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ECOLOGICAL CAUSES OF LIFE HISTORY VARIATION TESTED BY META-ANALYSIS,  
COMPARISON, AND EXPERIMENTAL APPROACHES

By

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B.S., The Evergreen State College, Olympia, Washington, 2001

Dissertation

presented in partial fulfillment of the requirements  
for the degree of  
Doctor of Philosophy  
in Biological Sciences, Organismal Biology and Ecology

The University of Montana  
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## Ecological Causes of Life History Variation

Chairperson: Dr. Thomas E. Martin

The ecological causes of life history variation among taxa and the arrangement of such variation along geographic gradients is enigmatic despite the proximity of life history traits to fitness and implications for understanding basic and applied population ecology. One classic explanation for the arrangement of avian life histories along a ‘slow-fast’ gradient, where species at low latitudes have ‘slow’ life history traits (low fecundity and mortality) and species at high latitudes have ‘fast’ life history traits (high fecundity and mortality), is the increase in seasonality of resources with increasing latitude (Ashmole’s hypothesis). Despite broad acceptance, this hypothesis has been supported only indirectly. I tested two key predictions of this hypothesis – that most mortality occurs in winter and that most mortality is caused by starvation – using meta-analysis. Surprisingly, in many populations, the season of greatest mortality was summer, and most mortality was caused by predation. These results suggest alternative explanations for life history variation should remain under consideration despite support for Ashmole’s hypothesis.

The relationship between provisioning behavior and offspring number was long recognized to integrate key life history tradeoffs between number and quality of offspring and between current and future reproductive success. Studies of the response of parental provisioning behavior to brood size variation played a formative role in the development of life history theory. Yet, the inference of such experiments for explaining among-species differences has always been limited by lack of comparative context. I expanded predictions of alternative ecological explanations (food limitation, nest predation, adult mortality) for life history variation to an among-species context and test these predictions using a comparative-experimental design across a broad range of bird species from three continents. I found resource limitation and adult mortality risk interact to explain variation among species in responses to natural and experimental variation in brood size, with the degree of food limitation appearing to vary across a gradient of adult mortality risk. This result helps to explain the potentially conflicting results of previous studies and suggests a pluralistic approach to understanding what factors explain life history variation may be fruitful.

Understanding variation among species in mortality rates may thus be pivotal to understanding ecological causes of life history variation. To this end, I compared differences in spatiotemporal variance in survival among three temperate-breeding species with differing migratory strategy. I found that migratory behavior may be associated with reduced spatial variance in annual survival because resident species disperse less, reducing population connectivity. I also found that migratory behavior is associated with increased temporal variance in survival, counter to expectations of general theory. Given the potential importance of mortality risk in life history evolution, expanded geographic comparisons of annual and within-year patterns of variance in survival rates is likely key to understanding variation among species in life history traits.

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## Table of Contents

Abstract	ii
Acknowledgments	iii
Table of Contents	iv
List of Figures and Tables	v
Chapter 1: Introduction	1
Chapter 2: Timing and proximate causes of mortality in wild bird populations: testing Ashmole's hypothesis	8
Chapter 3: Causes of world-wide variation in parental provisioning behavior relative to offspring number	44
Chapter 4: Dispersal and spatiotemporal variance in survival in three passerine bird species with differing migratory behavior	71
Appendix 1: Species, season of greatest mortality, and data sources used in meta-analyses of seasonal mortality rates	104
Appendix 2: Data sources used in meta-analysis of proximate mortality causes	117
Appendix 3: Phylogenetic analyses of the relationships between ecological factors and reaction norms of feeding rate to offspring number	130

## List of Tables and Figures

### Chapter 2

Table 1. Studies reporting support for Ashmole's hypothesis	30
Table 2. Season of greatest mortality and study methodology across 69 year-round studies of mortality in wild bird populations	34
Table 3. Relationship between monthly actual evapotranspiration (AE) and monthly survival rates in 39 wild bird populations	35
Figure 1. Graphical representation of Ashmole's hypothesis	39
Figure 2. Forest plot of the proportion of known mortalities due to predation in populations of adults in winter	40
Figure 3. Forest plot of the proportion of known mortalities due to predation in populations of mixed age classes studied year-round	41
Figure 4. Forest plot of the proportion of known mortalities due to predation in populations that were not categories as adults in winter, fledglings, or year-round studies	42
Figure 5. Forest plot of the proportion of known mortalities in populations studies daily or less frequently	43

