fails to reject the null hypothesis that the instruments are uncorrelated with the error term (i.e., the instruments are valid), and that the instruments that were excluded were correctly excluded from the estimated equation (J = 0). Similarly, Table 1.08 shows that the Hansen J has a p-value of 0.568, which fails to reject the null hypothesis of J= 0 as well. Hence the over identification restrictions in my DPD model are valid. However, in a model containing large sets of excluded instruments, Hansen J statistics might not be sufficient.

This is because the Hansen J statistic evaluates the entire set of overidentifying restrictions. Hence, Hansen C statistics are also included in the results. The Hansen C test is a mechanism for testing a subset of the original set of orthogonality conditions. The test statistic is calculated as the difference between the Sargan statistic for the restricted regression (regression using a smaller set of restrictions) and the Sargan statistic for the unrestricted regression (regression using the full set of overidentifying restrictions). Table 1.07 shows that the p-value for Hansen C is 0.502. Likewise, table 1.08 shows that the p-value is 0.811, which fails to reject the null hypothesis that the specified orthogonality conditions are satisfied. I have also included the Fstatistics from the first stage in the tables which shows that the instruments used in the regressions are strong since the p-values are significant at the 1% level.

As pointed out by Arellano and Bover (1995) and Blundell and Bond (1998), the lagged levels are not the best instruments for first differenced variables. Arellano and Bover proposed two estimators- one step and two-step estimators and showed that two-step is more efficient. Hence, I have included the two step system GMM estimator, which avoids the problems of weak instruments because it uses equations in both the differences (equation 3) and the level (equation 4). One and two-step GMMs differ in the way their weighting matrices are constructed. Whereas a one-step estimator is based on an initial weight matrix, a two-step estimator uses estimates derived from that initial weight matrix to construct a new weight matrix from which re-estimates of the parameters are obtained. In other words, for the one-step GMM, one would perform the GMM regression using a weight matrix W = I and get the estimator, $\hat{\beta}_1$. To construct the weight matrix for the two-step GMM, one would obtain the residuals from the one-step GMM: $\hat{u}_t = u(x_t; \hat{\beta}_1)$ and use the residuals to construct a weighting

matrix $\hat{\Omega}_{\hat{\beta}_1}$. Finally, one would then use the new weighing matrix to estimate the GMM estimator. Hence, the two-step uses the consistent variance co-variance matrix from first step GMM and is more asymptotically efficient than the one-step estimators. The one step system GMM estimator is also equivalent to the 2SLS estimator, while the two-step system GMM is equivalent to the 3SLS estimator because of the use of the consistent variance-covariance matrix from the first step process. However, in the presence of a large number of instruments, the two step standard errors can have a downward bias. The bias in the two step standard errors are corrected by Windmeijer's (2005) finite correction procedure. For this analysis, both one-step and two-step system GMM are conducted and the results are shown in tables 1.07-1.10.

The DPD is most useful when the data contain large individuals (N) and short time periods (T). In the paper the panel data is cross-sectionally dominant with 28 individual units over 21 time periods. If the moment conditions are satisfied, the GMM estimator will be consistent.

One of the important diagnostic tests in the any dynamic panel estimation is the AR test of autocorrelation of the residuals. By design, we expect a statistically significant p-value for the AR1 test because the residuals of the differenced equation should have first order serial correlation. However, if there are no serial correlation in the original error terms, the differenced residuals should not have a statistically significant p-value for the AR2 test. Therefore, for the second lags of the endogenous variables to be a valid of of instruments, we should have a statistically insignificant AR2 statistic. The p-values for the AR1 and AR2 tests are included in the results. All of the diagnostic tests validate the DPD approach of estimation for this dataset.

2.5 RESULTS

Table 1.06 shows the results of the fixed effect regressions that were used to verify the mechanisms through which the political quotas might be able to influence the enrollment rates of the SCs. In all cases, I use the state and time fixed effects and the reported standard errors are the robust standard errors clustered on states. Column 1 shows that a 10 % increase in the previous year's SC quota share would result in a 1.95% increase in the total number of primary schools in the current year. Likewise, column 2 shows that a 10% increase in the previous year's SC quota share would result in a 3.86% increase in the number of upper-primary schools. Columns 3 and 4 shows that there is a statistically significant lagged impact of the quotas on the number of secondary and higher secondary schools. It suggests that a 10% increase in the SC quota share, two years prior, will increase the number of secondary schools by a 7.3% and the number of higher secondary schools by a 19.7%. The impact of the 10% increase of quota share is a 7.82% increase in the state's education expenditure (column 3) but the state's welfare expenditure is not influenced by the political quota (column 4).

Tables 1.07- 1.11 in the appendix show the results of the Blundell and Bond DPD model. I have used the xtabond2 command developed by Roodman (2009). In all cases, the Arellano and Bond test of autocorrelation in the first differenced residuals rejects the null hypothesis that there might be a presence of second-order autocorrelation in the data. Also, all of the reported estimates show that the Hansen's statistic for over identifying restrictions is valid. I have included the AR1, AR2, the Hansen statistic and the Hansen statistics for difference in the tables. The table also includes the number of instruments and observations. All of the results are derived from the one and two step system GMM method and the stated standard errors are the robust standard errors.

Table 1.07 shows the estimates of the DPD for SC primary GER. Columns 1 and 2 provide the estimates using the one step system GMM and columns 3 and 4 provide estimates using the two step system GMM. I find a positive and significant impact of QS on the SC primary GER. The coefficients on the lagged dependent variable are very close to one, which suggests the unit root problem in these results. Therefore, I interpret the corresponding beta estimates with a slight caution. The results from the two-step system GMM (column 3) show that if none of the proposed mediator variables are used as controls, a 10% increase in QS will likely produce a 3.84% increase in the gross enrollment rates of SC in primary schools. However, if the mediator variables are included as controls, column 4 shows that a 10% increase in QS will likely increase the SC primary gross enrollment rate by 5.55%. This result is significant at the 1% level. The estimates from the one-step system GMM (columns 1 and 2) yield a similar result. Column 1 shows that if none of the proposed mediator variables are used as controls, a 10% increase in QS will likely produce a 3.87% increase in the gross enrollment rates of SC in primary education. However, if the mediator variables are included as controls, column 2 shows that a 10% increase in QS will likely increase the SC primary gross enrollment rate by 3.57%. This result is significant at the 5% level.

The estimates of the DPD for the SC upper primary GER are depicted in table 1.08. Columns 1 and 2 show the estimates using the one step system GMM and columns 3 and 4 show the estimates using the two step system GMM. I find a positive and significant impact of QS on the SC upper primary GER. The results from the two-step system GMM (column 3) show

that if none of the proposed mediator variables are used as controls, a 10% increase in QS will likely produce a 3.98% increase in the gross enrollment rates of SC in upper primary schools. However, if those variables are included as controls, column 4 shows that a 10% increase in QS will likely increase the SC upper primary gross enrollment rate by 7.64%. This result is significant at the 10% level. Likewise, column 2 in table 1.08 shows that with all the mediator variables included, a 10% increase in QS will likely produce a 2.75% increase in gross enrollment rates of SC in upper primary schools. This result is statistically significant at 10% level.

Table 1.09 shows the estimates of the DPD for SC secondary GER. Columns 1 and 2 show the estimates using the one step system GMM and column 3 and 4 show the estimates using the two step system GMM. I find a positive and significant impact of QS on the SC secondary GER. The results from the two-step system GMM (column 3) indicate that if none of the proposed mediator variables are used as controls, a 10% increase in QS will likely produce a 12.03% increase in the gross enrollment rates of SC in secondary schools. This result is significant at the 1% level. However, if the mediator variables are included as controls, the result is positive but not statistically significant. The estimates from the one-step system GMM (columns 1 and 2) provide a similar result with a slightly lower magnitude. Column 1 shows that if none of the proposed mediator variables are used as controls, a 10% increase in QS will likely produce a 12.18% increase in the gross enrollment rates of SC in secondary education. However, if the mediator variables are included as controls, column 2 shows that a 10% increase in QS will likely increase the SC secondary gross enrollment rate by 9.43%. This result is significant at the 10% level.

Table 1.10 shows the estimates of the DPD on SC higher secondary GER. Columns 1 and 2 show the estimates using the one step system GMM and columns 3 and 4 shows the estimates using the two step system GMM. I find a positive and significant impact of QS on the SC higher secondary GER. The results from the two-step system GMM (column 3) indicate that if none of the proposed mediator variables are used as controls, a 10% increase in QS will likely produce a 16.51% increase in the gross enrollment rates of SC in higher secondary schools. This result is significant at 1% level. However, if the mediator variables are included as controls, a 10% increase in QS will likely produce a 3.62% increase in the gross enrollment rates of SC in higher secondary schools (column 4.) The estimates from the one-step system GMM (columns 1 and 2) provide a similar result. Column 1 shows that if none of the proposed mediator variables are used as controls, a 10% increase in QS will likely produce a 8.19% increase in the gross enrollment rates of SC in higher secondary education. However, if the mediator variables are included as controls, column 2 shows that a 10% increase in QS will likely increase the SC higher secondary gross enrollment rate by 10.23%. This result is significant at the 10% level.

Finally, table 1.11 in the appendix shows the results for the drop-out rates of SC students at various levels of schooling. Column 2 shows that a 1% increase in the quota share of SCs would likely decrease the SC primary level dropout rate by 4.1%. Similarly, the dropout rates would likely decrease by 5.2% and 1.9% for upper primary and secondary levels of schooling respectively. These results are all statistically significant at 10% level.

These results show that there is a significant impact of the SC political quotas on the enrollment of SC children in various levels of schooling. For example, an increase of GER in primary level of schooling by 5.55% for every 10% increase in QS means that the mean GER

increases by 5.87. This translates into an increase in enrollment of SC children in primary school by 39,448 students. Likewise, the results show that a 10% increase in QS would result in an increase in enrollment of 20696, 27624, and 28264 of the SC children in upper-primary, secondary, and higher secondary level of schooling respectively. The sample box-plot and the corresponding regression coefficients with 95% confidence intervals are given in figure 1.3 and 1.4 in the appendix.

2.5.1 ROBUSTNESS CHECK

I have interpreted the coefficients described in the results as being the result of a causal relationship between the quota and the enrollment rates of SC students in various levels of schooling. This interpretation comes from the way that the quota variable (QS) is designed, the within-state and cross-time variations and the methods used. In this section, I highlight the three other approaches I have taken to show a similar causal relationship between the quota variable and the enrollment rates of the SC (i.e. to show robustness). I was unable to perform these checks for secondary and higher secondary enrollment rates due to limitations in the data. *2.5.1.1* Dynamic Panel Model with district-level data:

It is also important to verify the results using a narrower data set, especially since the SC quotas are assigned at the district level. Since each state in India is divided into administrative districts, we can use the DPD model used earlier in the paper with the district level data. Due to the limitation in the availability of district level data, however, I will not be using any controls to identify the causal impact. The model identification is identical to the model described in equation 3 and 4 except that the Xs are not controlled for. The DPD model is more suited for the district level data because of the large N and small T in this dataset. The results of this model

are shown in Table 1.12 in the Appendix. The sign and the magnitude of the coefficient of interest are as expected to verify the previously obtained core results. Table 1.12 shows that a 10% increase in the QS will likely result in an increase of 1.06% in the SC primary school enrollment rates. Similarly, a 10% increase in the QS will likely result in an increase of 8.32% in the SC upper primary school enrollment. The results of this district data DPD model indicate that my core results are robust and not impacted by the type of data I use.

2.5.1.2 Dynamic Panel Model with state-level election cycle data:

Another robustness check I have performed is to run the same DPD model outlined earlier, but with the election cycle data instead. For every state, I average the values of each variable between the election years. If my results truly reflect the causal effect of quotas on enrollment, I expect to see results similar to my core results by using this dataset instead. To further test the robustness of the results, I drop the education budget and the welfare budget variables from this test, as these variables had minimal impact on my core results. By taking the average of between-elections values for all the variables, I attempt to eliminate the possibility that election year variabilities are driving the enrollment rates. Hence, I expect to see a smaller coefficient on my variable of interest (QS) compared to the core results. The Hansen J and C statistics are all valid and are listed in table 1.13. The results are summarized in the table as well. The estimates show that a 10% increase in the QS causes the enrollment rates to go up by 12.76% in the primary level and 17.17% in the upper primary level of schooling.

2.5.1.3 Dynamic Panel Model without Union Territories and newer states:

Finally, I check whether my results are driven primarily by the Union Territories and the newer states that were added to the union. To do this, I only use the state-level data for the

states that were created before November, 2000 and exclude all of the Union Territories. I then use the same DPD method (explained above in the empirical section) to get my coefficient estimates. The test statistics and the coefficient estimates are given on table 1.14. The estimates show that a 10% increase in the QS causes the enrollment rates to go up by 27.27% in the primary level and 22.74% in the upper primary level of schooling. These results show that the estimates I obtained are robust to the changes of states and union territories.

2.6 CONCLUSION AND DISCUSSION

Although various forms of Affirmative Action policies have been in place in many countries in the world, the true socio-economic impact of those policies on the targeted populations has not been definitively assessed. This paper fills the void in the existing literature by finding conclusive evidence of the impact of one such AA policy in India on India's minority SC population. By assessing the impact of India's political reservation system on its minorities' education outcomes, I have contributed to the larger discussion on the effectiveness of affirmative action policies in general. If the goal of the political reservation system in India is to uplift the socio-economic status of its minorities, and eventually erode away the underlying social and economic discrimination that still exists based on the caste system, then the reservation system in India should be considered a "success. "

This paper finds that the reservation system in India has been fairly effective in improving the education outcomes of the minorities. One of the major findings of this paper is that the SC quota has a positive and statistically significant impact on the primary, upper primary, secondary, and higher secondary enrollments of SC students. However, the impact is highest at the upper primary level. This might be explained by the role-model effect I discussed

earlier. If a SC parent sees that a SC member from his/her community has been elected as a State legislator, he/she is more likely to be encouraged to send their children to a higher level of schooling. Hence, the effect on upper primary level is higher compared to the primary level of schooling. We can also notice that the impact on secondary and higher secondary SC enrollments is not as big as it is for upper primary enrollment. One of the reasons for this might be the fact that free and compulsory education in India only applies to the primary and upper primary levels of schooling. Hence, even if the SC parents would be willing to send their children to higher education, cost of education might be impeding on the enrollment in those levels.

The gross enrollment ratios (GERs) in various levels of schooling would increase 1.61% -7.64% for every 10% increase in the quota share of the SC population in a state's Legislative Assemblies. Likewise, the dropout rates among the SC students would fall 1.88% - 5.25% for every 1% increase in the Quota Share. Overall, that's a very promising indicator that the quota system is effective in improving the education outcomes of the targeted groups. This means that in some states, even an increase of one seat for the SC would mean increased school enrollments and decreased dropouts.

