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Stock flow adjustments in sovereign debt dynamics: The role of fiscal frameworks[☆]



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ABSTRACT

We assess, by means of system GMM, how Stock Flow Adjustments (SFA) affect the debt-to-GDP ratio in a panel of 65 countries (covering both developed and emerging and low-income countries) between 1985–2014. In addition, we inspect the role of fiscal rules in affecting SFAs. We find that SFAs positively contribute to the change in the debt-to-GDP ratio with a coefficient close to one. Fiscal rules in general did not lead governments to a systematic use of SFAs to lower budget deficits; however, the existence of fiscal rules with monitor compliance contributes to lower the debt level, although the cyclical deficit partly counteracts this desirable effect. The time period matters: the fall in the debt ratio due to the presence of fiscal rules before the crisis was between 1.7 and 4.2 percent of GDP while after the Global Financial Crisis, revenue and debt-based rules did not contribute to the reduction of debt, which was reinforced with large SFAs.

1. Introduction

The biggest driver of public debt increases is not primary deficits, nor economic activity, nor interest payments. Instead, the main driver is large stock flow adjustments (SFAs), the residual term in a traditional debt decomposition exercise (Jalles, Jaramillo, Mulas-Granados, 2017a). These SFAs can be considered as blind spots in public debt dynamics because they cannot be properly modelled or accurately forecasted (Jaramillo, Kimani and Mulas-Granados, 2017b). Moreover, they are typically associated with a lack of transparency in fiscal accounts (Weber, 2012). Hence, the relevance of assessing to what extent SFAs affect government debt dynamics, which is the main purpose of this paper.

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Many public finance scholars have explored the drivers of debt increases, but the analysis of SFAs in debt dynamics and the relationship between them and fiscal frameworks and institutions, within a fiscal reaction function framework, has received little attention. This is a particularly relevant policy question since fiscal frameworks and institutions – which are meant at constraining the behaviour of governments² – can lead to creative accounting to circumvent such aspects (Milesi-Ferretti, 2003).

Among the initial papers that have studied the role of SFAs on public debt accumulation, the most comprehensive article is the one by Campos, Jaimovich, and Panizza (2006), who assembled a dataset of debt spikes in 117 countries for the period 1972 to 2003. They concluded that debt spikes have little to do with budget deficits, but instead arise from stock flow adjustments, which can be partly explained by contingent liabilities and balance sheet effects. However, they noted that these two components only explain 20 percent of the intra-country variance of SFA, and concluded that there is still much that we do not understand about SFA.

In addition, Abbas et al. (2011) looked at 60 episodes of debt increases between 1880 and 2007 and found that key contributors to public debt surges during non-recessionary periods were both primary deficits and stock-flow adjustments. Finally, Weber (2012), using data for 163 countries between 1980 and 2010, showed that stock-flow adjustments were a significant source of debt increases, while they played only a minor role in explaining debt decreases. SFAs could only be partly explained by balance sheet effects and the realization of contingent liabilities, and significant differences existed in average stock-flow adjustments across countries reflecting country-specific factors. Weber (2012) concluded that fiscal transparency has a major role to play in this area since fiscally transparent countries tend to have a smaller magnitude of SFA in their debt increases.

Using a sample of European Union countries, von Hagen and Wolff (2006) - the paper closest to ours - show how governments use SFA (a form of creative accounting according to their paper) to circumvent the fiscal rules put in place by the European Economic and Monetary Union. They emphasize the need to improve fiscal transparency and reinforce the monitoring of these supranational rules, to reduce off-budget operations.

Against this background, this paper makes several contributions to the literature. First, we run our analysis for the period 1985–2014 encompassing the Global Financial Crisis, after which the relevance of fiscal rules came even more to the forefront of the discussion. Second, while previous literature has largely focused on advanced economies or a sample of European countries, we extend the coverage to include also emerging and low-income countries, an aspect previously unexplored. Third, in addition to inspecting the role of fiscal rules in affecting SFAs, we go deeper in the analysis by looking at different types and design characteristics of those rules. Fourth, for a subsample of countries, we rely on a recent dataset on fiscal institutions (Gupta & Yloutinen, 2014) and inspect which matter the most for the build-up of SFAs.

Our main findings are: (i) SFAs contribute to the change in the level of the debt-to-GDP ratio with a coefficient close to one. (ii) Fiscal rules in general did not led governments to a systematic use SFAs to lower deficits in our country sample. (iii) Countries with some form of macro-fiscal forecasting tool in place have allowed governments to use SFAs to lower deficit. (iv) The existence of fiscal rules in which monitor compliance exist, contributes to lower the debt level, although the cyclical deficit partly counteracts this desirable effect. (v) The magnitude of the fall in the debt ratio due to the presence of fiscal rules before the crisis was between 1.7 and 4.2 percent of GDP. (vi) After the Global and Financial Crisis (GFC), revenue and debt-based rules contributed positively to the accumulation of debt, a fact that is reinforced via large SFAs.

The remainder of the paper is organized as follows. Section 2 presents the key accounting identity on the decomposition of government debt changes. Section 3 provides details on the empirical methodology and data. Section 4 discusses the main empirical results. Section 5 concludes.

2. Accounting identity

The standard equation for decomposing debt changes (see Escolano, 2010 for further details) can be mathematically represented as follows:

$$D_T - D_0 = \sum_{t=1}^T \frac{r_t - G_t}{1 + G_t} D_{t-1} + \sum_{t=1}^T d_t + \sum_{t=1}^T sfa_t. \quad (1)$$

Equation (1) states that the change in the debt-to-GDP ratio ($D_T - D_0$) between time 0 and time T, is the sum of three components: (i) the product of the lagged debt ratio (D_{t-1}) and the difference between the nominal effective interest rate on debt (r_t) and the nominal GDP growth rate (G_t), cumulated over the years under scrutiny; (ii) the cumulative government deficit to GDP (d_t); and (iii) a cumulative stock-flow adjustment (sfa_t) or debt-deficit adjustment term which captures valuation effects and “below-the-line” fiscal-financial operations (for example, financial sector recapitalization,³ or privatizations receipts or the impact of exchange rate changes on foreign denominated debt), as well as errors and omissions.⁴ In von Hagen and Wolff’s (2006) simpler notation, we have:

² These types of mechanisms, watchdogs or rules are introduced to reduce rent-seeking behavior of policy makers, to mitigate common pool problems and prevent undesirable fiscal outcomes (von Hagen, 2002).

³ Note, however that recapitalization might just be recorded “above-the-line” if certain conditions (e.g. insufficient expected return, accumulated losses or negative net worth) are met. Take the case of Dexia’s Group: (http://ec.europa.eu/competition/publications/cpn/2010_2_11.pdf). We thank an anonymous referee for this point.

⁴ This debt decomposition measures only the direct effect of real GDP on the denominator of the debt to GDP ratio. It does not, however, measure the indirect effects of real GDP growth on other subcomponents (such as the primary balance and SFA), which could be significant. For example, Bova, Ruiz-Arnan, Toscani, and Ture (2016) find that realizations of contingent liabilities (often reflected in SFA) tend to occur during periods of economic stress.

$$D_t - D_{t-1} - d_t = sfa_t. \quad (2)$$

A positive SFA means that the stock of government debt has risen between period t and $(t-1)$ by more than the budget deficit recorded in period t . Typical official definitions tend to treat SFA as a statistical residual, which should cancel out over time. However, “large and persistent stock-flow adjustments (especially if they always have a negative impact on debt developments) should give cause for concern, as they may be the result of the inappropriate recording of budgetary operations and can lead to large ex-post upward revisions in deficit levels” (EC, 2003, p. 79).

Naturally, both financial transactions and non-financial transactions are not perfect substitutes but rather should be taken together to better evaluate fiscal dynamics and sustainability.⁵

On the other hand, one should be aware that a higher frequency in the occurrence of SFAs is not necessarily related to an attempt to reduce the budget deficit, but they might be simply the result of liquidity management in turbulent times, and/or the active utilisation of favorable market conditions (see notably, Rybáček, 2015).

3. Empirical methodology and data

3.1. Empirical approach

According to Milesi-Ferretti (2003) fiscal rules (and to our larger purposes, fiscal frameworks and institutions) may induce governments to engage in “bad” or even “ugly” creative accounting. To empirically test this proposition, we study the relation between deficits and SFAs in a large panel of countries between 1985 and 2014.