### Chapter 3

Figure 1. Among-species variation in reaction norms of per-offspring provisioning to natural variation in offspring number	66
Figure 2. Phylogeny, standard name, geographic site, and life history traits of 29 bird species studied	67
Figure 3. Among-species variation of per-offspring provisioning rate to natural variation in offspring number and relationship to ecological factors	68
Figure 4. Among-species variation in responses of total provisioning rate to experimental brood size reductions and relationship to adult mortality risk	69
Figure 5. Among-species variation in responses of per-nestling provisioning rate and offspring mass to experimental brood size reductions	70

## Chapter 4

Table 1. Model selection criteria for alternative multi-state mark-recapture models of the effects of species, sex, strata, and distance between strata on survival, recapture probability, and movement probability	96
Table 2. Posterior probability distribution for annual survival and spatial process variation in annual survival for three species of birds	98
Figure 1. Model-averaged maximum likelihood estimates and confidence intervals of the relationship between breeding dispersal probability and distance between strata for three species of passerine bird	101
Figure 2. Model-averaged maximum likelihood estimates and confidence intervals of the relationship between breeding dispersal probability and distance between strata by sex for two species of passerine bird	102
Figure 3. Comparison of estimates of the temporal process coefficient of variation in annual survival rates of 22 populations of migrant and resident passerine birds	103

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**Chapter 1: Introduction**

The expression of life history traits, or age-specific components of fecundity and mortality, determines the fitness of individuals and the growth rates of populations (Roff 2002). This deterministic relationship between life history traits, fitness, and population growth motivates a central role for life history theory in evolutionary ecology (Stearns 1992). A widely recognized yet still poorly-explained pattern in the expression of life history traits is variation among populations or species in fecundity and mortality (Stearns 1992, Ricklefs 2000, Roff 2002, Martin 2004). This pattern is perhaps best described as a ‘slow-fast’ gradient, along which ‘slow’ populations express low fecundity and low mortality, while ‘fast’ populations express high fecundity and mortality (Bennett and Owens 2002). Despite the near-universality of this pattern in vertebrates (Dunham and Miles 1985, Gaillard et al. 1989, Clobert et al. 1998, Rochet et al. 2000) and its consequences for evolution (e.g. Stearns 1992), population ecology (e.g. Sæther and Bakke 2000), and conservation biology (e.g. Heppell 1998), our understanding of the causes of this pattern remains poor at best.

Life history theory largely explains this broad pattern of variation among populations through tradeoffs between traits (Stearns 1989), ecological differences among populations that cause natural selection on life history traits (Roff 2002), and phylogenetic effects of shared evolutionary history on differences among species in expression of traits (Owens and Bennett 1995). Tradeoffs are invoked to explain why variation is constrained to occur along a slow-fast gradient despite persistent natural selection for increased fecundity and decreased mortality (Stearns 1989). Yet, tradeoffs do not alone explain why some species are ‘fast’ and some are ‘slow’ (Roff 2002). Ecological differences among populations in resource availability (Lack



25 1947) or extrinsic mortality factors (Moreau 1944, Skutch 1949, Stearns 1992) are widely  
26 invoked to explain the positioning of species along this slow-fast gradient. This line of reasoning  
27 is supported by geographic patterns in variation among species in life history traits (Jetz et al.  
28 2008). Life history traits frequently covary strongly with latitude, suggesting that some  
29 ecological factor that also covaries with latitude is responsible (Hussell 1985). In terrestrial  
30 birds, species that occur at high latitudes display ‘fast’ life history traits (high fecundity and  
31 mortality) while species that occur at low latitudes display ‘slow’ life history traits (low  
32 fecundity and mortality), independent of shared phylogenetic history (Martin et al. 2000, Martin  
33 and Ghalambor 2001, Jetz et al. 2008). The recognition of the strength and enigmatic nature of  
34 this pattern in terrestrial birds (Moreau 1944, Lack 1947, Skutch 1949) drove the early  
35 development of general life history theory (Stearns 1992, Ricklefs 2000, Martin 2004) and  
36 remains an active area of research 60 years later.