These results are somewhat consistent with those of previous literature. Most of the previous papers have found little or no effect of SC quotas on poverty or other socio economic outcomes. However, Chin and Prakash (2011) found that the political reservation for SCs did not have any impact on poverty reduction of the states but Pande (2003) found that there is evidence of transfers of resources from high caste to low caste population. My results are different compared to that of Chin & Prakash since they found no impact of the SC political

reservation on poverty in India. However, if we believe that education helps mitigate poverty in the long run, the findings of my paper is crucial in understanding the long-term effect of political reservation on poverty. Given that over the past several years the Government of India has been investing in the education and welfare of SCs, and that Pande (2003) found evidence of transfer between the castes, I did expect to see a positive and significant impact of SC reservation on SC education outcomes. Hence, the results of this paper are significant as I have shown that the SC quota does have a fairly significant impact on the enrollment of SC students in all levels of schooling. My second contribution to the literature is the establishment of the mechanisms through which an elected SC legislator via political reservation might have an impact on the education outcomes of SCs. First, the use of the fixed effect method validates the mechanisms through which the political set-aside have an impact on the education outcomes of the SC minority. To further strengthen the argument, I used the SEM mediation analysis (refer to appendix C), which also shows that an elected SC legislator have a significant influence on the allocation of state's resources towards education, welfare, and the number of schools built annually.

Given the findings of previous studies and this paper, it is clear that political reservation is not just pro-poor but also pro-minority. If the SC reservation is able to drive up SC enrollment in schools, it contributes towards India's twelfth five year plan and it's Sustainable Development Goals (SDGs) of achieving universal primary education. A more robust and aggressive effort from all branches of governments (local, state and central) to complement the constitutionally granted political reservation, could potentially integrate the SC population into mainstream India in all socio-economics sectors of life

CHAPTER 3: HEALTH

3.1 INTRODUCTION

The health outcomes of people in most of developing countries are still poor compared to those in developed counties despite the huge amount of received foreign aid, modern medical advances, and the many NGOs working in the health sector in developing countries. Two health outcomes that remain especially poor in developing countries are infant mortality rate (IMR) and under-five mortality rate (U5MR)⁴¹. According to the UN World Population Prospects Report of 2004, 83% (5.3 billion) of the world's population lives in developing countries. According to the same UN report, India is home to 21% of the developing world's population, and 121 out of every 1000 Indian children will die before the age of 5.⁴²

In many developing countries, minorities usually suffer from socio-economic disparities, which affect their economic mobility. In India, the Scheduled Caste (SC) population represents 16.6% (206 million) of the total Indian population, thereby forming the largest minority group. Historically, caste affiliation and social groups have had substantial implications among various aspects of life in India, including educational outcomes, economic well-being, and career perspectives. However, with affirmative action and anti-discriminatory laws, access to resources for SC and other minorities has improved. Nevertheless, studies show that the SC

⁴¹ Infant mortality rate is defined as the ratio of number of deaths within the first year of life to every 1000 live births. Under-five mortality rate is defined as the ratio of number of deaths within the first 5 years of life to every 1000 live births.

⁴² Hence, India's under-five mortality rate is 121..

population continues to suffer from economic disparity (Deshpande 2001) which impedes socioeconomic mobility. Various factors including economic well-being, health infrastructures, and education, contribute to the prevalence of high IMR and U5MR among the people in developing countries. Moreover, various papers have shown that the IMR and U5MR are significantly different among various social groups. In particular, research has shown that IMR and U5MR are especially high among the minorities in the developing countries⁴³. This disparity is noted by Das, Bhavsar, and Patel (2000) in 1993. They find that in the Indian state of Gujarat, U5MR varied from 98 among the highest caste's population to 127 among the lower caste (including scheduled caste) population. Lower caste, indicative of a minority social group, could thus be clearly linked to higher U5MR.

Health outcomes, as with many socio-economic outcomes, of India's SC population, have been poor compared to the non-SC population. In a caste-dominated country such as India, SC makes up more than a sixth of the population of India (about 160 million). Even after the passage of numerous anti-discriminatory laws since 1989, SC population has seen a continuous violation of its basic human rights for decades. According to a 2007 report by the United Nations (UN), "[m]ore than 165 million people in India continue to be subject to discrimination, exploitation and violence simply because of their caste." The same report goes on to say that "caste-based divisions continue to dominate in housing, marriage, employment and general social interaction—divisions that are reinforced through economic boycotts and physical violence." It has been documented that these minority groups experience structural

⁴³ Das, N.P., Bhavsar, S. & Patel, R. (2000, November). The growth and development of scheduled caste and scheduled tribe population in Gujarat and future prospects. International Union for the Scientific Study of Population

discrimination in various forms in the society. Hanchinamani (2001) noted that these groups experience physical, psychological, emotional, and cultural abuses in their daily lives. These actions are in-turn legitimized by the society at large.

Various papers and reports have documented that there exists a physical segregation of living spaces in India, most commonly in the rural areas, forcing poor minority groups⁴⁴ to live in some of most unhygienic and inhabitable conditions⁴⁵. This could affect their access to healthcare and quality of health service received which impacts their overall health status and outcomes. According to a report published by the Center for Enquiry into Health and Allied Themes (CEHAT)⁴⁶ in 2007, along with poor healthcare access, some groups among the SC population are especially vulnerable to emotional distress and common diseases. In the presence of continued structural and social discrimination, obtaining equal access to health services might be an impossibility. The report also highlights the possibility that the social views of healthcare professionals in rural areas might impede access to healthcare for the SC population. The doctor's attitudes towards the social structure of the caste system might also act as a barrier for the SC to receiving quality healthcare from the already available healthcare system if SC members suspect they will face discrimination.

Although data and studies on the nature of discrimination faced by SC in the health sector are very limited, there is indirect evidence of discrimination reflected in some of the health indicators among the SC population. Some of the most pressing health concerns for the

⁴⁴ The mentioned reports suggests that these rural dwellers are mostly Scheduled Caste population and religious minorities.

⁴⁵ Minorityrights.org; Development of Scheduled Castes and Scheduled Tribes in India, 2008; National Campaign on Dalit Human Rights (NCDHR) 2008.

⁴⁶ Chatterjee, C., & Sheoran, G. (2007). Vulnerable groups in India The Centre for Enquiry into Health and Allied Themes (CEHAT), 1-21.

SC are various measures of child mortality. According to the NFHS-3 survey conducted in 2005-2006, 66 out of 1000 SC babies die before their first birthday, compared to 57 out of 1000 babies among the all population of India. Similarly, for every 1000 new births, 88 SC children died before their fifth birthday compared to 74 for all population of India. Likewise, other indicators, including NMR⁴⁷, PMR⁴⁸ and CMR⁴⁹ are significantly higher for the SC children compared to the all India population. Table below, depicts statistically significant SC versus ST and all India comparison of the various indicators related to children mortality based on the NFHS-3 2005-2006 surveys.

Social	Neo-natal	Perinatal	Infant	Childhood	Under 5	Childhood
Group	mortality	mortality	mortality	mortality	mortality	vaccination
	(NMR)	(PMR)	(IMR)	(CMR)	(U5MR)	(full
						immunization)
						(%)
SC	46.3	55.0	66.4	23.2	88.1	39.7
ST	39.9	40.6	62.1	35.8	95.7	31.3
All India	39.0	48.5	57.0	18.4	74.3	43.5

However, it is also important to understand that the health outcomes are dependent on the type and extent of health infrastructures available to the SC population. Although it is difficult to quantify the number of health centers or doctors working exclusively for the SC population, examining the health infrastructure in the rural areas of India, where 80% of SC live, provides an indication of the resources available to them. According to the RCSH survey of

⁴⁷ Neo-natal mortality rate (NMR) is the number of deaths among the children less than 28 days old, out of every 1000 live births.

⁴⁸ Perinatal mortality rate (also referred to as Pre neonatal mortality rate) is the number of fetal deaths (stillbirth), out of every 1000 live births.

⁴⁹ Childhood mortality rate (CMR) is the number of deaths among children aged 1-4, out of every 1000 live births.

2004, only 68.3% of the rural villages in India had at least one Primary Health Center (PHC). Similarly, about 79% of the villages had a government hospital. According to the same survey, close to a hundred percent of urban areas have PHCs and government hospitals. However, many of those facilities in rural areas are not accessible throughout the year because of poor road conditions- a third of the villages are affected. According to a different survey (NSSO 2004), the total medical expenditure per treated patient in rural India was INRs. 257 while that in urban India was INRs. 306. Of that amount, people in rural India spent just INRs. 26 per capita in private medical facilities, compared to INRs. 299 in urban India. These figures indicate that minorities living in rural India have much less infrastructure to help them improve their health outcomes.

The goal of this paper is to assess, through an indirect effect, the impact of the political reservation on the IMR and U5MR of the SC population. Since some of the previous literature has shown that a minority representation in local, state, and federal branches of governments have influence in policy matters, we should expect similar influence in the health sector. Increased representation of a specific minority caste in an executive function in a governmental sector might improve the provision of public goods and health infrastructures for members of that caste. Some of the ways that a minority representative might have an impact on the IMR and the U5MR of its people is through local and state health expenditures, the number of PMCs and governmental health facilities and other health infrastructures. Additionally, a direct effect may be present, for example in the case of a "role model" effect, wherein the successful election of an SC member inspires SC parents to imagine greater potential for their children and invest in them accordingly. Hence, both indirectly and directly, an SC representative has the potential

to positively influence the IMR and U5MR for the SC population. I study the impact of such representation on IMR and U5MR based on the DHS health survey of 2005-2006 using the Cox Proportional Hazard Model.

3.2 LITERATURE REVIEW

Many papers have highlighted the determinants of childhood mortality in developing countries. However, very few have studied the impact of affirmative action policies on childhood mortality. Van Den Berg, Bergemann, and Sumedha (2010) created a proportional hazard model using data from the National Family Health Survey (NFHS 3) to identify the impact of employment quotas in India on the childhood mortality rates. The employment quotas in India went into effect in 1950 for the SC and ST population and in 1993 for the OBC population, reserving a sizable portion (just under 50%) of the public sector jobs for these castes. The authors control for various biological, environmental, and socio-economic factors, as well as for some of the macroeconomic indicators of specific Indian states. They find that the employment quotas are able to decrease childhood mortality among the SC children. They also find that this decrease is substantially greater in India's rural areas as opposed to its urban areas. This paper conducts an analysis similar to Van Den Berg, Bergemann, and Sumedha (2010), except examining the effect of the political reservation instead of the employment quotas. As was the case with employment quotas, I expect political quotas to have a negative and significant impact on the mortality of SC children.

The determinants of childhood mortality have been debated in the medical science field as well as in the social science sector. However, there is a consensus on some of the biological, environmental, and socio-economic factors that can drive mortality in children. Sonia Bhalotra

has published several papers studying the impact of birth spacing (Bhalotra and van Soest, 2007), economic growth (Bhalotra, 2008), state health expenditure (Bhalotra, 2007) and religion (Bhalotra, Valente, & van Soest, 2008) on childhood mortality. In all cases, Bhalotra and her coauthors use a dynamic panel data model to study the impact of these determinants, ultimately concluding that longer birth spacing, economic growth, and higher state health expenditure each reduce childhood mortality. They also find that religion plays a role, wherein a Muslim child in India faces a higher risk of childhood mortality than a Hindu child.

Masset and White (2003) also examined several determinants of mortality in children. Restricting their analysis to the Indian state of Andhra Pradesh, they used a Cox proportional hazard model to identify specific determinants of childhood mortality. They then compared the hazard ratios and the coefficients to those of the state of Kerela. Their analysis indicated that key determinants, including birth spacing, a mother's health, and environmental factors, act as driving causes of childhood mortality. As discussed by Grossman (1972) in his book <u>The</u> <u>Demand for Health: a theoretical and empirical investigation</u>, a household's income drives its demand for healthcare. However, health has negative income elasticity. Grossman's explanation for this is that "the income elasticities of the detrimental inputs in the health production function exceed those of the beneficial inputs". In order to test the income as one of the determinants of childhood mortality, I use the "wealth index" as proposed by the Demographic and Health Surveys (DHS), and as a necessity, because DHS does not collect data on household income.

The mechanism through which an elected minority representative might have an impact on childhood mortality is important to identify in this analysis. As highlighted in section 2.2 on

chapter two, the established mechanism is the type of policy and budget influence that an elected official might have. For example, Downs (1957) concluded that an elected official might have a bias towards the agenda of the group he/she represents. Likewise, Osborne and Slivinski's (1996) Citizen-Candidate model revealed that the identity of the candidate running for public office has an impact on policy determination. This contributes to my understanding that a minority legislator could influence certain policy measures affecting the associated minority population, and that these measures, in turn, could have implications for that population's health outcomes. As an example, the elected minority candidate could be involved in determining the amount of a state's annual health and welfare budget, allocated to the represented minorities, potentially improving the well-being of those minorities. Pande (2003), Duflo and Chattopadhyay (2004), and Chin and Prakash (2011) have studied the various possible mechanisms and the impact of the political reservation laws on various measures of well-being of the targeted minority groups. Assuming that the elected SC representatives care for the policy preferences of the group he/she is representing, or they care about their reelection prospects, I hypothesize that such policy preferences will improve the health outcomes (namely, IMR and U5MR) of the represented group.

3.3 DATA

The health data used in this paper comes from the National Family Health Surveys (NFHS 3) of 2005. The dependent variables of interest are infant mortality rate (IMR) and underfive mortality rate (U5MR), as measured by the probability of SC children dying within 1 year and 5 years respectively. Outside of the NFHS surveys, data on mortality rates among the SC

population for various states are difficult to find. Although the IMR and U5MR can be estimated from data from the Sample Registration System (SRS) and demographic surveillance systems, these systems are not widely used in developing countries. India, does in fact, release data from a SRS system, however, these data are considered to be less comprehensive, and less reliable, and are of uncertain accuracy. Additionally, the SRS system does not record all socioeconomic information, making an analysis of the socioeconomic inequalities in mortality impossible. Furthermore, information specific to the SC population and their health service provided is inadequate and inconsistent, within the SRS system, because of missing data on many crucial variables. Therefore, the best source of data on childhood mortality for this analysis is a household survey such as the NFHS.