Looking at equation (2) above, the change of the public debt level in percent of GDP in country i at time t ($\Delta b_{it} = (D_{it} - D_{it-1}) / Y_{it}$) is the sum of SFA in percent of GDP (sfa_t) and the deficit in percent of GDP (d_t). If one takes the following equation:

$$\Delta b_{it} = \alpha_0 + \alpha_1 sfa_{it} + \varepsilon_{it}. \quad (3)$$

Then α_1 is algebraically given by: $\alpha_1 = 1 + \frac{\text{cov}(d_t, sfa_t)}{\text{var}(sfa_t)}$.

Assuming that $\alpha_1 = 1$ implies that the covariance between deficits and SFAs is zero. A coefficient smaller (larger) than one implies a negative (positive) covariance between sfa and d . Borrowing from von Hagen and Wolff (2006), the following reduced-form regression equation will be used to empirically estimate the impact of fiscal frameworks and institutions:

$$\Delta b_{it} = \alpha_0 + \alpha_1 sfa_{it} + \alpha_2 FI_{it} + \alpha_3 sfa_{it} * FI_{it} + \varphi_i + \varepsilon_{it} \quad (4)$$

where FI is our fiscal framework or institution proxy, φ_i are country fixed effects to account for unobserved cross-country heterogeneity and ε_{it} is a disturbance term satisfying standard conditions of zero mean and constant variance. If the hypothesis of no relation between sfa and d holds true and $\alpha_1 = 1$, the coefficient α_3 measures directly the covariance between deficits and SFAs when a given fiscal framework or institution is in place. If $\alpha_3 < 0$ then an increase in the SFA would lower the deficit.

The existence of a fiscal rule (or fiscal institution) should theoretically constrain the debt dynamics, by putting additional oversight on the growth of the stock of government debt. This could be clearer for explicit debt rules, for government spending rules or for budget balance rules. In the case of fiscal revenue rules, that desirable effect might be less obvious since trying to push up revenues might not necessarily bring down the fiscal imbalance, given the uncertainties of notably tax collections, linked partly to economic conditions. Moreover, when fiscal rules are in place, one would expect that the leeway for governments to take advantage of less appropriate recording of budgetary operations be curtailed. Hence, the interactive term in (4) might reflect that effect of a fiscal rule on the development of the stock-flow adjustment itself.

To separate the effects of structural from cyclically adjusted deficits, we run an alternative regression equation, given by:

$$\Delta b_{it} = \beta_0 + \beta_1 d_{it} + \beta_2 FI_{it} + \beta_3 d_{it} * FI_{it} + \varphi_i + \varepsilon_{it}. \quad (5)$$

The treatment effect of FI can be identified by the coefficient β_3 . A negative value of this coefficient means that an increase in deficits leads to a lower SFA as a consequence of the presence of a given FI . Coefficients β_3 and α_3 should have the same sign as they reflect the same covariance. To uncover the effect of the structural and cyclical part of the deficit, equation (5) is augmented as follows:

$$\Delta b_{it} = \gamma_0 + \gamma_1 d_{it}^{ca} + \gamma_2 d_{it}^c + \gamma_3 FI_{it} + \gamma_4 d_{it}^{ca} * FI_{it} + \gamma_5 d_{it}^{ca} * FI_{it} + \varphi_i + \varepsilon_{it} \quad (6)$$

where d_{it}^{ca} is the cyclically adjusted deficit while d_{it}^c denotes the cyclical component. Milesi-Ferretti (2003) model predicts γ_5 to have a larger coefficient than γ_3 as creative accounting is expected to be more strongly used during bad times.

The models described above are reduced-forms and do not allow making causal statements or even quantifying the clean effect of SFAs on debt, meaning that the use of instruments is required. While adding other covariates partly corrects for these biases, endogeneity can still arise from other omitted variables (unobserved heterogeneity and selection effects), measurement errors in variables

⁵ In addition, the statistical methodologies to record financial and non-financial transactions have changed over time. That said, the inclusion of time fixed effects should be able, at least partially and imperfectly, to capture such changes. Naturally, if rules change and improve at each iteration, moving towards an optimal/desirable direction, then room for creative accounting diminishes.

and reverse causality (simultaneity). Since causality can run in both directions, some of the right-hand-side regressors may be correlated with the error term. Our equations are first estimated using Generalized Method of Moments estimator with robust standard errors clustered at the country level. The first-differenced Generalized Method of Moments (GMM) estimator can behave poorly if time series are persistent (which is the case for debt). Hence, we use the more efficient system GMM estimator that exploits stationarity restrictions. This method jointly estimates Equation (6) in first differences, using as instruments lagged levels of the dependent and independent variables, and in levels, using as instruments the first differences of the regressors (Arellano & Bover, 1995; Blundell & Bond, 1998).⁶ GMM estimators are unbiased, and compared with ordinary least squares or fixed effects (within-group) estimators, exhibit the smallest bias and variance (Arellano and Bond, 1991).⁷

As robustness checks we also employ alternative estimators. More specifically, we rely on pooled Ordinary Least Squares, panel within-group estimator and bias-corrected least-squares dummy variable (LSDV-C) estimator by Bruno (2005).⁸

3.2. Data and stylized facts

Our sample, for which the macro data come from the IMF World Economic Outlook (WEO) database, covers 65 countries observed over the period 1985–2014. We also rely on IMF's WEO measures of the cyclically adjusted balance (deficit) and use it to construct the structural component and the difference with the unadjusted balance (deficit).

We use equation (2) similarly to notably von Hagen and Wolff (2006) to compute the data on the SFA. Following this method, we compared the final debt level as of 2014 with the accumulated deficits (that is, the sum of the debt level at the first year of available data – which may differ from country to country – and all budget deficits between that first year and 2014, as a percentage of 2014 GDP). These computations are displayed in Fig. 1 for all countries covered in our analysis. It shows that most countries have regularly had positive, and in some cases, quite large SFAs over time. For instance, Finland and Luxembourg have 68 and 41 percentage points of GDP more debt than their budget data suggest, respectively. SFAs are negative mainly in Eastern European Countries.

In addition, Fig. 2 exemplifies the accounting identity for some countries, on a yearly basis, where we show the specific relevance and magnitude of the stock flow adjustment, which tend to be rather persistent over time in many cases, therefore blurring to some extent the link between (primary) budget balances and government debt dynamics.

For fiscal rules as well as their types and characteristics, we make use of the datasets created by the IMF. The first dataset was introduced by Schaechter, Kinda, Budina, and Weber (2012) and its most recent available updates discussed in detail by Lledó, Yoon, Fang, Mbaye, and Young (2017). The rules are classified according to the following typology: expenditure rules (ER), revenue rules (RR), budget balance rules (BBR) and debt rules (DR). Additionally, we created a dummy variable FR_I , denoting the existence of any of these fiscal rules in a given country in a given year. Moreover, the dataset contains information on such features of the rules as existing escape clauses, enforcement procedures or independent monitoring councils or their transparency.

In the analysis, we include 65 countries, which had at least one of the rules in place during the period of analysis. Overall, during the 31 years of the timespan at least one rule in place was observed in 1076 cases (on 2015 possible), the most frequent being the budget balance rule (974 cases), followed by debt rule (772 occurrences), expenditure rule (399), the least frequent being the revenue rule (186). Only a handful of countries (Germany, Indonesia, Japan, Malaysia, and Singapore) had at least one rule in place for the entire time span. In all of these cases, it was the balanced budget rule, additionally completed by an expenditure rule (for Germany) and debt rule (in Malaysia). If a given rule was in place, the debt rule was present in a given country for almost 16.5 years, balanced budget rule for 15.7 years, revenue rule for 13.3 years (but it was present only in 13 countries) and expenditure rule for 9.7 years. The dataset additionally contains information about monitoring, enforcement and escape clause for each type of rules. We use this data on somewhat more aggregate level, i.e., if any of the fiscal rules applied in a country had a monitoring of compliance in place, the variable $FR_monitor$ assumes value 1 and zero otherwise. The same is the case for formal enforcement procedure and escape clauses whereas independent monitoring body and transparency are taken “as they are” from the IMF database.⁹

As far as fiscal rules are concerned we can plot the absolute number of new rules (of any type) over time by income group, and we get

⁶ We equally tried estimating the key equations with a difference GMM estimator but decided against it because the lagged dependent variable was not significant. Moreover, the tenor of the results is very similar to the system GMM. More specifically, we run the two-step system-GMM estimator with Windmeijer standard errors. The significance of the results is robust to different choices of instruments and predetermined variables.

⁷ As far as information on the choice of lagged levels (differences) used as instruments in the difference (level) equation, as work by Bowsher (2002) and, more recently, Roodman (2009) have indicated, when it comes to moment conditions (as thus to instruments) more is not always better. The GMM estimators are likely to suffer from “overfitting bias” once the number of instruments approaches (or exceeds) the number of groups/countries (as a simple rule of thumb). In the present case, the validity of instruments was examined using Hansen's test of overidentifying restrictions. Intuitively, the system GMM estimator does not rely exclusively on the first-differenced equations, but exploits also information contained in the original equations in levels.