37         Despite enduring research interest in explaining the ecological factors responsible for  
38 latitudinal gradients in the life history variation of birds, numerous questions remain. Several  
39 leading explanations for geographic patterns in life history variation invoke unvalidated  
40 assumptions and are supported only by relatively indirect evidence (Stearns 1992, Roff 2002). In  
41 Chapter 2, I test two generally untested and unrecognized mechanistic predictions of Ashmole’s  
42 hypothesis (Ashmole 1961, 1963), a leading explanation for geographic variation in bird life  
43 histories. This hypothesis is predicated on density-dependent mortality caused by geographic  
44 variation in the seasonal dynamics of resource availability (Ricklefs 1980, Jetz et al. 2008,  
45 Ricklefs 2010). My results are contrary to two simple predictions of this hypothesis: first that  
46 most mortality should occur in seasons of low resources (i.e. winter) and second, that most  
47 mortality should be caused by starvation. Instead, a diversity of terrestrial bird species appear to

48 commonly display greatest seasonal mortality rates in the summer, and the leading proximate  
49 cause of mortality is predation (Chapter 2). This raises key questions about the validity of this  
50 hypothesis despite general acceptance (Jetz et al. 2008, Ricklefs 2010) based on more indirect  
51 forms of support (reviewed in Chapter 2) and suggests consideration of alternative explanations,  
52 or proposal of new explanations, for geographic patterns in life history variation.

53         One classic context in which alternative ecological explanations for variation in the  
54 expression of life history traits have been tested is studies of the response of parental behaviors  
55 to natural or experimentally-induced variation in brood size (Nur 1984, Linden and Moller 1989,  
56 VanderWerf 1992). While many of these experiments were designed to test explanations for  
57 differences among species in life history variation, their design has been limited to single  
58 species, making generalization to an among-species context uncertain (Martin 2004). I expanded  
59 the predictions of this now-classic study design to an among-species context to allow tests for  
60 general ecological explanations for among-species variation in life history traits (Chapter 3). I  
61 then tested these predictions using a combination of comparison (29 species) and comparative  
62 experiment (9 species) from 4 study sites on 3 continents (Chapter 3). Species differed in the  
63 response of parental care, measured as provisioning rates, to natural and experimental variation  
64 in brood size, and the differences among species were explained by apparent food limitation *and*  
65 differences in adult mortality risk. My results suggest that food limitation and adult mortality  
66 risk may interact to explain the observed geographic variation in bird life history traits, which  
67 helps resolve potentially conflicting results of previous studies (e.g. Golet et al. 1998) and  
68 emphasizes that pluralist approaches may be important to ultimately understanding what  
69 ecological factors play a role in life history evolution.

70           Given the potential importance of adult mortality risk in explaining geographic patterns  
71 in life history variation (Chapter 3) and the role that spatiotemporal variance in mortality risk  
72 may play in population growth (Tuljapurkar 1982) and life history evolution (Roff 2002), I  
73 examined how differences among species in dispersal behavior and migratory behavior may  
74 influence spatiotemporal variance in survival rates (Chapter 4). I found that increasing dispersal  
75 among three species with differing migratory behavior (from residency to long-distance  
76 migration) was associated with increasing migratory distance and with reduced spatial variance  
77 in survival rates (Chapter 4). Yet, temporal variance in survival was not related to migratory  
78 distance as previously predicted (Greenberg 1980). Here, I have begun to address an important  
79 question raised by this dissertation that will likely be a productive target of future research:  
80 considering the potential importance of mortality rates in life history evolution, what explains  
81 geographic variation within and among species in mortality risk?

82           My results suggest mortality risk is important in explaining geographic variation in life  
83 histories, yet we do not know why mortality risk varies geographically. Is this variation driven  
84 by resource dynamics, predation risk, an interaction between the two, or some other ecological  
85 factor that may covary with latitude? Do differences among species in other traits, such as  
86 migratory behavior, partly explain differences in mortality risk? Do species with alternative life  
87 history strategies resolve life history tradeoffs, other than the tradeoff between number and  
88 quality of offspring I have compared among species here, differently? While I have addressed  
89 the timing and proximate causes of mortality in published studies of wild bird populations  
90 (Chapter 2) and compared how three species differ in spatiotemporal variance in survival  
91 (Chapter 4), there is a clear paucity of such information for tropical bird species, and thus

92 understanding whether and how such patterns in mortality risk vary geographically remains an  
93 open question.

94

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153

154 **Chapter 2: Timing and proximate causes of mortality in wild bird populations: testing**