The NFHS-3 dataset provides information on birth and death histories of the respondents and their children. To estimate child mortality rates from NFHS dataset, we can use the fertility histories of the women respondents in the survey. These can be constructed based on responses to questions asked to women of child-bearing age about births and deaths of their children. The comprehensiveness of the NFHS data allows for a compilation of child-level data containing information for each child on the date of the mother's interview; specifically: the date of birth of the child, a binary variable indicating the child's status (alive or dead), and the survival time. Mortality estimation is based on deaths over a specific period of time—usually the 5 or 10 years before the survey date. The NFHS-3 dataset includes birth histories of 124,385 women (aged 15-49) out of which about 23,125 women were from the SC group. The summary statistics of all variables are listed in table 2.01 in the appendix. Each of those variables are discusses in detail in the next section.

I have also incorporated health infrastructure data into the analysis. The data on number of primary health centers and the number of doctors comes from the Ministry of Health and Family Welfare, which spans the period 1970 to 2005. I have also included the health expenditures by states as one of the controls. These data were collected from indiastat.com and from states' annual budget reports. The socio-economic covariates in the paper are derived from the NFHS surveys. I have used place of residence, age of mother, and medical attention for pregnancy and delivery as the socio-economic determinants of infant and under five mortality rates. I also control for the household's annual income. The following section describes the covariates used in the empirical analysis in more detail.

3.3.1 Covariates and Determinants of Childhood Mortality

The determinants of childhood mortality used in this paper are based on the findings reported in previous literature. The determinants can be consolidated into four different groups — biological, socio-economic, behavioral, and environmental variables — which are described further below and are used in the model described in the next section.

3.3.1.1 Biological Determinants:

The biological determinants of childhood mortality used in this paper are: *birth order of the children, mother's age at birth of the child, length of birth intervals, and gender of the child.* Although the relationship between birth order of a child and mortality is not conclusive, several health scholars expect mortality to be higher for a woman's first child. Sullivan (1994) noted that higher mortality among first born children in developing countries can be attributed to the fact that women in rural parts of developing countries may have their first child before they have

reached reproductive maturity. Additionally, a paper published by the World Health Organization (WHO) in 2000⁵⁰ shows that the IMR of first child in India is 93, while that of second and third child is 77 and 72 respectively. Likewise, U5MR among the first child is 117, while that for second and third child is 106 and 107 respectively⁵¹. In the dataset, 32.95% of the children reported were the first child to the mother. Among the SC respondents, 30.81% of the reported children were the first-born.

According to the same WHO report, IMR and U5MR are much higher among children born to adolescent mothers (aged 10-19) and to older mothers (aged 35 and up). Pebley and Strupp (1987) concluded that the higher mortality rates among children born to adolescent mothers might be due to the fact that these women have not yet reached reproductive maturity. Similarly, the higher mortality rates among children born to older mothers may be a result of natural age-based deterioration of these women's reproductive health. According to the NFHS3 survey, 27% of surveyed women gave birth to a child before they turned 20. IMR among children born to mothers less than 20 years of age is 107 and among the children born to mothers age 40-49 is 112. Likewise, the U5MR is 141 for adolescent mothers and 163 for older mothers. All other mother age-groups have significantly less IMR and U5MR. Given that the relationship between mother's age and childhood mortality is U-shaped, wherein childhood mortality is highest at lower age groups and at higher age groups, I use the mother's age and mother's age squared as controls. In the survey sample that is used in this paper, the mean age of the mother (SC and otherwise) was 23 years.

⁵⁰ Claeson, M., Bos, E., Mawji, T. & Pathmanathan, I. (2000). Reducing child mortality in India in the new millennium. Bulletin of the World Health Organization, 78(10), 1192-1199.

⁵¹ All of the IMR and U5MR numbers are expressed as the number of deaths out of 1000 new births.

One of the most important biological determinants of childhood mortality is the interval between births. Arulampalam & Bhalotra (2004) find that if a child is born in a household after the death of a previous child, the probability of the child dying as a newborn increases. The United Nations International Children's Emergency Fund (UNICEF) determined in a 1999 report on developing countries that "a child born within 24 months of the previous child has a 68% higher risk of dying within the neonatal period and a 99% higher risk of dying in the postneonatal period." Furthermore, the risk significantly increases if the previous child had died. The 2000 WHO report also shows that the IMR among children born within 24 months of the previous birth is 130 and the U5MR is 178. These rates become significantly smaller if the birth interval increases to more than 48 months (IMR of 42 and U5MR of 57). Boerma and Bicego (1993) analyzed the relationship between birth interval and mortality and concluded that the reason the mortality rate is high if the birth interval is low, is that the mother's body is not fully recovered from physical and nutritional depletion if the interval between births is less than 24 months. In the survey sample that is used in this paper, the average birth interval among SC mothers is 32 months. Finally, gender of the child is also an important determinant of childhood mortality. There is a general consensus that male children have a higher mortality probability at all ages of childhood because of genetically determined characteristics. However, in many developing countries, including India, the existence of a cultural preference for a male child over a female, causes me to expect that the IMR and U5MR for female will be higher. Sullivan (1994) and Croll (2001) both noted this relationship. In the sample, 48.04% of the SC respondents were female.

3.3.1.2 Socio-economic Determinants:

The socio-economic determinants of childhood mortality used in this paper are: *the mother's highest education level and her household's economic well-being*. The relationship between a mother's education and her child's probability of dying is well documented⁵². As Sullivan (1994) discovered, children of illiterate mothers have a higher probability of dying at all ages than those of educated mothers. This may be attributed to the fact that, compared to a poorly educated mother, a well-educated mother is likely equipped with a better understanding of relevant environmental threats and the appropriate usage of healthcare methods and facilities. The WHO report (2000) on India supports this finding, indicating that if a mother is illiterate, her children's average IMR and U5MR are 101 and 141 respectively. These mortality rates drop significantly if the mother has a high school or higher education; in this scenario, IMR and U5MR are 37 and 43 respectively. In my dataset, 64% of the SC women did not have any formal education.

Since the NFHS survey data do not contain information on the respondent's income, I use the wealth index calculated based on the ownership of different asset types as a measure of "economic well-being", as recommended by DHS. According to DHS guidelines, a Principal Components Analysis (PCA) is first conducted and the coefficient recorded for each of the assets. Second, the PCA coefficient, mean and the standard deviation for each asset are used to find a wealth index associated with possession of each asset. Finally, the individual scores are added to obtain the wealth index of the households. I then use DHS's guideline to divide those wealth indices into five quintiles — poorest, poorer, middle, richer, and richest for each of the

⁵² Retherford (1989), Palloni & Tienda (1986), Pebley & Strupp (1987), Shiva Kumar (1995), Boerma & Bicego (1993), Bourne & Walker (1991)

households. I expect the "poorest" quintile to have the higher childhood mortality rates and "richest" quintile to have the lowest childhood mortality rates. In the sample, 22% of the SC respondents were classified as "poorest", 22% were "poorer", 20% were "richer", and 14% were "richest".

3.3.1.3 Behavioral Determinants:

The behavioral determinants of childhood mortality used in this paper are: *tetanus immunization, the number of antenatal visits, mother's knowledge of oral rehydration salts (ORS), and the total absence of breast feeding.* According to the UNICEF report published in 1999, 73% of U5 mortality in India is a result of low-cost treatable diseases. Among them, diarrhea, respiratory diseases, tetanus, and measles account for 60% of all deaths of children under-five years of age. In my dataset, only 11% of the SC respondents had used tetanus immunization on their children. The report also notes that 60% of childhood mortality would be preventable by the use of antibiotics, immunization, and oral rehydration methods (ORS). Gupta (1990) found that immunization and mother's knowledge of ORS are two of the key behavioral determinants of IMR and U5MR. Similarly, many scholars have noted that antenatal care is very important in the early health of a child⁵³. Hence, I use number of antenatal visits and the mother's knowledge of ORS as two of the controls in my regression. Only 1.5% of the SC respondents in the dataset had used ORS and 11% had used some antenatal care. Additionally, I

⁵³ Acharya, L., & Cleland, J. (2000). Maternal and child health services in rural Nepal: does access or quality matter more?. Health Policy Plan, 15(2), 223–229.

Addisse, M. (2003, January). Maternal and Child Health Care. University of Gondar. Retrieved from https://www.cartercenter.org/resources/pdfs/health/ephti/library/lecture_notes/ health_science_students/ln_maternal_care_final.pdf.

have chosen to incorporate a dummy variable of whether or not the child was breastfed as a behavioral determinant of childhood mortality. There is a consensus in the medical science literature that as a mother's milk is highly nutritious and rich, it can prevent many illnesses and diseases for a breastfed child. Palloni & Tienda (1986) found empirical evidence that a breastfed child has a lower probability of dying compared to a non-breastfed child. Hence, I expect to find in this analysis that a breastfed child has a lower probability of dying. In the dataset used, 75% of the SC women did not breastfed their child.

3.3.1.4 Environmental Determinants:

The environmental determinants of childhood mortality used in this paper are: *absence of a toilet, absence of an accessible and safe source of potable water, and use of unsafe* (*potentially toxic*) *fuel sources such as dung, wood, and coal for cooking*. As Mosley and Chen (1984) noted in their framework for the study of childhood survival in developing countries, medical science has reached an overwhelming consensus that the environment a child is born into has an enormous impact on the child's probability of dying. A report published by the International Institute for Population Sciences (IIPS) in 2000 observed that diarrhea and acute respiratory infections are two of the major causes of mortality among children under five years of age in India. Water and air quality are known to be major contributors to diarrhea and respiratory infections. In a fact sheet report published in 2006, UNICEF indicates that IMR and U5MR worldwide are considerably lower among children born in households with safe and accessible drinking water and safe and improved toilets. According to the NFHS3 (2005-2006) India survey, 43% of the respondents did not have any toilet facility and 9.6 % of respondents

were using unsanitary toilet facilities. Many scholars have found that the use of safe water and toilet facilities have a decidedly positive impact on childhood mortality. Kishor and Parasuraman (1998) found that in India, households using safe water and toilet facilities had a significantly lower IMR and U5MR. Merick (1985) and Brockerhoff and Derose (1996) found similar results in urban Brazil and five East African countries respectively. In this analysis, I consider absence of a safe water source and absence of a sanitary toilet facility as possible causes of diarrhea, and use these as two environmental controls for childhood mortality. In the sample, over 93% of the SC respondents reported they had safe drinking water source. However, 57% of the SC respondents did not have access to any toilet facility.

As previously mentioned and as noted in the Mosley-Chen Framework, air quality is also an important factor affecting IMR and U5MR rates. In this paper, I use the type of fuel used by households as a determinant of air quality experienced by the child. If the household uses coal, wood, grass, or animal dung as its primary source of cooking fuel, then we expect the immediate air quality in the household is unsafe, in terms of air quality. Thus, I expect to find that household's use of an unsafe fuel will result in higher IMR and U5MR. In the sample, over 74% of the SC respondents used unsafe fuel for cooking.

3.3.1.5 Political and Administrative Determinants:

The political and administrative determinants used in this paper are *quota shares and the number of primary health centers (PHCs)*. Many papers have highlighted the importance of state capacity and a robust political development in the fight against childhood mortality. Dawson (2010) points out that the state's capacity to promote law and assign public goods is an

important factor in childhood mortality. Similarly, Houweling, Kunst, Looman & Mackenbach (2005) indicate that socioeconomic development and a government's commitment to citizen's health are important factors in curbing childhood mortality.

The quota variable used in this paper is a modification to the QS variable described in section 1.2.2 on chapter one. Quota Share (QS) variable represents the ratio of the SC legislative seat proportion to the SC population proportion. In this paper, I have divided the QS variable into five categorical groups and is referred to as QD. They are:

$$QD = \begin{cases} 1, & QS < 0.50 \\ 6, & 0.50 < QS < 0.60 \\ 7, & 0.60 < QS < 0.70 \\ 8, & 0.70 < QS < 0.80 \\ 9, & 0.80 < QS < 0.90 \end{cases}$$

This redefinition of the quota variable changes it from a continuous variable to a categorical variable. QD of 6, for example, means that compared to the SC population proportion, there is a 50-60% representation of SCs in the state's legislative branch. As an example, if QS for the state X in the year Y is 0.60, and if SCs represented all of the population of India, there is a 60% representation in the assembly for a 100% representation in population. More realistically, if the SCs represent about 20% of the Indian population, the QS of 0.60 would translate into a 12% representation of the assembly seats. This construction is more suited for the paper since more information can be obtained on the impact of quotas on childhood mortality if analyzed within several quota ranges instead of a continuous quota. If the variable was a continuous variable (QS), then the hazard ratio given by the hazard model would be the ratio of hazard rates for one unit increase in the QS. However, that conclusion might be inappropriate if there is any critical threshold in quota that might have a more dramatic impact in reducing the child mortality

rates. For example, an increase of QS from 0.49 to 0.50 is a unit increase in the QS and the hazard model would give us the reduction in IMR as a result of that one unit increase in QS. Instead, if categorical variable (QD) was used, then we might be able to capture the marginal effect of crossing a threshold, whereby the same one unit increase in QS from 0.49 to 0.50 might have a more dramatic effect if 0.50 is a critical threshold. Hence, we can look at the results from using QS as the total effect of the quotas while the use of QD would result in the marginal effects- which is what is of interest in this paper. Therefore, the impact is clearer if we can identify a quota range that has the most impact on the reduction of childhood mortality. Thus, I am now able to compare the hazard rates of individuals born in states with varying quota groups.

I expect the critical threshold of representation to be 50-60% (QD = 6). The QD of 6, means that at least 50% of the SC populations are represented in the legislatures. The graph (figure 2.3) shows that if the QD =6, the slope of the line of proportion of children surviving against the survival time is the steepest. In other words, the 50-60% quota marks the most influential quota to decrease childhood mortality. If at least 50% of the SC population is represented, then the SC legislators might have a higher influence over the legislative outcomes, thereby influencing the policies that his/her voters demand the most. Likewise, if the elected SC members were to have a higher chance of re-election, QD of 6 means that the legislator should follow the mandate of the plurality of his/her voters. Therefore, if SC legislator were rational, we would expect him/her to cater to its voters by working on policies designed to improve the health outcomes of his/her voters. Hence, the QD of 6 should have a higher chance of reduction of IMR and U5MR.

3.4 MODEL AND EMPIRICAL STRATEGY

The estimates of the impact of the political quota, along with several determinants of childhood mortality, on IMR and U5MR, can be obtained using a logistic regression model or a hazard model. However, when the mortality rates are analyzed using a logistic regression model, only children who were born five years prior to the NFHS survey (for U5MR) can be included in the study. This is because children born more recently may not yet have been fully exposed to the same environmental and other risks as those faced by children who were born five years earlier. Therefore, I would predict a downward bias in the mortality measurements if children born less than five years ago were included in the study. However, if these observations are discarded, the data suffers some loss of information. More importantly, the exclusion of these observations causes an increase in measurement bias because it assumes that some of the determinants of IMR and U5MR have been applied to the whole life of the child, when in reality, the determinants were obtained at the time of the survey, and the period of exposure is unknown. Using a hazard rate model instead of a standard regression model avoids these pitfalls by treating censored observations differently. Unlike the logistic regression model, hazard rate model analyses the time to an event (death in this case), thereby, being able to account for the censoring. The hazard function is the probability that a person will experience the event (death) in the next instant if he/she survives to time t. The hazard model used in this paper is the Cox Proportional Hazard Model (Cox & Oakes 1984, and Collet 1994).