⁸ Kiviet (1995) used asymptotic expansion techniques to approximate the small sample bias of the standard LSDV estimator for samples where N is small or only moderately large. Bruno (2005) extended the bias approximation formulas to accommodate unbalanced panels with a strictly exogenous selection rule.

⁹ The most frequent and relatively persistent design feature is the existence of an enforcement mechanism, which was in place in 28 countries on average for slightly more than 10 years. Marginally least popular is monitoring (25 countries, on average in place for 9.6 years), Transparency requirements were present in 21 countries, notably on average for the longest period, i.e. for almost 11 years. Independent monitoring body was in place in 22 countries, but as a relatively recent mechanism, its average duration only slightly exceeds 5 years. Finally, some form of escape clause is present in 12 countries, on average for 7.5 years.

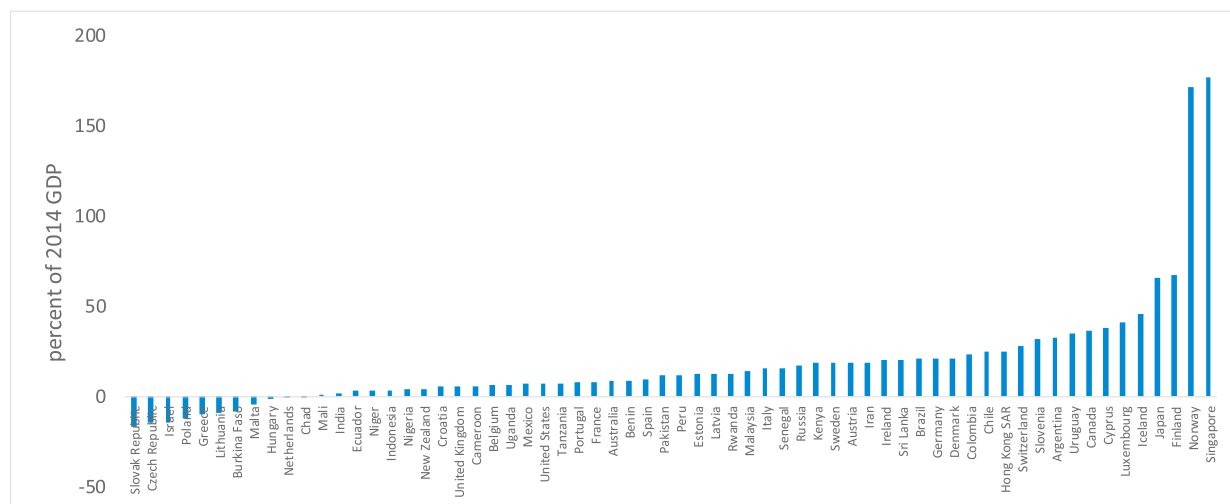


Fig. 1. Total Stock Flow Adjustments in percent of 2014 GDP by country.

the pattern observed in Fig. 3. Looking at Advanced Economies, while countries have implemented fiscal rules since the mid-1980s, most of them followed the Maastricht Treaty in 1992 (in adherence to the EU convergence criteria) as well as after the Global Financial Crisis. In non Advanced Economies, the absolute number of fiscal rules is lower than the advanced economies sample, and most of them were implemented starting in the early 2000s.

Gupta and Yloutinen (2014) made available another dataset on fiscal rules, which we also use. They analyse fiscal institutional frameworks in G-20 economies complemented by six low-income countries (Kenya, Mozambique, Myanmar, Uganda, Vietnam and Zambia). In particular, aspects covered in this database include: fiscal reporting (*fr*), macro fiscal forecasting (*mf*), independent fiscal agency (*ifa*), fiscal objectives (*fo*), medium term budget framework (*mbf*), budget execution (*be*), understanding the scale and scope of the fiscal challenge (understanding), developing a credible fiscal strategy (*developing*) and implementing the fiscal strategy through the budget process (*implementing*). Except for *ifa*, which is present only in 17 out of the 26 countries, all of these institutions are to a smaller or larger extent present in at least 24 countries.¹⁰

4. Empirical results

4.1. Baseline with fiscal rules

We start by estimating equations (4) and (5) for the different fiscal rules proxies using the entire sample of countries and time span. Results are displayed in Tables 1 and 2 respectively. Starting with Table 1, we observe that SFAs, as the accounting identity suggests, contribute to the change in the level of the debt-to-GDP ratio with a coefficient close to one. The existence of a fiscal rule leads of a fall in the debt level but the coefficient estimate is not statistically different from zero (specification 1). Since the coefficient on the SFA is not statistically different from one, the estimated coefficients on the interaction terms represent the covariances between the SFA and the deficit when a specific rule or a rule with a certain characteristic is in place. In the cases of the rules *FRmonitor*, *indepmonit* and *transp* the interaction term comes out positive but with a small magnitude and statistically significant at the 10 percent level. Contrary to the evidence presented in von Hagen and Wolff (2006) - who used a different definition for the fiscal rule dummy variable, closer to an event study approach, for a smaller country group - these findings suggest that fiscal rules in general did not led governments to systematically use SFAs to lower deficits in our sample of 65 countries. Moreover, in Table 2 we still have that fiscal rules do not statistically significantly affect the change in the debt level. As in Table 1, a positive covariance appears. An increase in the deficit by one percentage point is associated with an increase of the SFA by an amount between [0.4–1.3] depending on the fiscal rule proxy under scrutiny.

As far as diagnostics are concerned, we can see that they are satisfactory. We first examine the serial correlation of the error structure. We reject first order but do not reject second order serial correlation.¹¹ Regarding the first stage diagnostics, the Hansen Test indicates that the instruments in the system GMM are correctly excluded. Though system GMM does perform better in the presence of persistence of series it may suffer from a drawback in that the additional instruments brought about in the process may overfit the model and may not be able to remove the endogenous component from the estimation (Roodman, 2009). Still, the Hansen test for overall exclusion of the instrument, with somewhat high p-values, might be seen as a caveat regarding identification.

¹⁰ Descriptive statistics are presented in the Appendix.

¹¹ Recall that the null hypothesis is that the errors exhibit no second-order serial correlation.