155 **Ashmole's hypothesis**

156

157 Daniel C. Barton and Thomas E. Martin

158

159 **Abstract**

- 160 1. Fecundity in birds is widely recognized to increase with latitude across diverse  
161 phylogenetic groups and regions, yet the causes of this variation remain enigmatic.
- 162 2. Ashmole's hypothesis is one of the most broadly accepted explanations for this pattern.  
163 This hypothesis suggests that increasing seasonality leads to increasing overwinter  
164 mortality due to resource scarcity during the lean season (e.g., winter) in higher latitude  
165 climates. This mortality is then thought to yield increased per-capita resources for  
166 breeding that allow larger clutch sizes at high latitudes. Support for this hypothesis has  
167 been based on indirect tests, whereas the underlying mechanisms and assumptions remain  
168 poorly explored.
- 169 3. We used a meta-analysis of over 150 published studies to test two underlying and critical  
170 assumptions of Ashmole's hypothesis: first, that adult mortality is greatest during the  
171 season of greatest resource scarcity, and second, that most mortality is caused by  
172 starvation.
- 173 4. We found that the lean season (winter) was generally not the season of greatest mortality.  
174 Instead, spring or summer was most frequently the season of greatest mortality.  
175 Moreover, monthly survival rates were not explained by monthly productivity, again

176 opposing predictions from Ashmole's hypothesis. Finally, predation, rather than  
177 starvation, was the most frequent proximate cause of mortality.

178 5. Our results do not support the mechanistic predictions of Ashmole's hypothesis, and  
179 suggest alternative explanations of latitudinal variation in clutch size should remain under  
180 consideration. Our meta-analysis also highlights a paucity of data available on the timing  
181 and causes of mortality in many bird populations, particularly tropical bird populations,  
182 despite the clear theoretical and empirical importance of such data.

183

184 **Key-words** seasonal mortality, cause-specific mortality, life history, evolution, latitudinal  
185 gradient

186

## 187 **Introduction**

188 Explaining life history variation among species is a principle goal of evolutionary biology due to  
189 the importance of life history traits to fitness (Roff 2002) and population dynamics (Sæther &  
190 Bakke 2000). A long-recognized yet enigmatic pattern in life history variation is the increase in  
191 fecundity (clutch size) among terrestrial species birds with increasing latitude (Moreau 1944;  
192 Lack 1947; Skutch 1949). Replication of this pattern across diverse phylogenetic groups of birds  
193 on multiple continents suggests that an environmental factor that covaries with latitude is  
194 responsible (Martin 1996; Jetz, Sekercioglu & Böhning-Gaese 2008). Alternative environmental  
195 factors proposed to explain this pattern include day length (Hussell 1985), food availability  
196 (Lack 1947), nest predation risk (Skutch 1949), adult mortality risk (Law 1979; Martin 2004),  
197 and seasonality of resources (Ashmole 1961, 1963).



198 Ashmole's hypothesis is one of the most widely accepted explanations for latitudinal  
199 variation in clutch sizes of birds (Table 1). Ashmole's hypothesis explains clutch size variation  
200 as a function of varying seasonality in resource availability among latitudes. Under this  
201 hypothesis, population sizes are thought to be limited by mortality caused by low levels of  
202 resources during the winter season, and reproduction (clutch size) is limited by per-capita  
203 resources available during the breeding season (Ashmole 1961, 1963; Ricklefs 1980; Martin  
204 1996; McNamara et al. 2008; Figure 1). Through this mechanism, mortality from scarce  
205 resources during the lean season (i.e. winter) is argued to regulate population size at a level  
206 substantially below the summer carrying capacity at high latitudes. The resulting increase in per-  
207 capita resources for the breeding population thus explains the latitudinal increase in clutch sizes  
208 of terrestrial birds (Figure 1).

209 Ashmole's hypothesis has been supported via a diversity of tests (Table 1). In particular,  
210 tests of the predicted correlation between degree of seasonality and fecundity (e.g. Ricklefs  
211 1980; Jetz, Sekercioglu & Böhning-Gaese 2008), high seasonal variability in population sizes  
212 (e.g. Ashmole 1961) and simulation-based modeling approaches (e.g. Griebeler & Böhning-  
213 Gaese 2004; McNamara et al. 2008) are suggested to support this hypothesis of clutch size  
214 evolution (Table 1). Yet, these tests have three important shortcomings. First, tests of the  
215 predicted correlation between variation in seasonality and clutch size are indirect in that they do  
216 not test the actual mechanism of the hypothesis. Seasonality of resources as well as many other  
217 environmental factors all co-vary with latitude and thus a relationship between seasonality of  
218 resources and clutch size cannot be considered strong evidence for causation (Dunn & MacInnes  
219 1987; Hussell 1985; Koenig 1986). Second, population sizes can vary extensively even when  
220 mortality is constant year-round if reproduction is seasonal (Fretwell 1972). Thus, population

221 variation is not an appropriate means of testing this hypothesis. Third, simulation-based  
222 modeling approaches have accepted at least one of the assumptions of Ashmole's hypothesis (i.e.  
223 overwinter food limitation of population size) to be true *a priori* (e.g. Griebeler & Böhning-  
224 Gaese 2004; McNamara et al. 2008). While patterns predicted to result from Ashmole's  
225 hypothesis are supported, other hypotheses also can explain these patterns (Martin 1996, 2004).  
226 Ultimately, direct tests of the assumptions and mechanism of Ashmole's hypothesis are needed.