The model identification comes from the way that the political quotas are assigned, as discussed in section 1.2.2. The quotas are assigned based on the proportion of the SC population in each state. The quotas do not change based on any of the socio-economic or health outcomes

of the minorities; they are strictly based on their population proportion. This exogenous variation in the Quota variable allows us to use the Quota as one of the controls in the Cox Proportional Hazard Model for each state, and hence the impact of quota on childhood mortality can be estimated.

In this study, child mortality is calculated for two age groups- mortality of SC children from birth to the age of 12 months (Infant Mortality Rate) and mortality of SC children from birth to the age of 60 months (Under-5 Mortality Rate). Those variables will be referred to as "IMR" and "U5MR" respectively. Hence, one of the dependent variables, IMR, is the risk of SC children dying within the age interval 0-12 months. Likewise, dependent variable, U5MR, is the risk of SC children dying within the age interval 0-60 months. The independent variables are all of the biological, environmental, behavioral and political variables mentioned in the previous section. Continuous-time hazard rates are estimated using Cox Proportional Hazard Model. I present my results of the survival analysis with and without several of those variables.

In any survival analysis, the calculation of the dependent variable requires setting up an event indicator and a measure of time at risk for the event. In this paper, "event" occurs when the child dies and the measure of time at risk for the event is the age of the child until the "failure" has occurred, in months. The time measure is calculated as follows:

$$survival time = \begin{cases} date \ of \ interview \ (CMC) - date \ of \ birth \ (CMC), \\ age \ at \ death, \end{cases}$$
 child is alive child is dead

NFHS reports all dates in century month code (CMC) in the dataset. CMC is the number of months since the start of 20th century. For example, December 1900 will have a CMC of 12, and January 1901 will have a CMC of 13. Therefore, if the date of interview in CMC was 1279 and

the date of birth of a child in CMC is 1218 and the child is still alive, then at the time of interview, the child is 61 (1279 – 1218) months old. However, since I am interested in infant-mortality (death before 12 months of age), and under-five mortality (death before 60 months of age), I censor the data if the child is over that age limit and still alive. This is important because those individuals do not experience the "event" (death) before the study ends. This is one of the critical reasons that logistic regression is not appropriate in this situation. Refer to figure 2.1 and 2.2 in the appendix for the censor illustration.

Cox (1984) introduced a variety of models that focuses on hazard functions and has been widely used in biostatistics and social sciences. One of the models he suggested is the proportional hazard model, where the hazard for any group of individuals is a fixed proportion of the hazard for any other groups. The generic proportional hazard model, thus, looked like this:

$$h_i(t|x_i) = h_0(t)\exp(\sum \beta_i X_{ij})$$
 2.1

where $h_i(t|x_i)$ is the hazard function at time t for individuals with covariates X. This model could be rewritten as:

$$\ln\left(\frac{h_i(t)}{h_0(t)}\right) = \sum (\beta_j X_{ij})$$
 2.2

The β coefficients can, thus, be calculated by maximizing the following partial log likelihood function.

$$L_{j}(\beta) = P(individual \ j \ fails \ | someone \ from \ R(X_{j}) \ fails)$$
$$L(\widehat{\beta}) = \prod_{j=1}^{k} \frac{exp \sum_{i=1}^{m} \beta_{i} X_{ij}}{\sum_{p \in R} exp \sum_{i=1}^{m} \beta_{i} X_{pi}}$$
2.3

In this paper, let $\lambda(t|B_{ij}, E_{ij}, V_{ij}, S_{ij}, P_{st})$ denote the hazard function for the *i*th child in *j*th household at time t, i = 1, 2, ..., n; where, B_{ij} denotes the biological determinants, E_{ij} denotes the environmental determinants, V_{ij} denotes the behavioral determinants, S_{ij} represents socioeconomic determinants, and P_{st} denotes the political determinants. The baseline hazard function at time t is denoted as $\lambda_0(t)$ such that:

$$\lambda(t|B_{ij}, E_{ij}, V_{ij}, S_{ij}, P_{st}) = \lambda_0(t) \exp(\sum \beta_j X_{ij}); \qquad 2.4$$

where X_s are all the independent variables (B, E, V, S and P) that might cause the hazard (death). The political determinants include the "number of PHCs" variable and the quota variable. As mentioned in earlier section, all of the covariates under B, E, V, and S come from the NFHS-3 survey. The summary statistics of all of the categorical variables are listed on table 2.01. The table provides the number of observations for each of the covariates and their subcategories along with the percent of the dataset they represent. In order to test the collinearity among all of the covariates, I use the chi square test of independence. Under the test, the null hypothesis is that the categorical variables used are independent. The choices of all of the B, E, V, S, and P covariates are based on the test results, which are shown on table 2.04.

The covariates under P are merged in the dataset such that both the "number of PHCs" and the quota correspond to the date of birth of the children in the survey. For example: if a SC child is reported born in 1995 in the state of West Bengal, the number of PHCs and the SC quota in West Bengal in 1995 is merged with the data for that respondent. Because of the way that the political and administrative data are merged in the dataset, both of those covariates are time-invariant. The P variables are assigned for each child in the dataset depending on the year they were born and remains fixed for the remainder of their lives.

The hazard ratio, $\lambda_1(t)/\lambda_0(t)$ is the ratio of the risk of the event (death) occurring at time t for a group to the risk for the baseline group. Hence, the log of the ratio, i.e. the hazard function of the group divided by the baseline hazard function at time t, is a sum of the linear combination of parameters and regressors and is given by:

$$log\left(\frac{\lambda(t|B_{ij},E_{ij},V_{ij},S_{ij},P_{st})}{\lambda_0(t)}\right) = \left(\sum_{j=1}^K \beta_j B_{ij} + \beta_j E_{ij} + \beta_j V_{ij} + \beta_j S_{ij} + \beta_j P_{ij}\right)$$
2.5

The ratio of hazard functions is therefore just a ratio of risk functions. Therefore, the proportional hazards regression model can be interpreted as a function of relative risk among different groups. This is different from the logistic regression models because the logit models are a function of an odds ratio. Changes in any of the covariate have a multiplicative effect on the baseline risk. The β estimators can be obtained by maximizing the log likelihood function given above.

To test the proportionality assumption of the hazard rates, I use the Schoenfeld PH test in Stata. The Schoenfeld test is a three step process where the hazard model is first run and the Schoenfeld residuals for each predictor are saved. Second, the variables are ranked according to the failure times. In other words, the child with the earliest event of death gets a rank of 1 and the next gets 2 and so on. Finally, correlation between the ranked failure times from step two and the observed failure times from step one are tested. Therefore, the null hypothesis of the test is that there is no correlation between the residuals from the first step and the ranked failure times from the second step. In other words, the proportional hazard assumption is valid. The results of the test are on table 2.02 and 2.03. With the p-values for all covariates higher than the significance level, we can conclude that the proportionality assumptions are valid. To ensure the validity of the model, we need to first test the fit of the model. We can use the Wald test of overall goodness of fit. The null hypothesis of the test is that the hazard ratio is equal to one which is equivalent to regression coefficients being equal to zero. In other words, it tests how far the estimated parameters are from zero. The Wald statistics is then expressed

as
$$Z = \frac{\widehat{\beta}_j}{se(\widehat{\beta}_j)}$$
 where Z is distributed as χ^{2} (1). All of the Wald statistics and their

corresponding p-values are listed in the results table (tables 2.05, 2.06, 2.07, and 2.08). Since all of the Wald p values are less than the significance level, we can conclude that the model is fit. A second set of test of fit of the model is the test of the assumptions of proportionality hazard (PH). STATA performs this test using the stphtest command. The null hypothesis of this test is that the Cox PH assumptions are valid. Since all of the rho values are statistically insignificant (p-values are higher than the significance level), we meet the basic proportionality assumption of the model. The results are given on table 2.05, 2.06, 2.07, and 2.08 in the appendix section. The results show that we have the expected signs and magnitude for all of the determinants of childhood mortality. Since the focus of this paper is the quota system, only the coefficients for the quota variable and relevant variables are listed in the table. By construction, using equation 2.2 above, $\exp(\hat{\beta})$ is the ratio of the group hazard to the baseline hazard. In other words, if the baseline group is a group of states with no SC representation, the $\exp(\hat{\beta})$ will tell us the increase or decrease of the hazard rate compared to that group. Therefore, if we are interested in the hazard rates of a group compared to some base group, we should take the exponential of the beta coefficients. If the coefficient for QD of 5 is -1.20, the hazard ratio = exp(-1.20) = 0.301. In

other words, if the representation is 50-60%, the hazard (risk) of SC infant deaths is 0.699 of the SC population without political representation. This is equivalent of saying that the risk of SC infant deaths decreases by 69.90% if the SC representation went from 0% to the 50-60%.

Table 2.05 shows the Cox Proportional hazard coefficients for the IMR of the Scheduled Caste children. If none of the covariates are included in the regression, I find that if the quota variable is 50-60%, instead of no representation, the risk of SC infants dying decreases by 68.56%. Similarly, if all of the determinants of childhood mortality are included in the regression, I find that if the quota variable is 50-60%, instead of no representation, the risk of sc infants dying decreases by 78.47%. If the socio-economic variables are dropped, I find that the risk of SC infants dying decrease by 77.21% with 50-60% representation. The decrease in SC infant's risk of dying is 77.39%, 73.07% and 75.87% if the environmental, biological and behavioral determinants are dropped respectively. We can also note from table 2.05 that the wealth of the SC household plays a big role in the IMR. For instance, if the SC household falls under the richest category, the risk of infant mortality decreases by 32.56%.

Table 2.06 shows the Cox Proportional hazard coefficients for the U5MR of the Scheduled Caste children. If none of the covariates are included in the regression, I find that if the quota variable is 50-60%, instead of no representation, the risk of SC infants dying decreases by 66.57%. Similarly, if all of the determinants of childhood mortality are included in the regression, I find that if the quota variable is 50-60%, instead of no representation, the risk of SC infant dying decreases by 72.03%. If the socio-economic variables are dropped, I find that the risk of SC infants dying decrease by 73.07% with 50-60% representation. The decrease in SC infant's risk of dying is 70.48%, 70.24% and 69.33% if the environmental, biological and behavioral determinants are dropped respectively. We can also note from the table that the wealth of the SC household plays a big role in the U5MR. For instance, if a SC household falls under the richest category, the risk of under-five mortality decreases by 37.12% while if you fall under the "poorest" category, the risk of under-five mortality increases by 83.50%

Table 2.07 shows the IMR for SC children by the type of residence of the households. With all of the covariates included, I find that in rural areas, if the quota variable is 50-60%, instead of no representation, the risk of SC infant dying decreases by 80.38%. In the urban areas, the risk decreases by 82.32% if the representation is 50-60%. If none of the covariates are included in the regression, I find that the risk of SC infants in rural areas dying decreases by 61.74%- that number is 75.21% in urban areas.

Table 2.08 shows the U5MR for SC children by the type of residence of the households. With all of the covariates included, I find that in the rural areas, 50-60% quota representation decreases the risk of SC infants in rural areas dying by 83.05%. Likewise, in the urban areas the risk decreases by 70.83% with the representation of 50-60%. If none of the covariates are included in the regression, I find that the risk of under-five SC children in rural areas dying decreases by 74.23%- that number is 69.97% in urban areas.

In my dataset, the IMR among SC children is 37.3 per 1000 live births while that among all Indian social groups is 30.8 per 1000 live births. Similarly, the U5MR among SC children is 65.3 per 1000 live births and that among all social groups is 52.4 per 1000 live births. The results show that if any state moves from no representation to at least 50% representation, the IMR among SC children could drop from 37.3 to as low as 8.03 per 1000 live birth. Likewise, U5MR among SC children could drop from 65.3 to as low as 18.26 per 1000 live births, if states in India moved from no SC representation to a 50% SC representation. It is important to note, however, that most states in India have some level of SC representation in the state's legislative assembly. In the dataset, about 6.5% of the observations had a QD of less than 50%. Therefore, the actual drop in IMR and U5MR might not be as dramatic as the results might suggest for all states. Likewise, overall improvements on health technologies could be attributed to some of the decreases in the IMR and U5MR. Refer to table 2.01 for the summary statistics on the variable.

3.6 CONCLUSION

The role of the reservation system in the SC child health is an important issue to understand. In this paper, I have tried to understand the causal impact of the quota system in India on the IMR and the U5MR of the SC children. The overall result shows that the quota system plays a positive and significant role in shaping the health of the SC children. I find that the impact of the 50-60% representation is crucial for the minorities. This is important because the results show that as soon as the SC population has the majority representation (50% or higher), the impact on the IMR and U5MR starts becoming more significant. In the dataset, only about 7% of the observations have less than a 50% representation. Hence, one of the key results from the paper is that a majority representation in the state legislature is all that is required to have a dent in the SC IMR and U5MR. I find that the impact of the quota system in reducing the IMR and U5MR are comparable. Finally, I find that the effect of the quota system is much smaller in the urban areas than in the rural areas. This is not very surprising since the urban areas tend to have more doctors, hospitals and better infrastructure compared to rural India. The private health care services tend to focus in urban areas because of the financial viability

and the accesses of basic infrastructure like road, electricity, and water. This leads to the larger supply of healthcare services in urban areas leaving the government to focus in rural areas. Hence, more representation is more likely to channel the public funds and infrastructure in the rural areas more easily than to the urban areas.

CHAPTER FOUR: DISCUSSION

Although various forms of Affirmative Action policies have been in place in many countries in the world, the true socio-economic impact of those policies on the targeted populations has not been definitively assessed. This dissertation fills the void in the existing literature by finding additional evidence of the impact of one such AA policy in India on India's minority SC population. By assessing the impact of India's political reservation system on its minorities' education and health outcomes, I have contributed to the larger discussion on the effectiveness of affirmative action policies in general. If the goal of the political reservation system in India is to uplift the socio-economic status of its minorities, and eventually erode away the underlying social and economic discrimination that still exists based on the caste system, then the reservation system in India should be considered a "success."