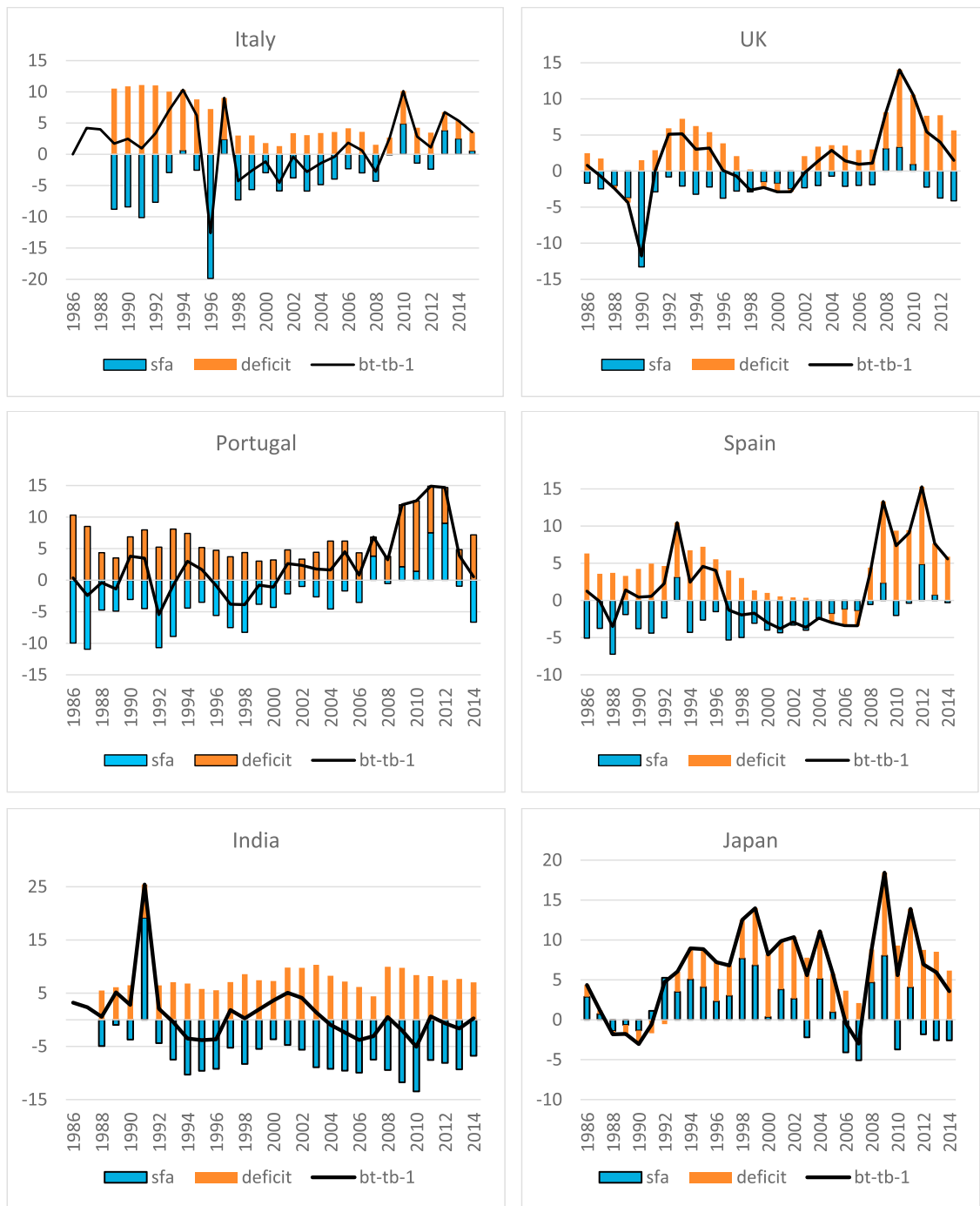


Fig. 2. Accounting identity (% of GDP). Note: “sfa” denotes stock flow adjustments; “deficit” denotes the budget deficit; “bt-tb-1” denotes the change in the public debt level. Source: authors’ calculations.

4.2. Baseline with fiscal institutions

In Table 3 we re-estimate equations (4) and (5) using the smaller sample stemming from Gupta and Yloutinen (2014) on fiscal frameworks. Still we obtain that in general fiscal rules did not significantly change the debt level in this group of 26 countries. Those countries containing a form of macro-fiscal forecasting tool (*mf*) seemed to have allowed governments to use SFAs to lower deficits (the interaction term is negative and statistically significant at the 5 percent level).

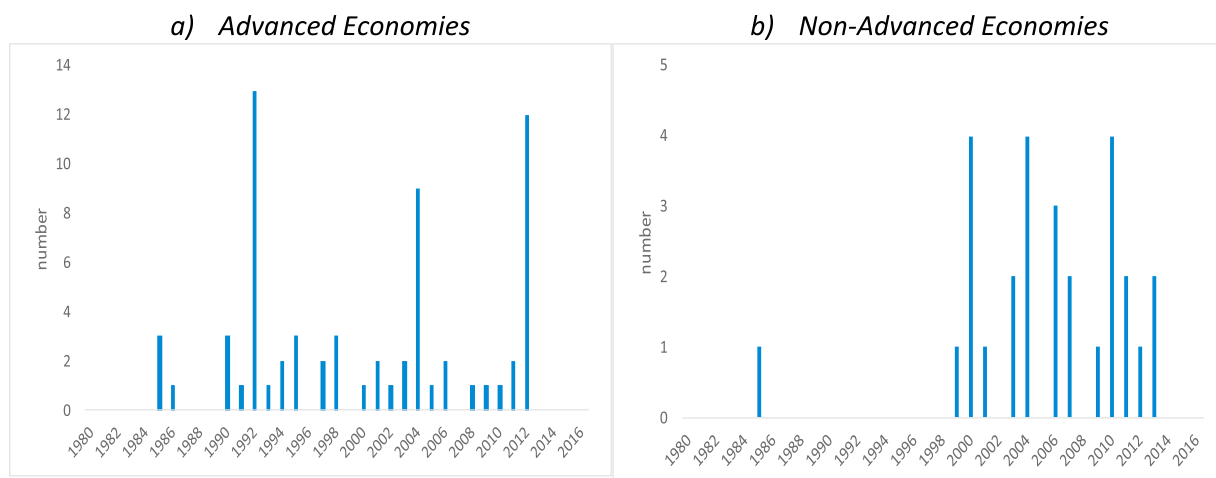


Fig. 3. Distribution of New Fiscal Rules implemented over time by Income GroupSource: International Monetary Fund's fiscal rule dataset.

A better and improved expectational view of the economy and public finances, suggests that SFAs have become a policy variable to control the deficit in countries where such framework was in place. All other characteristics of the fiscal system do not seem to matter neither individually nor in conjunction with the SFA variable. Results from estimating equation (5) in this setting yield statistically insignificant coefficients and, as hence, are omitted for reasons of parsimony (but available upon request).

4.3. Augmented version with fiscal rules

In order to assess to what extent the cyclical component of the budget balance plays a role in the analysis, we have used the cyclical and cyclically adjusted parts of the budget deficit separately. Therefore, capturing the structural and cyclical dimensions of the deficit leads to the estimation of equation (6). Looking at Table 4 for the case of fiscal rule and our large expanded sample, both components of the deficit positively clearly affect the debt level, particularly the structural part. While expenditures rules seem to negatively affect the debt level (despite not being statistically significantly different from zero), its interaction with the cyclical component of the deficit yields a positive and statistically significant coefficient. This means that an increase in the cyclical deficit when expenditure rules are in place lead to an increase in debt. A similar conclusion is also true with regard to the debt rule, but in this case, the impact of both deficit components is similarly positive. Finally, the existence of fiscal rules in which monitor compliance exist, contributes to lower the debt level, but the cyclical deficit partly counteracts this desirable effect.

4.4. Country sample and time split

As a first sensitivity exercise, we split our large heterogeneous sample between OECD and non-OECD countries. We observe in Table 5 – estimating equation (4) – that SFAs seem to be more important in positively changing the debt level in non-OECD countries (where the coefficient estimates are closer to 1 vis-à-vis the OECD). In addition, we see that in the OECD sub-sample the existence of debt rules lead to a rise in the debt level, an effect that is exacerbated when coupled with SFAs.

In addition, in Table 6 in turn, we split the period before and after the Global Financial Crisis (in 2008). The crisis was indeed a major structural break in the sense that before, most fiscal rules seem to lower debt levels and had a clear actively constraining role in keeping government debt from rising. The order of magnitude of the fall in the debt level due to the presence of fiscal rules before the crisis was between 1.7 and 4.2 percent of GDP. After the crisis, both revenue-based and debt-based rules starting contributing positively to the accumulation of debt, a fact that is reinforced with the existence of large SFAs. Indeed, one can hypothesise that although fiscal rules are in place, the fact the GFC put an extraordinary pressure on public accounts, simply nullified the contribution of such rules to the containment of fiscal imbalances and fostered SFAs.

4.5. Other robustness exercises

Our final exercise relates to stress-testing our previous results to alternative estimators. In particular, we run a pooled OLS, a within fixed effects estimator and the bias corrected least squares dummy variable. Results in Table 7 confirm the relevance of SFAs for the change in government debt. Without accounting for potential endogeneity, we get the nice result that the simple existence of fiscal rules lowers the public debt level (specifications 1, 4, 7). In addition, as before, both components of the deficit positively affect public debt, with the positive effect of the structural component being reinforced when fiscal rules are present.

In addition to the robustness check to alternative estimators, we also employed an alternative method to extract the structural and cyclical components of the budget deficit. In particular, instead of relying on the IMF's WEO measure of output gap, we rather apply the recent filtering technique developed by Hamilton (2017). Once the output gap is obtained, we then used it to get a new measure of the

Table 1
Baseline, all countries, fiscal rules, equation (4), system GMM.