227         Ashmole's hypothesis makes two key assumptions regarding patterns of mortality in bird  
228 populations that have not been generally tested. First, Ashmole's hypothesis argues that  
229 mortality rates are highest during the season of lowest resource abundance (Figure 1; Ashmole  
230 1963; Ricklefs 1980; McNamara et al. 2008). Second, Ashmole's hypothesis assumes that the  
231 primary cause of mortality is starvation due to density-dependent resource scarcity during the  
232 lean season. Despite an emerging emphasis on Ashmole's hypothesis for explaining a major  
233 worldwide pattern in life history variation (Griebeler & Böhning-Gaese 2004; Jetz, Sekercioglu  
234 & Böhning-Gaese 2008; McNamara et al. 2008; Ricklefs 2010), the underlying assumptions and  
235 mechanistic basis of this hypothesis remain largely untested. Given a diversity of alternative  
236 hypotheses, as well as the increasing acceptance of Ashmole's hypothesis (Table 1), clear tests of  
237 these assumptions are critical to advancing our understanding of the ecological basis of  
238 geographic variation in life history strategies.

239         We used literature data compiled from a comprehensive review and meta-analysis to test  
240 these two assumptions of Ashmole's hypothesis across a diversity of species for the first time.  
241 First, we tested whether mortality rates of bird populations were lowest in winter by reviewing  
242 bird-banding and radio-telemetry studies that estimated seasonal mortality rates of bird  
243 populations year-round. We further tested whether seasonal variation in resource productivity

244 predicted seasonal variation in mortality rates. Second, we tested whether starvation was the  
245 primary cause of mortality in wild populations by performing a quantitative meta-analysis of  
246 studies which assigned known mortality causes to birds equipped with radio telemetry devices.  
247 Our results provide a broad characterization of the timing and sources of mortality in wild bird  
248 populations across diverse taxonomic groups.

249

## 250 **Methods**

### 251 **Seasonal mortality rates of bird populations**

252 We reviewed the literature for studies of seasonal mortality rates of bird populations from year-  
253 round mark-resight-recapture or radio-telemetry studies. We searched the literature for ‘seasonal  
254 mortality birds’ and ‘seasonal survival birds’ using the Google Scholar and ISI-Thompson Web  
255 of Knowledge databases, and manually searched years 1990-2010 of *Auk*, *Condor*, *Wilson*  
256 *Journal of Ornithology*, *Journal of Wildlife Management*, *Journal of Avian Biology*, *Journal of*  
257 *Field Ornithology*, *Ibis*, *Wildlife Society Bulletin*, and *Ardea*. We chose 1990 as the starting  
258 point of our manual searches because this year is near the beginning of an explosion in the  
259 number of published studies of survival using mark-recapture and radio-telemetry techniques in  
260 wildlife biology and ecology. We further supplemented our initial set of studies using forward-  
261 and backward-citations. We only included peer-reviewed studies and Master’s theses or PhD  
262 dissertations in our review.

263 Three criteria qualified a study for admission to our review and analysis. First, studies  
264 had to report estimates of survival from at least two seasons representing an entire calendar year,  
265 and these seasons needed to represent at a minimum ‘spring/summer’ and ‘fall/winter’. This  
266 criterion gave us our primary measure of the seasonality of survival rates – and allowed us to test

267 whether mortality was greatest during the season of predicted low resource abundance (i.e. fall  
268 and/or winter) or not. Some studies reported survival from as few as two seasons while others  
269 reported survival rates for bi-weekly intervals for an entire year. Second, studies admitted were  
270 of natural variation in survival reported from populations in more or less natural settings. We  
271 included estimates from hunted populations when estimates of survival in the absence of hunting  
272 were also reported or the relative ranks of seasonal survival were apparently unchanged by  
273 hunting (i.e. when hunting mortality was smaller than among-season differences in mortality).  
274 Third, we admitted studies of seasonal survival that used estimates of seasonal survival  
275 developed from a technique that could reasonably assume resighting probability was at or very  
276 near 1 (such as radio-telemetry; Pollock, Winterstein & Conroy 1989), intensive searches in  
277 closed or nearly-closed populations (e.g. Arcese et al. 1992), or a mark-recapture analysis that  
278 accounted for resighting probabilities of less than 1 (Lebreton et al. 1992). This eliminated  
279 numerous band-recovery estimates of seasonal survival (e.g. Dobson 1987) because they are  
280 likely biased in several key ways, including seasonal bias in reporting probabilities, which  
281 compromise explicitly seasonal comparisons.