This dissertation finds that the reservation system in India has been fairly effective in improving the education and health outcomes of minorities. The gross enrollment ratios (GERs) in various levels of schooling would increase 1.61% - 7.64% for every 10% increase in the quota share of the SC population in the state's Legislative Assemblies. More specifically, if the SC's population share did not change, I find that an increase of 1% of the SC quota in the state's legislative seat would increase gross primary enrollment rate by about 5.55, gross upperprimary enrollment rate by 5.69, gross secondary enrollment rate by 8.63, and gross higher secondary enrollment rate by 7.75. Likewise, the dropout rates among the SC students would fall 1.88% - 5.25% for every 1% increase in the quota share. Overall, that is a very promising indicator that the quota system has been effective in improving the education outcomes of the targeted groups. Similarly, I also find that the infant mortality rates among the SC children could fall 68% - 78% and the under-five mortality rates could fall 66% - 73% if the quota share is at least 50% of the SC's population share. This means that in some states, even an increase of one seat for the SC would mean increased school enrollments, decreased drop outs and decreased deaths of SC infants and children. Majority of the existing literature have already concluded that the current political reservation system in India have had positive impact in the poverty reduction and transfer of wealth for the SC population thereby likely improving the well-being of those minority groups. My dissertation adds a direct evidence of improved wellbeing of SC as a result of the political reservation to the literature. The findings of a positive and significant impact on the minority's education and health is a significant step towards understanding the real impact these reservation policies have had over the past few decades.

From the policy maker's point of view, this study finds that there are significant positive effects of the minority political quota system on the minority's education and health outcomes. Even though economists deal with many variables in the environment to study any issues-making it harder to estimate the exact impact of any laws, this study, coupled with previous findings on the impact of such policies from other authors, shows that the law is an important safety blanket for individuals under the SC category. In reality, between 1976 and 2008, the SC quota share increased from 13.83% to 14.73% (+6.5%). Between the same periods, the SC primary GER increased from 101.2 to 128.5 (+26.097%), and upper primary GER increased from 31.1 to 85.2 (+173.9%). Likewise, the IMR and U5MR among the SCs have also fallen during that

period. Between 1992 and 2005, the U5MR among the SC children have fallen from 149.1 to 88.1 (-40.9%) and the IMR have fallen from 86.3 to 71.0 (-17.7%). These statistics show that India has made progress, but much more could be done with a more carefully crafted political quota system.

Even though the Scheduled Castes (SCs) form the biggest portion of India's minority groups, the Scheduled Tribes (STs) and the Other Backward Classes (OBCs) form a significant minority. According to the National Sample Survey conducted in 2007 by the Government of India, 40.9% of India's population is OBC, 19.6% is SC, 8.6% is ST and 30.9% is the rest of the population. Even though the OBC population has no Constitutional political reservation in place, there are quotas assigned for them in the education and employment sectors. Future research in this field could focus on the impact of those quotas on OBC's socio-economic outcomes. Similarly, future papers should also focus on the impact of the political reservation policy that assigns legislative quotas to STs on the ST population's socio-economic outcomes.

However, from a policy maker's point of view, studying the impact of such policies on minorities' outcomes itself does not provide evidence of the "success" of the policy. Future research should also focus on the impact of those AA policies on the general population's outcomes. The question should be: "Has all of the reservation policies in many sectors of day-today lives, helped or hurt the general population?" The conclusive evidence of the "success" of the policy would be if the policy is Pareto efficient wherein the overall benefits to the minorities, outweigh the overall loss to the general population.

There is a clear positive and statistically significant impact of the political reservation on the education and health outcomes of the SC population. However, the fact that the

Delimitation Commission has only met four times since the inception of the Constitution means that the SC legislative seat proportion and the SC population proportion has been the same only 4 times in the past 67 years. Therefore, a major policy recommendation of this dissertation is that the Government of India should find a way to equate the SC seat proportion and the SC population proportion every year even if the Delimitation Commission might not meet that often.

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APPENDIX A

Census Year	Total population literacy rate	SC literacy rate
1961	24.0	10.3
1971	29.5	14.7
1981	36.2	21.4
1991	52.1	37.4
2001	65.4	54.7
2011	73.0	66.1

Table 1.01: Literacy rate differentials among social groups in India

Source: Census of India Digital Library

Census Year	India (%)	SC population (%)
1961	24.02	10.27
1971	29.45	14.67
1981	36.23	21.38
1991	52.10	37.41
2001	65.00	54.69

Source: Ministry of Human Resource Development/ Census of India digital library. Note: The enrollment rates (expressed as percentages) are the number of children enrolled in a level, divided by the population of the age group that officially corresponds to the same level.

Indicators	SC	ST	India
% Population below poverty line (2011)	29.4	43.0	22.0
Rural	31.5	45.3	25.4
Urban	21.7	24.1	13.7
Literacy rate (2011)	66.1	59.0	73.0
Male	75.2	68.5	80.9
Female	56.5	49.4	64.6
Gross enrollment ratio (GER)- primary			
1990	106.4	104.0	100.1
2004	115.3	121.9	107.8
Gross enrollment ratio (GER)- upper primary			
1990	52.7	40.7	62.1
2004	70.2	67.0	19.9
Dropout rates (grades 1-10)			
1990	77.7	85.0	71.3
2004	71.3	79.0	61.9

Table 1.03: Basic differentials among social groups

Variable	Obs	Mean	Std. Dev.	Min	Max
Gross Enrollment Ratio- primary (GERP)	476	105.84	31.22	11.02	160.57
Gross Enrollment Ratio- upper primary (GERUP)	486	74.59	25.02	12.63	119.73
Gross Enrollment Ratio- secondary (GERS)	157	71.71	28.28	12.13	151.15
Gross Enrollment Ratio- higher secondary (GERHS)	160	46.96	22.79	1.92	99.98
SC reserved seats	666	20.62	20.99	0	91
Quota Share (QS)	663	0.828	0.354	0	1.81
SC per capita income	627	19098.95	14887.7	3985.9	106373.7
State's education expenditure per pupil	576	9851880	9.86e07	2560.4	1.73e09
State's welfare expenditure per capita	569	37604.04	221531.8	0	2740719
Number of primary schools	624	23624.7	26825.9	14	147376
Number of upper primary schools	627	8749.25	11864.2	7	96797
Number of secondary schools	180	4050.55	4753.2	63	19053
Number of higher secondary schools	180	2242.83	2278.8	48	10739

Table 1.04: Summary Statistics- Education

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Table 1.05: Pairwise	correlation	among indene	ndent variables
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	Election Year	SC per capita income	SC population density	SC Quota Share	Education expenditures per pupil	Welfare expenditure per SC	Number of primary schools	Number of upper- primary schools	Number of secondary schools	Number of higher secondary schools
Election Year	1.000									
SC per capita income	-0.001	1.000								
SC population density	0.004	0.566	1.000							
SC Quota Share	0.021	0.255	0.169	1.000						
Education expenditures per pupil	-0.022	0.017	-0.037	-0.240	1.000					
Welfare expenditure per SC	-0.0080	0.031	-0.062	-0.410	0.081	1.000				
Number of primary schools	0.006	-0.035	-0.188	0.182	-0.079	-0.155	1.000			
Number of upper-primary schools	-0.020	-0.194	-0.194	0.165	-0.064	-0.119	0.709	1.000		
Number of secondary schools	-0.024	-0.279	-0.215	0.119	-0.121	-0.148	0.602	0.582	1.000	
Number of higher secondary schools	0.007	-0.268	-0.141	0.241	-0.141	-0.189	0.810	0.738	0.697	1.000

Table 1.06: Fixed-effect – mediator variables (the mechanism)

Variables/Method	FE (1)	FE (2)	FE(3)	FE(4)	FE (5)	FE (6)
Log Quota Share (lag 1)	0.195**	0.386**	0.201	0.217	0.275	0.261
	(0.076)	(0.181)	(0.618)	(0.422)	(0.965)	(0.561)
Log Quota Share (lag 2)	0.059	-0.017	0.732***	1.97***	0.782*	0.087
	(0.088)	(0.141)	(0.157	(0.124)	(0.435)	(0.437)
Constant	8.99	7.53	7.18	6.00	9.66	7.97
	(0.082)	(0.156)	(0.134)	(0.098)	(0.557)	(0.423)
State fixed effects	Yes	Yes	Yes	Yes	Yes	Yes
Year fixed effects	Yes	Yes	Yes	Yes	Yes	Yes
Number of observations	545	548	174	174	525	516
Number of groups	29	29	29	29	29	29
Overall R ²	0.028	0.044	0.053	0.164	0.019	0.119

Note: Dependent variable are- log number of primary schools (column 1); log number of upper primary schools (column 2); log number of secondary schools (column 4); log state's education expenditure (column 5); and log State's welfare expenditure (column 6). Robust standard errors clustered with States are in parentheses. *, **, *** indicates statistical significance at α = 0.10, 0.05 and 0.01 respectively.

Variables/Method	One step System GMM	One step System GMM	Two step System	Two step System
	(1)	(2)	GMM (3)	GMM (4)
Log SC Gross Enrollment Ratio- primary (lag 1)	0.958***	0.852***	0.958***	1.029***
	(0.026)	(0.120)	(0.027)	(0.156)
Log quota share	0.387*	0.357**	0.384*	0.555***
	(0.193)	(0.160)	(0.199)	(0.193)
Log SC income per capita		0.021		0.0034
		(0.043)		(0.058)
Log number of primary schools (lag1)		0.018		-0.013
		(0.027)		(0.048)
Log state education expenditure (lag1)		0.014		-0.036
		(0.031)		(0.055)
SC population density		-0.00005		-0.00009
		(0.00006)		(0.00007)
Election Year		-0.058		0.0001
		(0.071)		(0.061)
Other mediator variables	Ν	Υ	Ν	Y
AR (1)	-2.73	-1.80	-2.64	-1.58
(p-value)	(0.006)	(0.071)	(0.008)	(0.115)
AR (2)	1.61	1.50	1.56	1.27
(p-value)	(0.108)	(0.133)	(0.118)	(0.205)
Hansen J-statistic	22.89	18.85	23.86	12.92
(p-value)	(0.691)	(0.988)	(0.638)	(0.954)
Hansen C-statistic	20.50	14.36	21.28	10.32
(p-value)	(0.365)	(0.996)	(0.321)	(0.502)
F-statistic	53245.55	37970.81	52521.28	113634.71
(p-value)	(0.000)	(0.000)	(0.000)	(0.000)
No. of instruments	29	42	29	30
No. of observations	401	308	401	308

Table 1.07: Dynamic Panel model- GER- primary

Note: Dependent variable is Log SC Gross Enrollment Ratio- primary (GERP). Robust clustered standard errors are in parentheses. *, **, *** indicates statistical significance at α = 0.10, 0.05 and 0.01 respectively. Column 1 and 3 excludes all of the mediator variables. Column 2 and 4 includes the mediator variables. Other mediator variables include States' education expenditures, welfare expenditures and the interaction of those variables with the quota variable.

Variables/Method	One step System	One step System	Two step system	Two step system
	GMM (1)	GMM (2)	GMM (3)	GMM (4)
Log SC Gross Enrollment Ratio – upper primary	0.769***	0.811***	0.402*	0.855**
(lag 1)	(0.162)	(0.138)	(0.204)	(0.311)
Log quota share (lag 1)	0.200	0.275*	0.398*	0.764*
	(0.189)	(0.147)	(0.233)	(0.439)
Log SC income per capita		0.024		0.0009
		(0.380)		(0.089)
Log number of upper primary schools		0.027		0.016
		(0.023)		(0.045)
Log state education expenditure		0.022		0.001
		(0.047)		(0.064)
Election Year		0.032		0.072
		(0.035)		(0.058)
Other mediator variables	Ν	Y	Ν	Y
AR (1)	-2.11	-2.52	-1.91	-2.00
(p-value)	(0.035)	(0.012)	(0.056)	(0.046)
AR (2)	0.13	0.52	0.10	0.54
(p-value)	(0.900)	(0.600)	(0.918)	(0.589)
Hansen J statistic	16.51	13.40	11.49	13.44
(p-value)	(0.989)	(0.999)	(0.778)	(0.568)
Hansen C-statistic	11.75	14.47	5.98	3.73
(p-value)	(0.988)	(0.490)	(0.875)	(0.811)
F-statistic	71268.18	20927.62	59733.94	9446.49
(p-value)	(0.000)	(0.000)	(0.000)	(0.000)
No. of instruments	35	45	20	21
No. of observations	385	301	329	301

Table 1.08: Dynamic Panel model- GER- upper primary

Note: Dependent variable is Log SC Gross Enrollment Ratio- upper primary (GERUP). Corrected standard errors are in parentheses. *, **, *** indicates statistical significance at α = 0.10, 0.05 and 0.01 respectively. Column 1 and 3 are the results of one step and two system GMM respectively and excludes all of the mediator variables. Column 2 and 4 includes the mediator variables.

Variables/Method	One step System GMM	One step System GMM	Two step system GMM	Two step system GMM
	(1)	(2)	(3)	(4)
Log SC secondary enrollment per capita	0.823***	0.587***	0.823**	0.454***
(lag 1)	(0.049)	(0.087)	(0.051)	(0.157)
Log quota share	1.218***	0.943*	1.203***	1.494
	(0.296)	(0.491)	(0.292)	(0.990)
Log SC income per capita		0.180		0.004
		(0.120)		(0.142)
Log number of secondary schools		-0.037		0.118
		(0.072)		(0.101)
SC population density		-0.0001		0.0001
		(0.0002)		(0.0003)
Log welfare per capita				0.341**
				(0.158)
Election Year		-0.164*		-0.229
		(0.083)		(0.186)
Other mediator variables	Ν	Y	Ν	Y
AR (1)	-3.70	-2.93	-3.48	-1.90
(p-value)	(0.000)	(0.003)	(0.001)	(0.058)
AR (2)	-0.81	-1.19	-0.79	-0.54
(p-value)	(0.419)	(0.232)	(0.427)	(0.589)
Hansen J- statistic	24.33	18.86	24.33	46.42
(p-value)	(0.228)	(0.943)	(0.228)	(0.243)
Hansen C- statistic	23.50	14.68	23.50	5.40
(p-value)	(0.101)	(0.795)	(0.101)	(0.611)
F-statistic	17407.88	5023.67	7041.42	1636.58
(p-value)	(0.000)	(0.000)	(0.000)	(0.000)
No. of instruments	22	38	22	26
No. of observations	124	123	124	123

Table 1.09: Dynamic Panel model- SC Gross secondary enrollment ratio (GERS)

Note: Dependent variable is Log SC secondary school enrollment per capita. Corrected standard errors are in parentheses. *, **, *** indicates statistical significance at α = 0.10, 0.05 and 0.01 respectively. Column 1 and 3 are the results of one step and two system GMM respectively and excludes all of the mediator variables. Column 2 and 4 includes the mediator variables.