| Specification | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
|-----------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|
| Regressors | | | | | | | | | | |
| lagged debt | 0.16** (0.070) | 0.15** (0.072) | 0.16** (0.075) | 0.16** (0.070) | 0.16** (0.070) | 0.10* (0.057) | 0.12** (0.052) | 0.15** (0.058) | 0.11** (0.051) | 0.11** (0.050) |
| SFA | 0.83*** (0.098) | 0.93*** (0.052) | 0.89*** (0.093) | 0.80*** (0.104) | 0.77*** (0.123) | 0.92*** (0.053) | 0.92*** (0.046) | 0.96*** (0.030) | 0.91*** (0.051) | 0.91*** (0.051) |
| FR_1 | -0.11 (0.713) | | | | | | | | | |
| SFA x_FR | 0.13 (0.111) | | | | | | | | | |
| ER | | 0.47 (0.923) | | | | | | | | |
| SFA x_ER | | 0.04 (0.105) | | | | | | | | |
| RR | | | -0.04 (1.180) | | | | | | | |
| SFA x RR | | | 0.10 (0.108) | | | | | | | |
| BBR | | | | 0.20 (0.736) | | | | | | |
| SFA xBBR | | | | 0.16 (0.119) | | | | | | |
| DR | | | | | 1.85 (1.769) | | | | | |
| SFA x DR | | | | | 0.26* (0.148) | | | | | |
| FR_monitor | | | | | | 1.06 (1.011) | | | | |
| SFA x_FRmonitor | | | | | | -0.03 (0.157) | | | | |
| FR_enforce | | | | | | | 0.45 (0.828) | | | |
| SFA x FRenforce | | | | | | | -0.01 (0.140) | | | |
| FR_escape | | | | | | | | 0.75 (2.401) | | |
| SFA x FRescapte | | | | | | | | -0.57 (0.626) | | |
| Independent_monitor | | | | | | | | | 1.08 (1.244) | |
| SFA x indepmonit | | | | | | | | | 0.07** (0.028) | |
| Transparency | | | | | | | | | | -0.42 (0.987) |
| SFA x transp | | | | | | | | | | 0.04* (0.068) |
| Observations | 1336 | 1336 | 1336 | 1336 | 1336 | 1140 | 1378 | 865 | 1353 | 1378 |
| Hansen test (p-value) | 0.72 | 0.72 | 0.68 | 0.74 | 0.67 | 0.72 | 0.65 | 0.89 | 0.64 | 0.70 |
| ar2p | 0.17 | 0.43 | 0.17 | 0.11 | 0.02 | 0.91 | 0.87 | 0.58 | 0.99 | 0.86 |
| ar1p | 0.v04 | 0.05 | 0.05 | 0.04 | 0.05 | 0.05 | 0.05 | 0.10 | 0.05 | 0.05 |

Note: Dependent variable is the change in debt in percent of GDP. Robust standard errors clustered at the country level in parenthesis. The Hansen test evaluates the validity of the instrument set, i.e., tests for over-identifying restrictions. AR(1) and AR(2) are the Arellano-Bond autocorrelation tests of first and second order (the null is no autocorrelation), respectively. A constant term has been estimated but it is not reported for reasons of parsimony. *, **, *** denote statistical significance at the 10, 5 and 1 percent levels, respectively. "FR_1" if a country has at least one fiscal rule; "ER" = expenditure rule in place; "RR" revenue rule in place; "DR" = debt rule in place; "BBR" = budget balance rule in place, "monitor" = at least one of the rules in place monitor compliance exist; "enforce" = at least one of the rules in place formal enforcement procedure exist; "escape" at least of the rules in place escape clause exist. "Independent_monitor" = an independent body monitors implementation of the rules. "transparency" = Fiscal Responsibility Laws are in place ensuring transparency and accountability.
Fiscal rules dataset from [Ledó et al. \(2017\)](#).

Table 2
Baseline, all countries, fiscal rules, equation (5), system GMM.

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
|-----------------------|--------------------|--------------------|-------------------|--------------------|--------------------|-------------------|-------------------|--------------------|-------------------|-------------------|
| lagged debt | 0.28*** (0.091) | 0.26*** (0.084) | 0.28** (0.115) | 0.29*** (0.095) | 0.31*** (0.098) | 0.24** (0.117) | 0.28** (0.109) | 0.01 (0.238) | 0.30** (0.117) | 0.27** (0.102) |
| deficit | 0.89* (0.455) | 0.57*** (0.206) | 0.51** (0.200) | 0.83* (0.455) | 0.57** (0.261) | 0.09 (0.160) | 0.30** (0.139) | 0.60*** (0.151) | 0.13 (0.141) | 0.11 (0.118) |
| FR_1 | 3.46 (2.237) | | | | | | | | | |
| def_FR | -0.60 (0.483) | | | | | | | | | |
| ER | | 2.33 (1.686) | | | | | | | | |
| def_ER | | -0.14 (0.278) | | | | | | | | |
| RR | | | 4.54 (3.297) | | | | | | | |
| def_RR | | | -1.21 (0.901) | | | | | | | |
| BBR | | | | 3.09 (2.193) | | | | | | |
| def_BBR | | | | -0.60 (0.516) | | | | | | |
| DR | | | | | 0.99 (1.924) | | | | | |
| def_DR | | | | | -0.42* (0.223) | | | | | |
| FR_monitor | | | | | | 2.97 (3.221) | | | | |
| def_FRmonitor | | | | | | 1.36* (0.842) | | | | |
| FR_enforce | | | | | | | 8.86 (5.930) | | | |
| def_FRenforce | | | | | | | -1.46 (1.953) | | | |
| FR_escape | | | | | | | | 0.15 (3.397) | | |
| def_FRescapte | | | | | | | | -0.05 (0.971) | | |
| Independent_monitor | | | | | | | | | 5.33 (4.872) | |
| def_indepmonit | | | | | | | | | 1.16* (0.717) | |
| Transparency | | | | | | | | | | -0.55 (2.987) |
| def_transp | | | | | | | | | | 1.02 (0.948) |
| Observations | 1336 | 1336 | 1336 | 1336 | 1336 | 1140 | 1378 | 865 | 1353 | 1378 |
| Hansen test (p-value) | 0.84 | 0.85 | 0.62 | 0.85 | 0.79 | 0.78 | 0.69 | 0.69 | 0.59 | 0.73 |
| ar2p | 0.45 | 0.51 | 0.31 | 0.46 | 0.39 | 0.87 | 0.83 | 0.67 | 0.95 | 0.91 |
| ar1p | 0.01 | 0.02 | 0.01 | 0.01 | 0.01 | 0.02 | 0.03 | 0.01 | 0.03 | 0.02 |

Note: Dependent variable is the change in debt in percent of GDP. Robust standard errors clustered at the country level in parenthesis. The Hansen test evaluates the validity of the instrument set, i.e., tests for over-identifying restrictions. AR(1) and AR(2) are the Arellano-Bond autocorrelation tests of first and second order (the null is no autocorrelation), respectively. A constant term has been estimated but it is not reported for reasons of parsimony. *, **, *** denote statistical significance at the 10, 5 and 1 percent levels, respectively. “FR_1” if a country has at least one fiscal rule; “ER” = expenditure rule in place; “RR” revenue rule in place; “DR” = debt rule in place; “BBR” = budget balance rule in place, “monitor” = at least one of the rules in place monitor compliance exist; “enforce” = at least one of the rules in place formal enforcement procedure exist; “escape” at least of the rules in place escape clause exist. “Independent_monitor” = an independent body monitors implementation of the rules. “transparency” = Fiscal Responsibility Laws are in place ensuring transparency and accountability.

Fiscal rules dataset from [Lledó et al. \(2017\)](#).

cyclically adjusted balance. Reflecting the fact that the elasticity of government revenues (*REV*) to output growth is close to one while expenditure (*EXP*) is largely inelastic to growth ([Girouard & André, 2005](#)), we multiply government revenues by the factor $[1/(1 + OG/100)]$ to get *REV_adj* (revenue adjusted), with *OG* being the output gap obtain via the Hamilton filter. Then $CAB = REV_adj - EXP$.¹²

¹² Both revenues and expenditures come from the IMF WEO database.

Table 3
Baseline, all countries, fiscal institutions, equation (4), system GMM.

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|-----------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|
| lagged debt | 0.06 (0.040) | 0.06 (0.042) | 0.06 (0.042) | 0.06 (0.044) | 0.06 (0.040) | 0.04 (0.043) | 0.06 (0.041) | 0.06 (0.040) | 0.04 (0.039) |
| SFA | 0.99*** (0.049) | 1.01*** (0.020) | 1.01*** (0.017) | 1.03*** (0.013) | 1.00*** (0.022) | 1.02*** (0.119) | 1.00*** (0.030) | 1.00*** (0.035) | 0.99*** (0.096) |
| fr | −0.78 (1.680) | | | | | | | | |
| SFA_fr | 0.03 (0.068) | | | | | | | | |
| mf | | −0.76 (1.401) | | | | | | | |
| SFA_mf | | −0.01 (0.070) | | | | | | | |
| ifa | | | −0.59 (2.149) | | | | | | |
| SFA_ifa | | | −0.04 (0.183) | | | | | | |
| fo | | | | −0.78 (0.792) | | | | | |
| SFA_fo | | | | −0.05 (0.034) | | | | | |
| mbf | | | | | −1.49 (1.335) | | | | |
| SFA_mbf | | | | | 0.05 (0.085) | | | | |
| be | | | | | | 8.94** (4.238) | | | |
| SFA_be | | | | | | −0.01 (0.085) | | | |
| understanding | | | | | | | −0.59 (1.750) | | |
| SFA_understanding | | | | | | | 0.01 (0.090) | | |
| developing | | | | | | | | −1.07 (1.606) | |
| SFA_developing | | | | | | | | 0.03 (0.071) | |
| implementing | | | | | | | | | 4.96 (5.098) |
| SFA_implementing | | | | | | | | | 0.01 (0.073) |
| Observations | 562 | 562 | 562 | 562 | 562 | 562 | 562 | 562 | 562 |
| Hansen test (p-value) | 0.55 | 0.87 | 0.65 | 0.74 | 0.52 | 0.79 | 0.84 | 0.81 | 0.58 |
| ar2p | 0.15 | 0.17 | 0.14 | 0.25 | 0.31 | 0.24 | 0.41 | 0.19 | 0.23 |
| ar1p | 0.00 | 0.01 | 0.01 | 0.03 | 0.06 | 0.01 | 0.02 | 0.02 | 0.04 |