282         We combined the results of these studies into a meta-analysis to test the first assumption  
283 of Ashmole's hypothesis: that most mortality should occur in the winter. We scored each study  
284 for whether most mortality occurred in fall/winter (defined as the season of lowest resource  
285 abundance in the few tropical studies included in the analysis), a season other than winter, or  
286 whether support was equivocal (i.e. if mortality was similar year-round, or in the subset of  
287 studies of migrant populations, if migration confounded seasonal variation). To provide another  
288 more highly conservative test of Ashmole's hypothesis, we took all studies scored as equivocal  
289 and re-scored them as mortality highest in winter, because statistical power to detect differences

290 in survival rates among seasons may often be low. Even though it is unlikely all studies with  
291 equivocal seasonality reflect highest mortality in winter, especially ‘equivocal’ studies reporting  
292 that migration may be the season of greatest mortality, taking this conservative approach allows  
293 examination of whether any support for Ashmole’s hypothesis exists in this dataset. In the  
294 subset of studies of resident birds, we used logistic regression to test whether the probability that  
295 a study would report most mortality occurred in winter increased with latitude.

296 To test the predicted positive relationship between resource productivity and survival  
297 rates, we extracted monthly survival rates from studies included in this review when possible.  
298 We used survival rates in this analysis because transformation to normality to achieve the  
299 assumptions of regression was achievable with survival, but not mortality, rates. Monthly  
300 survival rates were extracted from studies by standardizing reported periodic tabular or graphical  
301 rates (using program DigitizeIt; I. Bormann, Germany 2006) to monthly rates. We tested  
302 whether monthly productivity, as measured by actual evapotranspiration (Mather 1962, 1963a,  
303 1963b, 1963c, 1964a, 1964b, 1964c, 1965; Ricklefs 1980), predicted monthly variation in  
304 survival rates using simple linear regression of arc-sine transformed monthly survival rates on  
305 log-transformed monthly actual evapotranspiration. We excluded year-round studies of  
306 migratory populations because it was often unclear which environment these populations  
307 occupied, and thus which values of monthly actual evapotranspiration the population  
308 experienced were ambiguous.

### 309 **Sources of mortality in wild bird populations**

310 We also reviewed the literature for studies that used radio-telemetry to assign mortality causes to  
311 wild bird populations. We employed the same search strategy described above except we used  
312 the following search terms: ‘known fate birds’, ‘mortality cause birds’, and ‘mortality source

313 birds'. Some studies located during our initial search for seasonal survival rates of bird  
314 populations included known fate data and thus were included in this meta-analysis as well.

315         Studies were admitted to this review and analysis based on two criteria. First, we only  
316 admitted studies that assigned fates (cause-specific mortality) to telemetered birds based on  
317 standardized criteria and that reported the exact numbers of birds assigned to each specific fate.  
318 These fates were our primary measure of the different causes of mortality in each bird  
319 population. Second, we only admitted studies of wild bird populations in more or less natural  
320 settings; thus, we did not admit any introduction or re-introduction programs to our review. For  
321 each study admitted, we collected and entered into a database: species, age class(es), season(s)  
322 studied, frequency of observation, total sample size of the study, and fates assigned to different  
323 categories. Ambiguous values were treated as missing values. These data were then used in a  
324 quantitative meta-analysis.

325         We conducted a meta-analysis of mortalities assigned to different causes to test the  
326 prediction of Ashmole's hypothesis that the principle cause of mortality should be starvation. A  
327 preliminary analysis (a simple summed proportion across all studies in the database) suggested  
328 predation was the major source of mortality, at least when summed across studies, which led us  
329 to design a conservative test of Ashmole's hypothesis. We summed the proportion of individuals  
330 assigned to two different fate categories (predation, and all other known causes including  
331 starvation, disease, and hunter kill) within each study. We then used a random effects model  
332 (DerSimonian & Laird 1986) of the Freeman-Tukey double arcsine transformed data (Freeman  
333 & Tukey 1950) to generate a pooled across-study 'incidence rate' and 95% confidence intervals  
334 for the proportion of known mortality caused by predation. We repeated this procedure for three  
335 different categories of studies. First, and providing the most direct test of the assumptions of

336 Ashmole's hypothesis, were studies of adults and adult/immature pools in winter. We also  
337 repeated this procedure for studies of adults and adult/immature pools year-round in which the  
338 data were not clearly separable by season and for studies that did not fit clearly into any of the  
339 other two categories (we refer to these as 'uncategorized'). We generated within-study estimates  
340 of proportions and confidence intervals using the exact binomial method.