Variables/Method	One step System	One step System	Two step system	Two step system
	GMM (1)	GMM (2)	GMM (3)	GMM (4)
Log SC higher secondary enrollment per	0.591***	0.735***	0.493***	0.821***
capita (lag 1)	(0.087)	(0.102)	(0.102)	(0.051)
Log quota share	0.819**	1.023*	1.651***	0.362*
	(0.307)	(0.506)	(0.389)	(0.210)
Log SC income per capita		0.067		0.106*
		(0.099)		(0.059)
Log education budget		-0.024		-0.088
		(0.087)		(0.053)
Log welfare budget		0.086		0.092
		(0.074)		(0.061)
Literacy rate		-0.0068		-0.0021
		(0.0145)		(0.008)
Election Year	0.078	-0.235**	0.264***	-0.014
	(0.072)	(0.110)	(0.079)	(0.086)
Other mediator variables	Ν	Y	Ν	Y
AR (1)	-0.39	-2.77	-1.24	-2.53
(p-value)	(0.695)	(0.006)	(0.216)	(0.011)
ĀR (2)	-1.04	-1.02	-1.04	-1.58
(p-value)	(0.298)	(0.430)	(0.296)	(0.100)
Hansen J-statistic	20.64	18.49	18.69	23.70
(p-value)	(0.965)	(0.555)	(0.849)	(0.966)
Hansen C-statistic	16.60	15.91	18.52	23.76
(p-value)	(0.921)	(0.195)	(0.553)	(0.533)
F-statistic	5059.10	2500.13	3494.40	5065.59
(p-value)	(0.000)	(0.000)	(0.000)	(0.000)
No. of instruments	38	28	30	45
No. of observations	98	126	98	126

Table 1.10: Dynamic Panel model- SC Gross higher secondary enrollment ratio (GERHS)

Note: Dependent variable is Log SC higher secondary school enrollment per capita. Corrected standard errors are in parentheses. *, **, *** indicates statistical significance at α = 0.10, 0.05 and 0.01 respectively. Column 1 and 3 are the results of one step and two system GMM respectively and excludes all of the mediator variables. Column 2 and 4 includes the mediator variables.

Variables/Method	One step System	Two step System	Two step system	One step System	Two step System
	GMM – primary	GMM – primary	GMM- upper	GMM- secondary	GMM- secondary
	(1)	(2)	primary (3)	(4)	(5)
Log SC dropout rate (lag	0.183	0.201	0.549**	0.314*	0.313*
1)	(0.179)	(0.224)	(0.253)	(0.171)	(0.167)
Log quota share	-4.256*	-4.139*	-5.247**	-2.23*	-1.877*
	(2.179)	(2.422)	(2.239)	(1.33)	(0.987)
Log SC primary	-4.815	-3.412			
enrollment per capita	(5.814)	(5.662)			
Log SC income per	0.415	0.354	0.142	-0.131	-0.139
capita	(0.329)	(0.358)	(0.168)	(0.239)	(0.309)
Log number of schools	0.189	0.214	0.362*	0.576	0.572
	(0.158)	(0.157)	(0.181)	(0.475)	(0.611)
Log student-teacher	0.473	0.355		0.223	0.217
ratio	(0.385)	(0.511)		(0.236)	(0.265)
Election Year	0.108	0.165	-0.0267		
	(0.266)	(0.243)	(0.110)		
AR (1)	-1.35	-1.07	-1.44	-1.74	-1.61
(p-value)	(0.178)	(0.284)	(0.150)	(0.081)	(0.107)
AR (2)	-0.66	-0.55	-1.02		
(p-value)	(0.508)	(0.582)	(0.310)		
Hansen	17.27	17.27	17.80	19.34	19.34
(p-value)	(0.635)	(0.635)	(0.962)	(0.500)	(0.500)
Hansen Diff	-1.98	-1.98	-3.05	1.30	1.30
(p-value)	(0.999)	(0.999)	(0.999)	(0.730)	(0.730)
No. of instruments	27	27	35	25	25
No. of observations	141	141	211	91	91

Table 1.11: Dynamic Panel model- SC Dropout rates

Note: Dependent variable is Log of dropout rates. Corrected standard errors are in parentheses. *, **, *** indicates statistical significance at α = 0.10, 0.05 and 0.01 respectively.

Variables	SC Primary	SC upper primary
	enrollment (1)	enrollment (2)
Log SC enrollment per capita (lag 1)	0.915***	0.646***
	(26.97)	(0.102)
Log quota share (lag 1)	0.106*	0.832***
	(0.059)	(0.214)
AR (1)	2.35	-2.90
(p-value)	(0.019)	(0.004)
AR (2)	1.48	-0.61
(p-value)	(0.139)	(0.540)
Hansen	8.23	1.57
(p-value)	(0.411)	(0.667)
Hansen Diff	0.48	1.14
(p-value)	(0.924)	(0.566)
No. of instruments	10	5
No. of observations	2924	2958

Table 1.12: Robustness check- DPD with district level data

Note: Dependent variables are Log SC primary and upper primary enrollments per capita. Corrected standard errors are in parentheses. *, **, *** indicates statistical significance at α = 0.10, 0.05 and 0.01 respectively. Both columns 1 and 2 are the results of two step system GMM

Variables	SC primary enrollment (1)	SC upper primary enrollment (2)
Log SC enrollment per capita (lag 1)	0.367	-0.153
	(0.458)	(1.09)
Log quota share	1.276*	1.717*
	(0.692)	(0.824)
Log SC income per capita	-0.132	-0.330
	(0.086)	(0.274)
Log number of schools	0.123**	0.186
	(0.055)	(0.157)
Log number of schools * Log quota share	-0.157**	-0.201
	(0.063)	(0.135)
Literacy rate	0.008	0.029
	(0.006)	(0.027)
AR (1)	-0.86	-0.18
(p-value)	(0.391)	(0.855)
AR (2)	-0.39	-0.41
(p-value)	(0.696)	(0.679)
Hansen	9.30	9.77
(p-value)	(0.504)	(0.461)
Hansen Diff	5.15	4.28
(p-value)	(0.525)	(0.639)
No. of instruments	16	16
No. of observations	106	108

Table 1.13: Robustness check- DPD with state-level election cycle data

Note: Dependent variables are Log SC primary and upper primary enrollments per capita. Corrected standard errors are in parentheses. *, **, *** indicates statistical significance at α = 0.10, 0.05 and 0.01 respectively. Both columns 1 and 2 are the results of two step system GMM

Variables	GPER (1)	GPER (2)	GUPER (3)	GUPER (4)
Log SC enrollment per capita (lag 1)	1.062***	-0.021	0.843***	0.265
	(0.115))	(0.148)	(0.089)	(0.462)
Log quota share	0.079**	2.727***	0.108*	2.274*
	(0.032)	(0.944)	(0.057)	(1.333)
Log SC income per capita		-0.200*		-0.189**
		(0.103)		(0.084)
Log number of schools		0.099		0.091
-		(0.088)		(0.119)
Log number of schools * Log quota share		-0.159		-0.204
		(0.116)		(.0130)
Literacy rate	-0.0005*			
	(0.0003)			
SC population density	-0.00002	0.0003	-0.00003	-0.0005
	(0.00003)	(0.0007)	(0.0001)	(0.0007)
Election Year dummy	-0.0165*	0.0016	0.0333	-0.0334
	(0.0087)	(0.008)	(0.0379)	(0.029)
Other mediator variables	N	Ŷ	N	Ŷ
AR (1)	-2.63	-1.06	-3.33	-1.39
(p-value)	(0.009)	(0.290)	(0.001)	(0.166)
ĀR (2)	1.43	-0.20	0.77	0.97
(p-value)	(0.151)	(0.841)	(0.443)	(0.330)
Hansen	16.42	5.80	16.93	5.59
(p-value)	(0.355)	(0.760)	(0.260)	(0.992)
Hansen Diff	7.93	2.48	4.65	-1.26
(p-value)	(0.440)	(0.779)	(0.589)	(0.999)
No. of instruments	21	20	18	28
No. of observations	442	342	474	335

Table 1.14: Robustness Check- DPD without UTs & newer states

Note: Dependent variables are Log SC primary enrollments per capita (column 1 and 2) and upper primary enrollments per capita (column 3 and 4). Corrected standard errors are in parentheses. *, **, *** indicates statistical significance at α = 0.10, 0.05 and 0.01 respectively. All the estimates are the results of two step system GMM.

		All survey sample		duled Caste
Categorical Variables	Frequency (N)	Percent (%)	Frequency (N)	Percent (%)
Gender of child				
Male	133,435	51.96	25,495	51.66
Female	123,348	48.04	23,857	48.34
Religion				
Hindu	182,752	71.17		
Muslim	39,955	15.56		
Place of residence	,			
Urban	101,981	39.72	18,336	37.15
Rural	154,788	60.28	31,016	62.85
Caste	104,700	00.20	51,010	02.00
SC	45,486	18.94		
ST	36,794	14.95		
OBC	86,033	34.84		
Mother's education				
No education	131,634	51.26	31,634	64.10
Some education	125,148	48.74	17,718	35.90
Wealth Index				
Poorest	41,057	15.99	10,684	21.65
Poorer	45,707	17.82	10,898	22.08
Richer	57,827	22.52	9,996	20.25
Richest	59,342	23.12	6,736	13.65
Quota	<i>07,042</i>	20.12	0,700	10.00
< 50%	14 525	6 47	908	2.02
	14,535	6.47		
50-60%	10,444	4.65	1,893	4.22
60-70%	6,154	2.74	968	2.16
70-80%	5,498	2.45	1,132	2.52
80-90%	49,474	22.03	10,586	23.61
Breastfed the child				
No	196,593	76.56	37,055	75.08
Yes	60,190	23.44	12,297	24.92
Place of delivery				
Home	23,308	11.04	6,029	12.22
Government facility	9,830	3.83	1,929	3.91
Private facility	10,676	4.16	1,353	2.74
Access to health care	10,070	1110	1,000	
Not a big problem	188,925	73.59	35,761	72.48
Big problem	67,795	26.41	13,576	27.52
ORS				
Never heard of it	67,482	26.28	14,271	28.92
Have used it	4,668	1.82	758	1.54
Antenatal care				
Yes	29,631	11.54	5,309	10.76
No	6,815	2.65	1,350	2.74
Immunized against tetanus				
No	5,453	2.12	1,103	2.23
Yes	30,994	12.07	5,578	11.30
Fuel used for cooking	00,774	12.07	0,070	11.00
Safe	80,948	31.52	11 552	22.41
			11,553	23.41
Unsafe	168,475	65.61	36,336	73.63
Toilet				
Unsanitary	24,728	9.63	2,906	5.89
No toilet	110,932	43.20	27,928	56.59
Water source				
Safe	231,126	90.01	46,070	93.35
Unsafe	17,769	6.92	1,651	3.35

Table 2.01: Summary statistics of covariates of the hazard model

	rho	chi^2	df	P>chi^2
Religion	0.0342	1.27	1	0.259
Breastfed	0.0321	1.27	1	0.259
Access to health	0.0529	2.91	1	0.088
care	0.0027	2.71	1	0.000
Birth order	0.0181	0.35	1	0.552
Mother's age	0.0218	0.54	1	0.464
Type of residence	-0.0128	0.16	1	0.687
Gap between	0.0202	0.70	1	0.401
births				
Wealth Index	-0.0446	1.95	1	0.163
Quota	0.0035	0.02	1	0.892
ORS	-0.0238	0.50	1	0.479
Antenatal care	0.0438	2.39	1	0.122
Access to water	-0.0220	0.55	1	0.459
GLOBAL		16	12	0.191

Table 2.02: Test of proportional hazard assumption- IMR

 Table 2.03: Test of proportional hazard assumption- U5MR

	rho	chi^2	df	P>chi^2
Breastfed	0.0294	2.33	1	0.127
Access to	0.0255	1.60	1	0.206
health care				
Type of	0.0153	0.56	1	0.455
residence				
Gender	0.0413	4.15	1	0.042
Quota	0.0226	1.95	1	0.163
ORS	0.0109	0.21	1	0.650
Antenatal	0.0153	0.70	1	0.402
care				
Access to	-0.0067	0.12	1	0.732
water				
GLOBAL		13.58	8	0.094

Variables	Water	Toilet	Fuel	Tetanus	Antenatal	ORS	Access	Place of	Religion
					care		to	delivery	
							health		
							care		
Water		6.631	0.901	2.631	2.631	0.262	2.619	1.609	31.908
		(0.036)	(0.343)	(0.105)	(0.105)	(0.877)	(0.106)	(0.447)	(0.000)
Toilet	6.631		62.860	1.741	1.081	15.763	15.871	34.071	8.255
	(0.036)		(0.000)	(0.419)	(0.582)	(0.003)	(0.000)	(0.000)	(0.220)
Fuel	0.901	62.860		4.971	4.971	8.788	4.534	24.274	4.902
	(0.343)	(0.000)		(0.027)	(0.027)	(0.012)	(0.033)	(0.000)	(0.179)
Tetanus	2.631	1.741	4.971		92.401	1.126	11.264	5.821	19.665
	(0.105)	(0.419)	(0.027)		(0.000)	(0.569)	(0.001)	(0.054)	(0.000)
Antenatal	2.631	1.081	4.971	92.401		1.126	1.950	5.821	19.665
care	(0.105)	(0.582)	(0.027)	(0.000)		(0.569)	(0.163)	(0.054)	(0.000)
ORS	0.262	15.763	8.788	1.126	1.126		0.091	7.677	1.416
	(0.877)	(0.003)	(0.012)	(0.569)	(0.569)		(0.956)	(0.104)	(0.965)
Access to	2.619	15.871	4.534	11.264	1.950	0.091		9.710	26.315
health	(0.106)	(0.000)	(0.033)	(0.001)	(0.163)	(0.956)		(0.008)	(0.000)
care									
Place of	1.609	34.071	24.274	5.821	5.821	7.677	9.710		6.245
delivery	(0.447)	(0.000)	(0.000)	(0.054)	(0.054)	(0.104)	(0.008)		(0.396)
Religion	31.908	8.255	4.902	19.665	19.665	1.416	26.315	6.245	
0	(0.000)	(0.220)	(0.179)	(0.000)	(0.000)	(0.965)	(0.000)	(0.396)	