Note: Dependent variable is the change in debt in percent of GDP. LDV denotes lagged dependent variable. Robust standard errors clustered at the country level in parenthesis. The Hansen test evaluates the validity of the instrument set, i.e., tests for over-identifying restrictions. AR(1) and AR(2) are the Arellano-Bond autocorrelation tests of first and second order (the null is no autocorrelation), respectively. A constant term has been estimated but it is not reported for reasons of parsimony. *, **, *** denote statistical significance at the 10, 5 and 1 percent levels, respectively. “fr” = fiscal reporting; “mf” = macro fiscal forecasting; “IFA” = independent fiscal agency; “fo” fiscal objectives; “MBF” medium term budget framework; “be” budget execution; “understanding” = understanding the scale and scope of the fiscal challenge; “developing” = developing a credible fiscal strategy; “implementing” = implementing the fiscal strategy through the budget process.

Fiscal rules dataset from [Gupta and Yloutinen \(2014\)](#).

The results from re-estimating equation (6) with system GMM and the new structural and cyclical deficit variables (and interaction terms) are not qualitatively different from the ones previously discussed (available upon request).

In addition, recognizing that other macro variables – not explicitly accounted for in equations (3) and (4) – may affect debt dynamics, we have re-estimated [Table 1](#) adding the inflation rate and the output gap (from the IMF WEO) as regressors. We have narrowed the macro variables to this set since they are the more potentially linked ones to the debt ratio dynamics. We do not add GDP explicitly as it enters as the denominator of fiscal variables (commonly used in ratios) and to avoid collinearity problems. In addition, by including country fixed effects we are (even if imperfectly) accounting for such factors. That being said, we have re-estimated the baseline regression by including explicitly a role for the price level (inflation) – known to affect fiscal dynamics – and economic slack (measured by the output gap). To reduce potential reverse causality these regressors enter lagged.

Table 4
Baseline Augmented, all countries, fiscal rules, equation (6), system GMM.

| Specification | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
|---------------------|--------------------|--------------------|--------------------|--------------------|--------------------|---------------------|--------------------|--------------------|--------------------|-------------------|
| lagged debt | 0.24*** (0.063) | 0.20*** (0.070) | 0.26*** (0.068) | 0.27*** (0.057) | 0.25*** (0.065) | 0.23*** (0.077) | 0.27*** (0.063) | 0.13 (0.089) | 0.27*** (0.084) | 0.19 (0.121) |
| cadef | 0.61*** (0.189) | 0.82*** (0.208) | 0.64*** (0.215) | 0.59*** (0.170) | 0.51* (0.279) | 0.46** (0.187) | 0.66*** (0.180) | 0.89*** (0.247) | 0.53*** (0.171) | 0.50** (0.191) |
| cyclicaldef_weo | 0.41** (0.194) | 0.27* (0.157) | 0.24** (0.114) | 0.41** (0.174) | 0.24 (0.168) | 0.20* (0.113) | 0.21 (0.132) | 0.41** (0.166) | 0.22* (0.124) | 0.18 (0.141) |
| FR_1 | 1.27 (1.349) | | | | | | | | | |
| cadef_FR | 0.14 (0.264) | | | | | | | | | |
| cycdef_FR | -0.22 (0.177) | | | | | | | | | |
| ER | | 2.59 (2.325) | | | | | | | | |
| cadef_ER | | -0.24 (0.458) | | | | | | | | |
| cycdef_ER | | 1.38** (0.523) | | | | | | | | |
| RR | | | -2.90 (2.679) | | | | | | | |
| cadef_RR | | | 0.72 (0.701) | | | | | | | |
| cycdef_RR | | | -0.48 (1.056) | | | | | | | |
| BBR | | | | 2.01 (1.508) | | | | | | |
| cadef_BBR | | | | 0.06 (0.283) | | | | | | |
| cycdef_BBR | | | | -0.28 (0.202) | | | | | | |
| DR | | | | | -0.16 (1.500) | | | | | |
| cadef_DR | | | | | 0.39* (0.241) | | | | | |
| cycdef_DR | | | | | 0.24* (0.153) | | | | | |
| FR_monitor | | | | | | -0.77*** (0.149) | | | | |
| cadef_FRmonitor | | | | | | 0.98* (0.561) | | | | |
| cycdef_FRmonitor | | | | | | 2.27** (1.041) | | | | |
| FR_enforce | | | | | | | 3.05 (1.979) | | | |
| cadef_FRenforce | | | | | | | 0.49 (0.682) | | | |
| cycdef_FRenforce | | | | | | | 0.17 (1.128) | | | |
| FR_escape | | | | | | | | 3.76* (1.861) | | |
| cadef_FRescape | | | | | | | | -0.24 (0.600) | | |
| cycdef_FRescape | | | | | | | | 0.70 (1.131) | | |
| Independent_monitor | | | | | | | | | -0.98 (2.164) | |
| cadef_indepmonit | | | | | | | | | -2.47 (2.217) | |
| cycdef_indepmonit | | | | | | | | | 8.10 (8.429) | |
| Transparency | | | | | | | | | | 1.75 (2.031) |
| cadef_transp | | | | | | | | | | 2.05 (1.887) |
| cycdef_transp | | | | | | | | | | 4.34 (4.692) |

(continued on next page)

Table 4 (continued)

| Specification | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
|-----------------------|------|------|------|------|------|------|------|------|------|------|
| Observations | 933 | 933 | 933 | 933 | 933 | 776 | 933 | 622 | 920 | 933 |
| Hansen test (p-value) | 0.78 | 0.65 | 0.84 | 0.67 | 0.68 | 0.79 | 0.86 | 0.89 | 0.80 | 0.76 |
| ar2p | 0.26 | 0.29 | 0.26 | 0.28 | 0.26 | 0.26 | 0.28 | 0.28 | 0.13 | 0.18 |
| ar1p | 0.08 | 0.09 | 0.07 | 0.09 | 0.07 | 0.09 | 0.08 | 0.12 | 0.04 | 0.03 |

Note: Dependent variable is the change in debt in percent of GDP. Robust standard errors clustered at the country level in parenthesis. The Hansen test evaluates the validity of the instrument set, i.e., tests for over-identifying restrictions. AR(1) and AR(2) are the Arellano-Bond autocorrelation tests of first and second order (the null is no autocorrelation), respectively. A constant term has been estimated but it is not reported for reasons of parsimony. *, **, *** denote statistical significance at the 10, 5 and 1 percent levels, respectively. “FR_1” if a country has at least one fiscal rule; “ER” = expenditure rule in place; “RR” revenue rule in place; “DR” = debt rule in place; “BBR” = budget balance rule in place, “monitor” = at least one of the rules in place monitor compliance exist; “enforce” = at least one of the rules in place formal enforcement procedure exist; “escape” at least of the rules in place escape clause exist. “Independent_monitor” = an independent body monitors implementation of the rules. “transparency” = Fiscal Responsibility Laws are in place ensuring transparency and accountability.

Fiscal rules dataset from [Lledó et al. \(2017\)](#).