341 Our meta-analysis thus tested whether a) studies found that most individuals died from  
342 predation or causes other than predation in winter, providing our direct, but conservative, test of  
343 Ashmole's hypothesis, and b) whether studies that combined winter mortality with other  
344 seasonal mortality found that most individuals died from predation or causes other than  
345 predation, which provided a less direct test of Ashmole's hypothesis. The use of simple  
346 proportions of known or total mortality from predation (binomial estimator), rather than  
347 estimators that allow for unequal exposure to risk due to staggered entry of subjects into the  
348 population or right-censoring of data series, could introduce two biases into our results (Heisey  
349 & Fuller 1985). First, the binomial estimator may underestimate mortality caused by mortality  
350 agents if there is staggered entry of subjects into the study population. Second, if both survival  
351 rates and sample size vary seasonally, the season with the largest sample size will have an  
352 inappropriately large effect on the overall estimate of mortality rate (Heisey & Patterson 2006).

353 These biases should not affect our tests of Ashmole's hypothesis for two reasons. First, if  
354 we underestimate mortality caused by predation, we are providing an even more highly  
355 conservative test of the hypothesis. Second, we divided our estimates of cause-specific mortality  
356 used in the meta-analysis into season/age categories (such as winter, the key test of Ashmole's  
357 hypothesis) when possible. We took this generally inclusive approach in our meta-analysis to

358 avoid excluding a very large proportion of studies that lacked sufficient information to calculate  
359 an estimator of cause-specific mortality other than the simple binomial estimator.

360 One methodological concern identified *a priori* was that radio-telemetry studies may be  
361 biased towards mis-assignment of cases of starvation as predation due to postmortem scavenging  
362 by predators (Brand, Vowles & Keith 1975). We tested whether scavenging may have biased the  
363 results of studies by testing whether studies that observed their subjects more often (daily vs. less  
364 often) were less likely to report cases of predation. We matched studies by species, age class,  
365 and season (“study category”) that varied in check frequency according to their methodological  
366 descriptions. We tested whether check frequency affected proportion of mortalities assigned to  
367 predation using a mixed-effect weighted regression model with a fixed effect of check frequency  
368 nested inside a random effect of study category. We used each individual study as a sample unit  
369 with exact binomial proportion of individuals estimated to have died due to predation as the  
370 response variable, weighted by the random effect weight  $W_i$  (DerSimonian & Laird 1986). We  
371 used a t-test to assess the statistical significance of the fixed effect of interest, check frequency.  
372 Computer code used to conduct the analysis is available from DCB.

373

## 374 **Results**

### 375 **Seasonal mortality of bird populations**

376 We obtained seasonal mortality estimates using radio-telemetry for 41 populations of 19 species,  
377 and using mark-recapture for 28 populations of 26 species (Table 2; see Appendix 1 in  
378 Supporting Information). Spring and summer were the most frequently reported season of  
379 greatest mortality (lowest survival) in radio-telemetry studies (Table 2). This result did not



380 change even when ‘equivocal’ studies were re-categorized as mortality being greatest in winter,  
381 providing a highly conservative test (Table 2).

382 Mark-recapture studies reported equivocal patterns of seasonality most frequently, but  
383 among studies that reported differences among seasons, spring and summer again were more  
384 frequently the season of greatest mortality than fall and winter (Table 2). Assigning greatest  
385 mortality in fall and winter to the large number of ‘equivocal’ mark-recapture studies caused fall  
386 and winter to become the season of greatest mortality (Table 2). Nonetheless, the pooled results  
387 across radio-telemetry and mark-recapture studies show that the most frequently reported season  
388 of highest mortality was during the spring or summer by a 2:1 margin, and that many studies  
389 showed no strong seasonal pattern (Table 2).

390 We tested whether latitude predicted the season of greatest mortality within the subset of  
391 48 studies of resident bird populations in which hunting did not confound seasonal mortality  
392 (Appendix 1). The probability that a study reported fall or winter as the season of greatest  
393 mortality was not predicted by the difference between maximum and minimum actual  
394 evapotranspiration (logistic regression:  $\beta = -0.009$ ,  $z = -0.994$ , 46 df,  $P = 0.32$ ). The probability  
395 that a study reported spring or summer as the season of greatest mortality also was not predicted  
396 by latitude (logistic regression:  $\beta = 0.009$ ,  $z = 1.105$ , 46 df,  $P = 0.27$ ). In short, latitude did not  
397 predict season of greatest mortality.