Table 2.04: Chi-square test of independence of categorical variables

Table 2.05: Co	IMR (1)	IMR (2)	IMR (3)	IMR (4)	IMR (5)	IMR (6)
Quota Share		11VIK (2)	INIK (3)			
(lag 2)						
(lag 2) 50-60%	-1.157***	-1.422***	-1.312***	-1.487***	-1.479***	-1.536***
50-0070	(0.317)	(0.376)	(0.326)	(0.373)	(0.364)	(0.373)
60-70%	-0.295	-0.510	-0.470	-0.562*	-0.519	-0.603*
00 70 /0	(0.293)	(0.347)	(0.295)	(0.346)	(0.344)	(0.346)
70-80%	(0.290)	-0.836**	-0.438	-0.902***	-0.884***	-0.954***
10 00 10		(0.342)	(0.269)	(0.339)	(0.337)	(0.339)
80-90%		-0.562**	-0.402**	-0.628***	-0.645***	-0.674***
00 2070		(0.236)	(0.208)	(0.234)	(0.233)	(0.233)
		(0.200)	(0.200)	(0.201)	(0.200)	(0.200)
Behavioral	No	No	Yes	Yes	Yes	Yes
variables						
Biological	No	Yes	No	Yes	Yes	Yes
variables						
Environmental	No	Yes	Yes	No	Yes	Yes
variables						
Socio-	No	Yes	Yes	Yes	No	Yes
economic						
variables						
Wealth Index						
Poorest		0.425***	0.545***	0.459***		0.443***
		(0.126)	(0.081)	(0.096)		(0.093)
Poorer		0.399***	0.433***	0.413***		0.406***
		(0.124)	(0.078)	(0.093)		(0.093)
Richer		-0.074	-0.061	-0.062		-0.064
		(0.154)	(0.089)	(0.108)		(0.108)
Richest		-0.412***	-0.462***	-0.388***		-0.394***
		(0.219)	(0.115)	(0.142)		(0.142)
Number of	40268	27974	40258	27986	27992	27986
observations						
Number of	1480	1036	1478	1036	1038	1036
deaths						
Log likelihood	-15594.74	-10307.06	-15444.47	-10298.91	-10345.51	-10293.91
Wald chi2 (p-	72.54	3052.76	4973.50	2935.49	305.46	2982.50
value)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)

Table 2.05: Cox Regression for IMR

Table 2.06: Cox	U5MR (1)	U5MR (2)	U5MR (3)	U5MR (4)	U5MR (5)	U5MR (6)
Quota Share						
(lag 2)						
50-60%	-1.096***	-1.182***	-1.212***	-1.220***	-1.312***	-1.274***
	(0.253)	(0.296)	(0.258)	(0.293)	(0.289)	(0.293)
60-70%	-0.593**	-0.717**	-0.702***	-0.738**	-0.763**	-0.781**
	(0.256)	(0.313)	(0.258)	(0.312)	(0.313)	(0.312)
70-80%		-0.480*	-0.360*	-0.509**	-0.518**	-0.566**
		(0.259)	(0.213)	(0.257)	(0.256)	(0.257)
80-90%		-0.483**	-0.429**	-0.520***	-0.568***	-0.572***
		(0.199)	(0.169)	(0.197)	(0.198)	(0.197)
Behavioral variables	No	No	Yes	Yes	Yes	Yes
Biological	No	Yes	No	Yes	Yes	Yes
variables						
Environmental	No	Yes	Yes	No	Yes	Yes
variables						
Socio-economic	No	Yes	Yes	Yes	No	Yes
variables						
Wealth Index						
Poorest		0.604***	0.658***	0.623***		0.607***
		(0.072)	(0.062)	(0.073)		(0.073)
Poorer		0.429***	0.422***	0.434***		0.426***
		(0.072)	(0.062)	(0.073)		(0.073)
Richer		-0.147*	-0.137*	-0.132		-0.134
		(0.086)	(0.072)	(0.086)		(0.086)
Richest		-0.488***	-0.531***	-0.457***		-0.464***
		(0.112)	(0.093)	(0.112)		(0.112)
Number of	40268	27974	40258	27986	27992	27986
observations						
Number of	2422	1766	2420	1766	1768	1766
deaths						
Log likelihood	-25373.22	-17473.99	-25099.41	-17457.91	-17537.61	-17448.64
Wald chi2 (p-	156.18	5062.28	7415.42	5001.19	513.48	5088.27
value)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)

Table 2.06: Cox Regression for U5MR

	IMR (1)		IMR (2)	
	Rural	Urban	Rural	Urban
Quota Share				
50-60%	-1.629***	-1.733***	-0.961*	-1.395***
	(0.615)	(0.487)	(0.592)	(0.380)
60-70%	-0.559	-0.845*	0.198	-0.768*
	(0.495)	(0.498)	(0.437)	(0.401)
70-80%	-0.556	-1.906***	0.484	-0.777*
	(0.433)	(0.668)	(0.381)	(0.451)
80-90%	-0.579*	-0.812**	0.310	-0.568**
	(0.341)	(0.325)	(0.326)	(0.265)
Behavioral variables	Yes	Yes	No	No
Biological variables	Yes	Yes	No	No
Environmental variables	Yes	Yes	No	No
Socio-economic variables	Yes	Yes	No	No
Wealth Index				
Poorest	0.500***	0.181		
	(0.116)	(0.235)		
Poorer	0.465***	0.214		
	(0.114)	(0.186)		
Richer	-0.079	-0.100		
	(0.169)	(0.148)		
Richest	-0.591*	-0.433**		
	(0.332)	(0.173)		
Number of observations	17478	10508	24601	15667
Number of deaths	705	331	992	488
Log likelihood	-6658.54	-2969.24	-9956.39	-4685.33
Wald chi2 (p-value)	1824.87	729.68	54.77	22.88
•	(0.000)	(0.000)	(0.000)	(0.001)

Table 2.07: Cox Regression for IMR – by residence

Rural -1.775*** (0.527)	Urban -1.232***	Rural	Urban
(0.527)	-1.232***		
(0.527)	-1.232***		
· /		-1.356***	-1.203***
0 710*	(0.383)	(0.509)	(0.304)
-0.710*	-1.024**	-1.608	-1.048***
(0.422)	(0.474)	(0.361)	(0.369)
-0.363	-1.047**		-0.625*
(0.332)	(0.444)		(0.327)
-0.515*	-0.664**		-0.551**
(0.275)	(0.285)		(0.226)
Yes	Yes	No	No
Yes	Yes	No	No
Yes	Yes	No	No
Yes	Yes	No	No
0.680***	0.389**		
(0.089)	(0.164)		
0.522***	0.183		
(0.090)	(0.142)		
-0.139	-0.208*		
(0.138)	(0.114)		
-0.519**	-0.558***		
(0.256)	(0.134)		
17478	10508	24601	15667
			777
			-7423.74
			40.98 (0.002)
	(0.332) -0.515* (0.275) Yes Yes Yes 0.680*** (0.089) 0.522*** (0.090) -0.139 (0.138) -0.519**	(0.332) (0.444) -0.515* -0.664** (0.275) (0.285) Yes Yes 0.680*** 0.389** (0.089) (0.164) 0.522*** 0.183 (0.090) (0.142) -0.139 -0.208* (0.138) (0.114) -0.519** -0.558*** (0.256) (0.134) 17478 10508 1210 556 -11353.97 -4977.09	$\begin{array}{cccccccccccccccccccccccccccccccccccc$

Table 2.08: Cox Regression	for USMR _ by r	osidonco
Table 2.06: Cox Regression	101003 M $ Dy 1$	esidence

APPENDIX B Figure 1.1: Literacy rate differentials among social groups

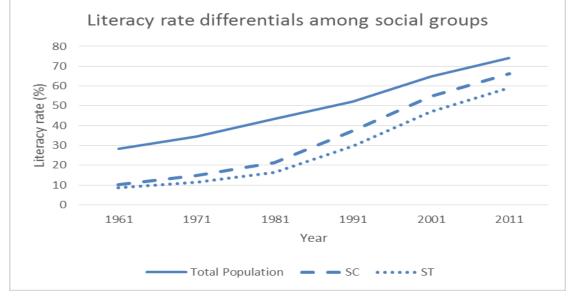
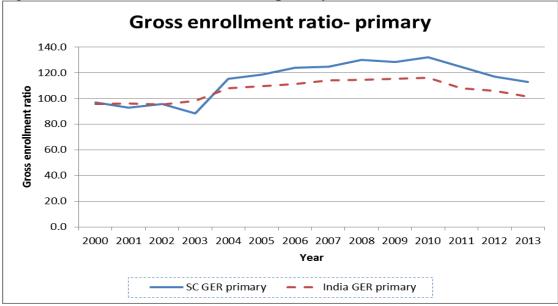


Figure 1.2: Gross enrollment ratio (GER)- primary schools



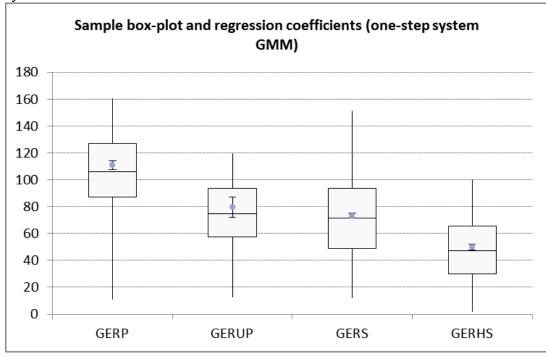
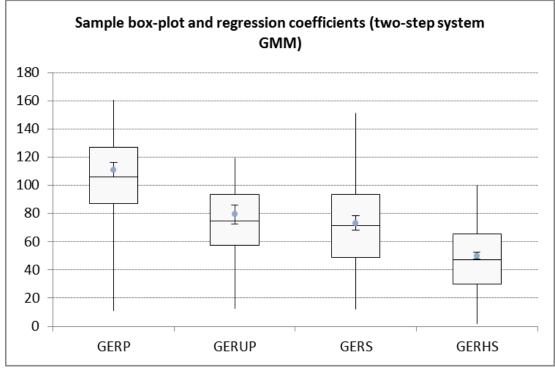
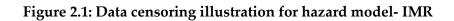


Figure 1.3: Sample box-plot and regression coefficients for all levels of schooling- one-step system GMM

Figure 1.4: Sample box-plot and regression coefficients for all levels of schooling- two-step system GMM





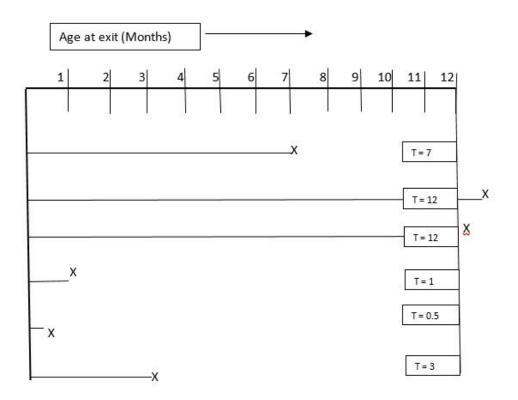
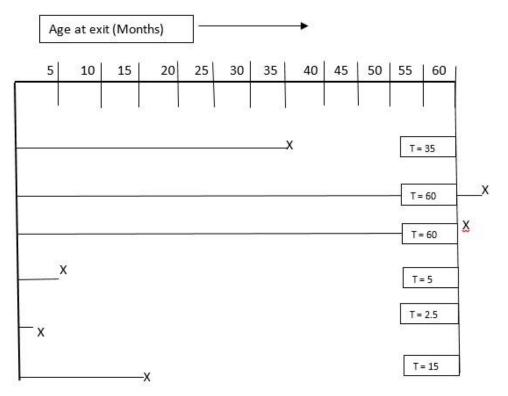


Figure 2.2: Data censoring illustration for hazard model- U5MR



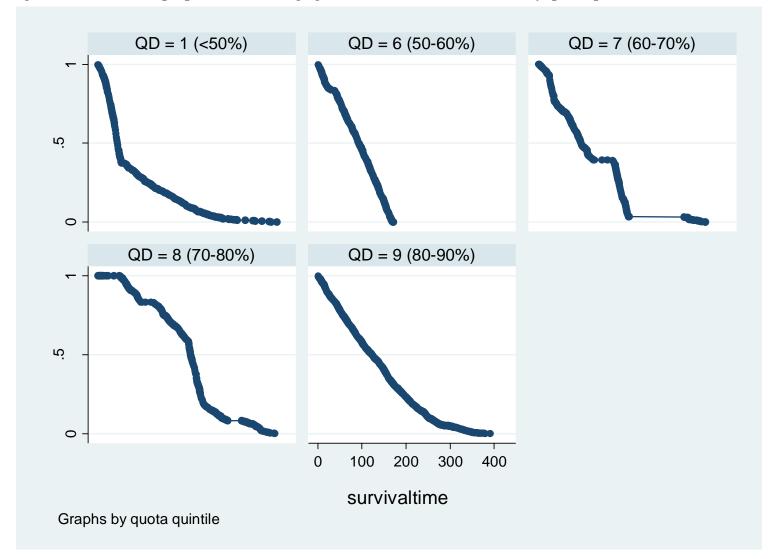


Figure 2.3: SC children proportion surviving against survival time in months- by quota quintile



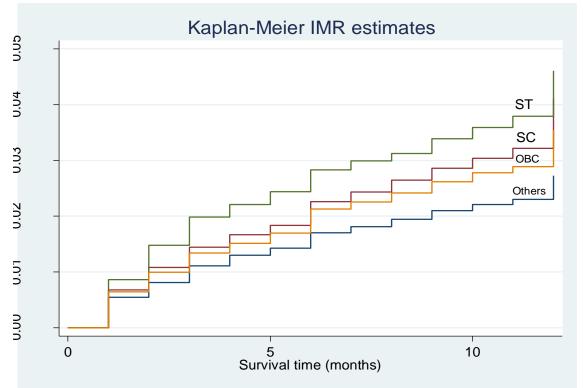
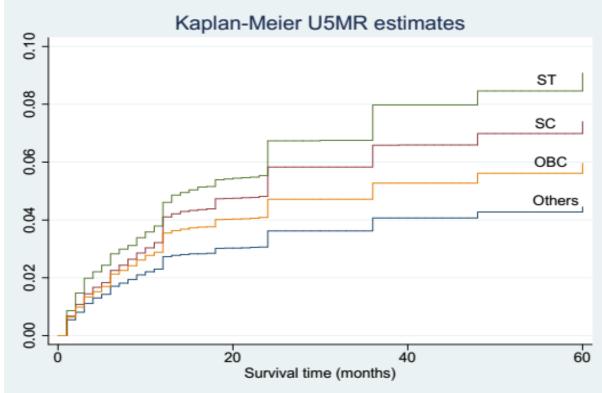