Table 5

Baseline, OECD vs non-OECD, fiscal rules, equation (4), system GMM.

| Specification | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
|-----------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|
| Sample | OECD | | | | | Non-OECD | | | | |
| lagged debt | 0.42*** (0.072) | 0.40*** (0.062) | 0.41*** (0.078) | 0.42*** (0.072) | 0.39*** (0.071) | 0.08** (0.039) | 0.07 (0.040) | 0.09* (0.042) | 0.08** (0.038) | 0.08* (0.042) |
| SFA | 0.60*** (0.133) | 0.50*** (0.152) | 0.58*** (0.190) | 0.67*** (0.231) | 0.49** (0.210) | 0.92*** (0.078) | 0.96*** (0.033) | 1.01*** (0.030) | 0.91*** (0.076) | 0.98*** (0.045) |
| FR_1 | 0.32 (0.715) | | | | | -1.27 (1.018) | | | | |
| SFA_FR | -0.04 (0.195) | | | | | 0.11 (0.102) | | | | |
| ER | | 0.88 (1.598) | | | | | -2.35 (1.567) | | | |
| SFA_ER | | 0.20 (0.212) | | | | | 0.15* (0.093) | | | |
| RR | | | 0.46 (3.085) | | | | | -2.63** (1.137) | | |
| SFA_RR | | | -0.14 (0.852) | | | | | 0.04 (0.077) | | |
| BBR | | | | 0.06 (0.992) | | | | | -1.54 (0.945) | |
| SFA_BBR | | | | -0.15 (0.185) | | | | | 0.13 (0.106) | |
| DR | | | | | 1.80** (0.934) | | | | | -1.12 (0.892) |
| SFA_DR | | | | | 0.38** (0.211) | | | | | 0.05 (0.081) |
| Observations | 822 | 822 | 822 | 822 | 822 | 514 | 514 | 514 | 514 | 514 |
| Hansen test (p-value) | 0.77 | 0.81 | 0.73 | 0.61 | 0.65 | 0.74 | 0.82 | 0.59 | 0.72 | 0.70 |
| ar2p | 0.16 | 0.15 | 0.14 | 0.20 | 0.32 | 0.31 | 0.57 | 0.47 | 0.30 | 0.43 |
| ar1p | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.04 | 0.04 | 0.05 | 0.04 | 0.05 |

Note: Dependent variable is the change in debt in percent of GDP. Robust standard errors clustered at the country level in parenthesis. The Hansen test evaluates the validity of the instrument set, i.e., tests for over-identifying restrictions. AR(1) and AR(2) are the Arellano-Bond autocorrelation tests of first and second order (the null is no autocorrelation), respectively. A constant term has been estimated but it is not reported for reasons of parsimony. *, **, *** denote statistical significance at the 10, 5 and 1 percent levels, respectively. “FR_1” if a country has at least one fiscal rule; “ER” = expenditure rule in place; “RR” revenue rule in place; “DR” = debt rule in place; “BBR” = budget balance rule in place.

Fiscal rules dataset from [Lledó et al. \(2017\)](#).

The results (see Appendix [Table A2](#)) show that inflation has no statistically significant effect on debt change, while the output gap comes out negative and significant contributing to lower government debt. All other variables remain qualitatively similar, therefore, not changing previous discussion of the key results. Note also, that the inclusion of country fixed effects partially accounts for this other country-specific idiosyncracies and structural aspects.¹³

¹³ We thank an anonymous referee for raising this point.

Table 6

Baseline, before vs after GFC, fiscal rules, equation (4), system GMM.

| Specification | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
|-----------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|
| Time period | Before GFC | | | | | After GFC | | | | |
| lagged debt | 0.07* (0.042) | 0.06 (0.044) | 0.07 (0.045) | 0.07* (0.042) | 0.06 (0.044) | 0.59*** (0.137) | 0.54*** (0.113) | 0.53*** (0.110) | 0.54*** (0.111) | 0.51*** (0.109) |
| SFA | 0.86*** (0.100) | 0.92*** (0.053) | 0.94*** (0.086) | 0.90*** (0.095) | 0.95*** (0.082) | −0.28 (1.010) | 0.51** (0.242) | 0.59** (0.289) | 0.53 (0.408) | 0.52 (0.370) |
| FR_1 | −1.48* (0.793) | | | | | −0.66 (2.166) | | | | |
| SFA_FR | 0.12 (0.118) | | | | | 1.36*** (0.123) | | | | |
| ER | | −4.79** (1.615) | | | | | −1.20 (1.433) | | | |
| SFA_ER | | 0.14 (0.089) | | | | | 0.72* (0.303) | | | |
| RR | | | −0.56 (2.273) | | | | | 1.51** (0.742) | | |
| SFA_RR | | | 0.05 (0.080) | | | | | 0.25 (0.512) | | |
| BBR | | | | −1.12** (0.351) | | | | | −2.02 (1.668) | |
| SFA_BBR | | | | 0.07 (0.108) | | | | | 0.34 (0.644) | |
| DR | | | | | 0.33 (1.425) | | | | | 2.54*** (0.541) |
| SFA_DR | | | | | 0.01 (0.074) | | | | | 0.60* (0.342) |
| Observations | 888 | 888 | 888 | 888 | 888 | 448 | 448 | 448 | 448 | 448 |
| Hansen test (p-value) | 0.54 | 0.48 | 0.50 | 0.58 | 0.51 | 0.76 | 0.85 | 0.75 | 0.73 | 0.54 |
| ar2p | 0.45 | 0.09 | 0.14 | 0.17 | 0.09 | 0.85 | 0.24 | 0.04 | 0.09 | 0.40 |
| ar1p | 0.03 | 0.04 | 0.04 | 0.04 | 0.05 | 0.00 | 0.00 | 0.00 | 0.00 | 0.01 |

Note: Dependent variable is the change in debt in percent of GDP. Robust standard errors clustered at the country level in parenthesis. The Hansen test evaluates the validity of the instrument set, i.e., tests for over-identifying restrictions. AR(1) and AR(2) are the Arellano-Bond autocorrelation tests of first and second order (the null is no autocorrelation), respectively. A constant term has been estimated but it is not reported for reasons of parsimony. *, **, *** denote statistical significance at the 10, 5 and 1 percent levels, respectively. “FR_1” if a country has at least one fiscal rule; “ER” = expenditure rule in place; “RR” revenue rule in place; “DR” = debt rule in place; “BBR” = budget balance rule in place.

Fiscal rules dataset from [Lledó et al. \(2017\)](#).

5. Conclusion and policy implications

We have assessed how SFA contribute to the path of the debt-to-GDP ratio in a panel of 65 countries in the period 1985–2014. Therefore, and vis-à-vis previous related literature, we extend the coverage beyond developed economies to include also emerging and low-income countries, an aspect previously unexplored. In addition to inspecting the role of fiscal rules in affecting SFAs, we go deeper in the analysis by looking at different types and design characteristics of those rules. Furthermore, for a subsample of countries, we also rely on a recent dataset on fiscal institutions and examine which matter the most for the build-up of SFAs.

Our main results are the following: (i) SFAs contribute to the change in the level of the debt-to-GDP ratio with a coefficient close to one. (ii) Fiscal rules in general did not led governments to a systematic use SFAs to lower deficits in our country sample. (iii) Countries with some form of macro-fiscal forecasting tool in place have allowed governments to use SFAs to lower deficit. (iv) The existence of fiscal rules in which monitor compliance exist, contributes to lower the debt level, although the cyclical deficit partly counteracts this desirable effect. (v) The magnitude of the fall in the debt ratio due to the presence of fiscal rules before the crisis was between 1.7 and 4.2 percent of GDP. (vi) After the GFC, both revenue-based and debt-based rules starting contributing positively to the accumulation of debt, a fact that is reinforced with the presence of large SFAs.

Our results have a number of policy implications. First, it is important to notice the effect of the GFC in reversing, to some extent, the performance of the fiscal rules in curbing government debt developments. Hence, policy makers would need to integrate this result in the implementation and redefinition of such fiscal frameworks. Second, the fact that in some cases countries used the SFA with an impact on the development of general government debt, raises the issue of being cautious when perceiving the dynamics of the intertemporal government budget constraint essentially via the primary balance, implying the need to further transparency in that context, to ensure the mitigation of such SFA.

Finally, and as a possible future research line, one could consider analysing the decomposition of SFAs to assess further the role of fiscal rules, although such data are less likely available for non-European economies.

Table 7
Robustness to other estimations, equations (4) and (5-6).

| Specification | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|-----------------|---------------------|--------------------|--------------------|---------------------|--------------------|--------------------|--------------------|---------------------|--------------------|
| Estimator | OLS | OLS | OLS | FE | FE | FE | LSDV | LSDV | LSDV |
| Equation | Eq. (4) | Eq. (5) | Eq. (6) | Eq. (4) | Eq. (5) | Eq. (6) | Eq. (4) | Eq. (5) | Eq. (6) |
| lagged debt | 0.16*** (0.040) | 0.10 (0.091) | 0.19* (0.108) | 0.12*** (0.029) | 0.02 (0.088) | 0.07 (0.092) | 0.01 (0.026) | -0.13*** (0.040) | -0.06 (0.054) |
| SFA | 0.88*** (0.057) | | | 0.98*** (0.037) | | | 0.99*** (0.017) | | |
| FR_1 | -1.04*** (0.282) | 0.83 (0.700) | 0.35 (0.476) | -1.14*** (0.230) | 1.58* (0.927) | 0.86 (0.936) | -0.85 (0.537) | 1.79*** (0.513) | 0.89 (0.784) |
| SFA_FR | -0.08 (0.076) | | | -0.00 (0.042) | | | 0.00 (0.024) | | |
| deficit | | 0.45*** (0.127) | | | 0.87*** (0.156) | | | 0.94*** (0.099) | |
| def_FR | | 0.16 (0.124) | | | 0.07 (0.123) | | | 0.05 (0.099) | |
| cadef | | | 0.35*** (0.098) | | | 0.69*** (0.121) | | | 0.75*** (0.149) |
| cyclicaldef_weo | | | 0.26* (0.160) | | | 1.05*** (0.267) | | | 1.19*** (0.254) |
| cadef_FR | | | 0.23** (0.101) | | | 0.35*** (0.122) | | | 0.35** (0.157) |
| cycdef_FR | | | 0.15 (0.187) | | | -0.21 (0.228) | | | -0.22 (0.220) |
| Observations | 1336 | 1336 | 933 | 1336 | 1336 | 933 | 1336 | 1336 | 933 |
| R-squared | 0.68 | 0.16 | 0.21 | 0.82 | 0.25 | 0.32 | 0.85 | 0.22 | 0.26 |

Note: Dependent variable is the change in debt in percent of GDP. Robust standard errors clustered at the country level in parenthesis. The Hansen test evaluates the validity of the instrument set, i.e., tests for over-identifying restrictions. AR(1) and AR(2) are the Arellano-Bond autocorrelation tests of first and second order (the null is no autocorrelation), respectively. A constant term has been estimated but it is not reported for reasons of parsimony. Specifications 4–6 include country and time fixed effects omitted for reasons of patrimony. *, **, *** denote statistical significance at the 10, 5 and 1 percent levels, respectively. “FR_1” if a country has at least one fiscal rule; “ER” = expenditure rule in place; “RR” revenue rule in place; “DR” = debt rule in place; “BBR” = budget balance rule in place.

Fiscal rules dataset from [Ledó et al. \(2017\)](#).

APPENDIX

Table A1
Summary Statistics

| Variable | Observations | Mean | Standard deviation | Minimum | Maximum |
|---------------------|--------------|-------|--------------------|---------|---------|
| Debt | 3708 | 57.17 | 48.38 | 0 | 789.83 |
| SFA | 3708 | -3.53 | 21.25 | -590.09 | 180.41 |
| Deficit | 3708 | 2.80 | 16.71 | -122.18 | 557.49 |
| cadef | 1377 | 2.42 | 3.78 | -11.92 | 18.70 |
| cyclicaldef | 1377 | -0.12 | 2.76 | -21.60 | 24.51 |
| FR | 1370 | 0.68 | 0.46 | 0 | 1 |
| ER | 1370 | 0.24 | 0.42 | 0 | 1 |
| RR | 1370 | 0.12 | 0.32 | 0 | 1 |
| BBR | 1370 | 0.62 | 0.48 | 0 | 1 |
| DR | 1370 | 0.51 | 0.50 | 0 | 1 |
| FR_monitor | 1170 | 0.17 | 0.38 | 0 | 1 |
| FR_enforce | 1413 | 0.16 | 0.36 | 0 | 1 |
| FR_escape | 888 | 0.07 | 0.26 | 0 | 1 |
| Independent monitor | 1388 | 0.6 | 0.2 | 0 | 1 |
| Transparency | 1413 | 0.14 | 0.35 | 0 | 1 |

Table A2
Robustness with explicit macroeconomic regressors included

| Specification | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
|----------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|
| Regressors | | | | | | | | | | |
| lagged debt | 0.32*** (0.059) | 0.30*** (0.063) | 0.32*** (0.067) | 0.32*** (0.057) | 0.29*** (0.065) | 0.29*** (0.076) | 0.31*** (0.069) | 0.30*** (0.100) | 0.32*** (0.068) | 0.31*** (0.069) |
| Inflation rate | 0.07 (0.081) | -0.03 (0.097) | -0.01 (0.117) | 0.09 (0.111) | 0.09 (0.122) | -0.05 (0.094) | 0.01 (0.100) | 0.05 (0.148) | -0.03 (0.105) | -0.02 (0.107) |

(continued on next page)

Table A2 (continued)

| Specification | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
|-----------------------|---------|---------|---------|---------|---------|---------|---------|---------|---------|---------|
| Output gap | −0.51* | −0.47** | −0.51* | −0.46 | −0.50** | −0.61* | −0.47** | −0.32 | −0.37 | −0.49* |
| | (0.307) | (0.207) | (0.290) | (0.287) | (0.246) | (0.330) | (0.222) | (0.313) | (0.262) | (0.280) |
| SFA | 0.67*** | 0.65*** | 0.74*** | 0.68*** | 0.72*** | 0.60*** | 0.71*** | 0.89*** | 0.64*** | 0.64*** |
| | (0.089) | (0.116) | (0.117) | (0.111) | (0.117) | (0.094) | (0.135) | (0.084) | (0.107) | (0.110) |
| FR_1 | 1.23 | | | | | | | | | |
| | (0.762) | | | | | | | | | |
| SFA x_FR | 0.13 | | | | | | | | | |
| | (0.130) | | | | | | | | | |
| ER | | 0.06 | | | | | | | | |
| | | (1.011) | | | | | | | | |
| SFA x_ER | | 0.25 | | | | | | | | |
| | | (0.163) | | | | | | | | |
| RR | | | −0.07 | | | | | | | |
| | | | (1.144) | | | | | | | |
| SFA x RR | | | 0.05 | | | | | | | |
| | | | (0.458) | | | | | | | |
| BBR | | | | 1.40 | | | | | | |
| | | | | (0.891) | | | | | | |
| SFA xBBR | | | | 0.16 | | | | | | |
| | | | | (0.113) | | | | | | |
| DR | | | | | 2.91 | | | | | |
| | | | | | (2.453) | | | | | |
| SFA x DR | | | | | 0.19 | | | | | |
| | | | | | (0.216) | | | | | |
| FR_monitor | | | | | | 1.78 | | | | |
| | | | | | | (2.010) | | | | |
| SFA x_FRmonitor | | | | | | 0.33*** | | | | |
| | | | | | | (0.088) | | | | |
| FR_enforce | | | | | | | 0.54 | | | |
| | | | | | | | (1.104) | | | |
| SFA x FRenforce | | | | | | | 0.19 | | | |
| | | | | | | | (0.155) | | | |
| FR_escape | | | | | | | | 0.06 | | |
| | | | | | | | | (1.842) | | |
| SFA x FRescape | | | | | | | | −0.25 | | |
| | | | | | | | | (0.440) | | |
| Independent_monitor | | | | | | | | | 1.37 | |
| | | | | | | | | | (0.860) | |
| SFA x indepmonit | | | | | | | | | 0.31*** | |
| | | | | | | | | | (0.104) | |
| Transparency | | | | | | | | | | 0.55 |
| | | | | | | | | | | (0.996) |
| SFA x transp | | | | | | | | | | 0.30*** |
| | | | | | | | | | | (0.108) |
| Observations | 1086 | 1086 | 1086 | 1086 | 1086 | 901 | 1086 | 713 | 1061 | 1086 |
| Hansen test (p-value) | 0.69 | 0.57 | 0.57 | 0.70 | 0.68 | 0.71 | 0.59 | 0.62 | 0.73 | 0.63 |
| ar2p | 0.12 | 0.17 | 0.13 | 0.19 | 0.17 | 0.23 | 0.29 | 0.12 | 0.22 | 0.15 |
| ar1p | 0.01 | 0.00 | 0.02 | 0.01 | 0.02 | 0.02 | 0.01 | 0.03 | 0.02 | 0.00 |

Note: Dependent variable is the change in debt in percent of GDP. Robust standard errors clustered at the country level in parenthesis. The Hansen test evaluates the validity of the instrument set, i.e., tests for over-identifying restrictions. AR(1) and AR(2) are the Arellano-Bond autocorrelation tests of first and second order (the null is no autocorrelation), respectively. A constant term has been estimated but it is not reported for reasons of parsimony. Specifications 4–6 include country and time fixed effects omitted for reasons of patrimony. *, **, *** denote statistical significance at the 10, 5 and 1 percent levels, respectively. “FR_1” if a country has at least one fiscal rule; “ER” = expenditure rule in place; “RR” revenue rule in place; “DR” = debt rule in place; “BBR” = budget balance rule in place.

Fiscal rules dataset from Lledó et al. (2017).

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