Figure 2.5: Kaplan-Meier hazard plot- U5MR



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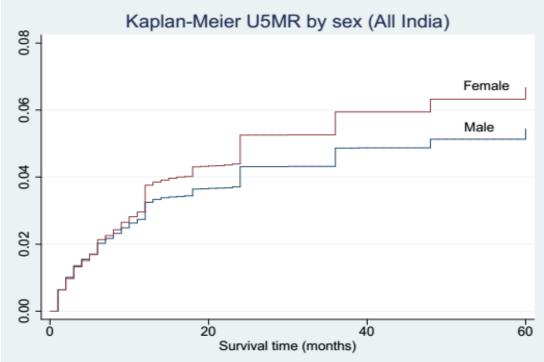
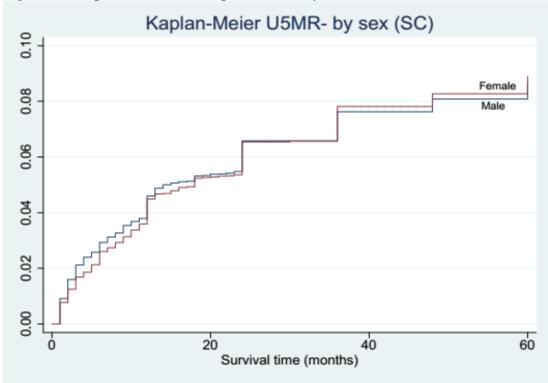


Figure 2.6: Kaplan Meier hazard plot- U5MR by sex (All India)

Figure 2.7: Kaplan Meier hazard plot- U5MR by sex (SC)



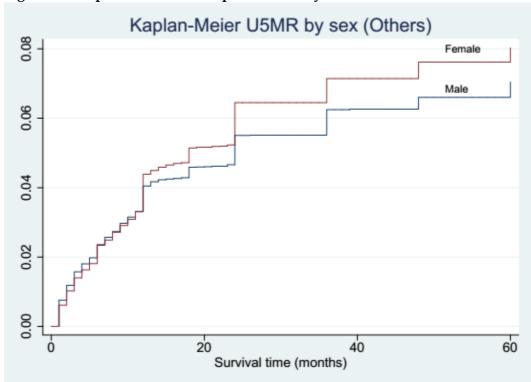


Figure 2.8: Kaplan Meier hazard plot- U5MR by sex (Others)

APPENDIX C: An alternative approach: The Structural Equation Model

To understand the mechanism through which the political reservation affects the education outcomes of SC, I use the mediation analysis with Structural Equation Model (SEM). SEM allows us to test the hypotheses about the relationship among different variables simultaneously. Hence, it is a multivariate technique that allows us to estimate a system of equations as a tool to estimate both the direct and indirect effects of the variable of interest on the dependent variable. It fits models using the observed covariances and the means. However, the use of SEM requires us to have a set of assumptions. First, the sample size has to be "large". The SEM model uses the Maximum Likelihood (ML) estimator. Since ML estimation relies on asymptotics, large sample size is required to get a consistent/efficient estimation of parameters. My data supports this assumption as I have an average of 640 observations per variable in my unbalanced panel data. Secondly, the SEM model assumes that the observed variable follow a multivariate normal distribution because the likelihood that is being maximized in the structural equation model uses the ML estimation. However, this assumption can be relaxed for the exogenous variables in the model. Finally, the SEM model assumes that there are no relevant variables that are omitted from the model -- i.e. the model is correctly specified. Thus, if the model is not fully specified, it might be prone to omitted variable bias, as in the linear regression case. This is one of the shortcomings in this SEM analysis as it is very hard to identify a complete path with all the possible variables. Therefore, this method is expected to have some bias on the coefficient estimates. However, the mediator variables are chosen based on the possible quantifiable mediation effect of QS on enrollments.

Mediation analysis is a part of the SEM model with which one can understand the underlying mechanism of a relationship between a dependent and independent variable through one or more mediator variables. It is important to understand, however, that because of limitation in fully specifying the model paths, the mediation model might not test the direct causal relationship between the dependent and independent variable. Instead, it treats the independent variable as the main cause of mediator variables, which in-turn causes the dependent variable (as depicted in the figure below.) Thus, the mediator variable explains the relationship between the dependent and the independent variable.

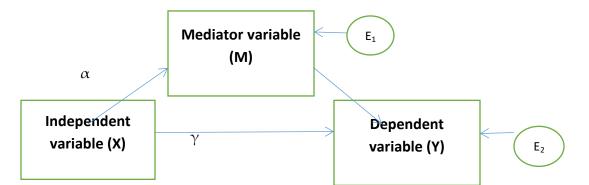


Figure above shows a simple mediation SEM model where the independent variable (X) causes the mediator variable (M), which in-turn causes the dependent variable (Y). Hence, there is a direct effect of X on Y and also an indirect effect from X to M to Y. In the figure above, direct effect = γ , indirect effect = $\alpha\beta$ and the total effect = $\gamma + \alpha\beta$. The following path diagram describes the perceived relationships between different variables which are modeled to analyze the presence and the extent of the impact that mediator variables have in my study. I am interested in the direct, indirect, and the total effects of the QS on two of my dependent variables (SCPEPC and SCUPEPC).

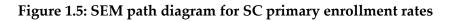
In the path diagram Figure 1.15 and Figure 1.16 (below), the independent variable of interest is LnQS (natural log of QS) and the two dependent variables are the LnSCPEPC and LnSCUPEPC (natural logs of SC primary and SC upper primary enrollments per capita). The mediator variables in the path diagram are STRP, LnNoPS, LnEBPP, LnWBPC, LnNoUPS and STRUP (described above). We can also see that LnQS can be affected by the 'ElectionYear' and the 'SCPopDensity' variables because of the way that the quota is assigned. Since the income of the SC families might have an impact on the primary and upper primary enrollments, LnSCPCInc is also included in the model.

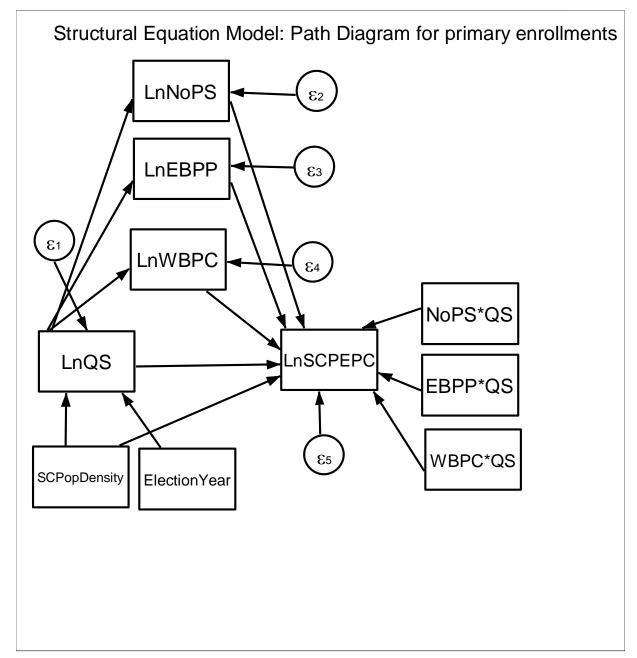
The estimates from the Structural Equation Model (SEM) are provided in Table 1.15 and Table 1.16 below. Since the robust clustered standard errors are used, the goodness of fit for the SEM model is given by the Standardized root mean squared residuals (SRMR). The coefficient of determination and the overall R² are also calculated. All of these test results are included on tables 1.15 and 1.16. Tables 1.15 and 1.16 also contain the direct, indirect and total effects of quotas and other variables on the SC enrollments in primary and upper primary schools respectively. Column 1 of Table 1.15 provides the direct effect of quotas and other variables on primary school enrollment. As expected, the direct impact of quotas on the enrollment rate of SC in primary education is positive and significant. It suggests that a 10% increase in the Quota Share (QS) will increase the enrollment in primary education by 3.3% and it is significant at the 10% level. Similarly, column 2 of Table 1.15 shows the indirect effect of quotas on enrollment. The indirect effects come from the various mediator variables previously described in the path diagram. The results suggest that a 10% increase in QS will result in a 9.34% increase in the SC enrollment in primary education and is significant at the 1% level. Finally, column 3 shows the total effects of QS and other control variables on the SC enrollment. It shows that a 10% increase in QS will likely increase the SC enrollment in primary education by about 13%. The coefficient is significant at the 1% level. It is also important to highlight the significance of the proposed mediator variables. The results listed in column 1 show that the number of schools, the education budget, and the welfare budget all play a significant mediating role. For example, a 1% increase in QS will likely increase in the number of primary schools by 14.4%, an increase in the education budget by 14.7%, and an increase in the welfare budget by 9.23%, all of which lead to an increase in SC primary school enrollment rates.

As previously stated, Table 1.16 provides the SEM results for SC upper primary enrollment rates. Column 1 provides the direct effect of quotas and other variables on the upper primary enrollment. As expected, the direct impact of quotas on the enrollment of SC in upper primary education is positive and small. However, it is not statistically significant. Similarly, column 2 shows the indirect effect of quotas on enrollment. The indirect effects come from the

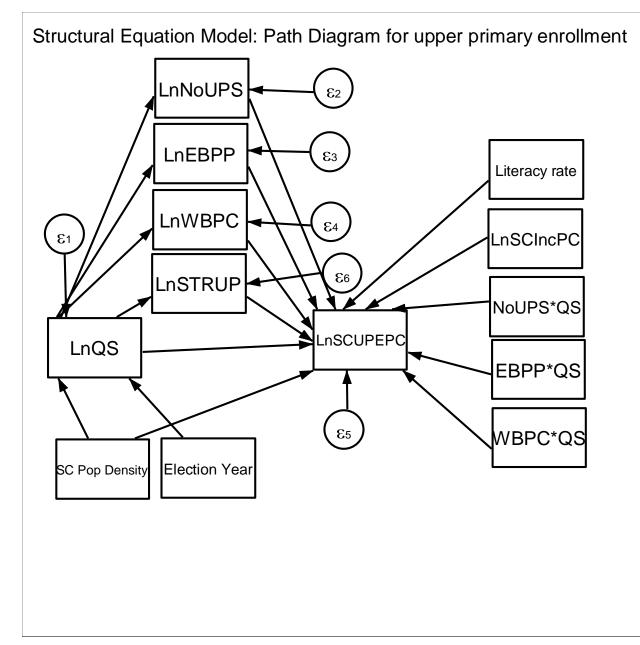
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various mediator variables previously described in the path diagram. The SEM results suggest that a 10% increase in QS will result in a 6.9% increase in the SC upper primary school enrollment; this is significant at the 1% level. Finally, column 3 shows the total effects of QS and other control variables on the SC enrollment rate. It shows that a 10% increase in QS will likely increase the SC enrollment rate in upper primary education by about 10%. The coefficient is significant at the 1% level. As before, the role of mediator variables is important to highlight. The results listed in column 1 show that the number of schools, the education budget, and the welfare budget all play a significant mediating role. For example, a 1% increase in QS will most likely increase the number of upper primary schools by 12.5%, the education budget by 14.8%, and the welfare budget by 9.5%.









Model	SEM (1)	Direct effects	Indirect effects	Total effects
LnNoPS <- LnQS	14.43***			
	(0.687)			
LnEBPP <- LnQS	14.72***			
	(0.392)			
LnWBPC <- LnQS	9.23***			
	(0.434)			
LnSCPEPC <-LnQS	0.333*	0.333*	0.934***	1.267***
	(0.197)	(0.197)	(0.030)	(0.185)
LnNoPS	0.519***	0.052***		0.052***
	(0.010)	(0.011)		(0.011)
LnEBPP	0.063***	0.064***		0.064***
	(0.024)	(0.024)		(0.024)
LnWBPC	-0.082***	-0.082***		-0.082***
	(0.027)	(0.027)		(0.027)
SCPD	-0.00005***	-0.00005***	0.0008**	0.0008*
	(0.00001)	(0.00001)	(0.0004)	(0.0003)
Election Year			0.696***	0.696***
			(0.110)	(0.110)
LnNoPS * LnQS	-0.079***	-0.079***		-0.079***
	(0.018)	(0.018)		(0.018)
LnEBPP * LnQS	-0.079**	-0.079**		-0.079**
	(0.032)	(0.033)		(0.033)
LnWBPC * LnQS	0.155***	0.155***		0.155***
	(0.044)	(0.044)		(0.044)

Table 1.15: Alternate method- Mediation analyzing using Structural Equation Model- primary enrollment rates

Note: Number of observations = 445. Robust standard errors are clustered at the state level and are reported in parentheses. *, **, *** indicates statistical significance at α = 0.10, 0.05 and 0.01 respectively. Coefficient of determination (CD) = 0.856; Standardized root mean squared residuals (SRMR) = 0.487; R² = 0.856 Refer to APPENDIX FIGURE 1.15 for the corresponding path diagram.

Model	SEM (2)	Direct effects	Indirect effects	Total effects
LnNoUPS <- LnQS	12.54***			
	(0.733)			
LnEBPP <- LnQS	14.83***			
-	(0.372)			
LnWBPC <- LnQS	9.45***			
	(0.402)			
LnSTRUP <- LnQS	5.11***			
	(0.168)			
LnSCUPEPC <-LnQS	0.318	0.318	0.688***	1.005***
	(0.326)	(0.326)	(0.028)	(0.336)
LnNoUPS	0.053**	0.053**		0.053**
	(0.027)	(0.027)		(0.027)
LnEBPP	0.085	0.085		0.085
	(0.055)	(0.055)		(0.055)
LnWBPC	-0.107*	-0.107*		-0.107*
	(0.064)	(0.064)		(0.064)
LnSTRUP	-0.043**	-0.043**		-0.043**
	(0.022)	(0.022)		(0.022)
SCPD	-0.00006***	-0.00006***	0.0007	0.0006
	(0.00002)	(0.00002)	(0.0006)	(0.0006)
LnSCIncPC	-0.040	-0.040		-0.040
	(0.031)	(0.031)		(0.031)
Literacy rate	0.007***	0.007***		0.007***
	(0.001)	(0.001)		(0.001)
Election Year	. ,		0.598	0.598
			(0.477)	(0.477)
LnNoUPS * LnQS	-0.064*	-0.064*		-0.064*
	(0.039)	(0.039)		(0.039)
LnEBPP * LnQS	-0.112*	-0.112*		-0.112*
	(0.066)	(0.066)		(0.066)
LnWBPC * LnQS	0.167*	0.167*		0.167*
	(0.092)	(0.092)		(0.092)

Table 1.16: Alternate method- Mediation analyzing using Structural Equation Model- upper primary enrollment rates

Note: Number of observations = 313. Robust standard errors are clustered at the state level and are reported in parentheses. *, **, *** indicates statistical significance at α = 0.10, 0.05 and 0.01 respectively. Coefficient of determination (CD) = 0.867; Standardized root mean squared residuals (SRMR) = 0.367; R² = 0.867 Refer to APPENDIX FIGURE 1.16 for the corresponding path diagram.