398 We tested the predicted positive relationship between monthly productivity and monthly  
399 survival based on monthly survival rates extracted from 40 resident bird populations. The  
400 relationship between arc-sine transformed monthly survival rates and log-transformed actual  
401 evapotranspiration varied widely among populations (Table 3). In 11 of the 40 populations  
402 examined, survival was significantly ( $P < 0.10$ ) positively related to actual evapotranspiration as

403 predicted by Ashmole's hypothesis. However, survival was significantly negatively related to  
404 actual evapotranspiration in 8 of the 40 populations, and was not significantly related to actual  
405 evapotranspiration in the majority (21 of 40) of studies. Thus, the positive relationship predicted  
406 by Ashmole's hypothesis was not supported among the 19 studies with significant patterns,  
407 which showed approximately equal numbers of positive (11) and negative (8) relationships  
408 between actual evapotranspiration and survival. This predicted positive relationship was also not  
409 supported across all studies regardless of whether the regression was significant, which showed  
410 approximately equal numbers of positive (17) and negative (23) relationships between actual  
411 evapotranspiration and survival. The proportion of populations showing a positive relationship  
412 between monthly actual evapotranspiration and survival did not differ between hunted (12 of 26  
413 positive) and unhunted (5 of 14 positive) populations ( $\chi^2 = 0.09$ , 1 df,  $P = 0.76$ ; Table 3).

#### 414 **Sources of mortality in wild bird populations**

415 Predation was by far the dominant cause of mortality across the studies included in our  
416 meta-analysis. During winter, the random-effects pooled proportion of known mortality caused  
417 by predation was 0.62 (95% CI = 0.52 – 0.72;  $n = 40$  populations of 20 species; Figure 2; see  
418 Appendix 2 in Supporting Information). Predation caused the largest proportion of known  
419 mortality in 23 of these 40 populations (Figure 2; Appendix 2). Among year-round studies of  
420 adult and juvenile mortality (studies that combined winter adult mortality with other seasons or  
421 age classes), the random-effects pooled proportion of known mortality caused by predation was  
422 0.78 (95% CI = 0.73 – 0.83;  $n = 39$  populations of 18 species; Figure 3; Appendix 2). Finally,  
423 for those cases that did not fit in any of the preceding categories, the random-effects pooled  
424 proportion of known mortality was 0.86 (95% CI = 0.76 – 0.94;  $n = 31$  populations of 25 species;

425 Figure 4; Appendix 2). In 27 of these 31 populations predation caused the greatest proportion of  
426 known mortality (Figure 4; Appendix 2).

427 Using a highly conservative approach where unknown mortality was assigned to non-  
428 predation, the random-effects pooled proportion of all mortality was still 0.49 (95% CI = 0.39 –  
429 0.60) during winter; 0.67 (95% CI = 0.63 – 0.71) for adults and juveniles studied year-round; and  
430 0.82 (95% CI = 0.71 – 0.91) for studies not fitting the preceding categories.

431 We tested whether studies that checked status of birds less frequently than daily affected  
432 assignment of mortality to predation using data from 27 studies of 9 species (Figure 5; Appendix  
433 2). The fixed effect of check frequency was not significant in a weighted random-effects mixed  
434 model with a random effect of species ( $t_{\text{check}} = -0.49$ , 16 df,  $P = 0.63$ ). Thus, in this subset of 27  
435 studies, check frequency was not associated with the estimated proportion of mortality due to  
436 predation.

437

## 438 **Discussion**

439 Ashmole's hypothesis is widely accepted as an explanation for latitudinal variation in the clutch  
440 sizes of terrestrial birds (Table 1). Our meta-analysis represents the first broad approach to  
441 testing key assumptions that form the mechanistic basis of this hypothesis (Figure 1). We could  
442 not confirm that winter is the season of greatest mortality (Table 2) or that starvation from scarce  
443 resources is the principle cause of mortality for adults in winter (Figure 2) or in any other season  
444 (Figures 3, 4), despite the vast majority of studies of seasonal mortality in birds having been  
445 conducted at latitudes greater than 30 degrees (Appendix 1). Furthermore, we did not find the  
446 predicted general positive relationship between seasonal survival rates and actual

447 evapotranspiration, and instead found that seasonal survival rates were as often negatively  
448 correlated with actual evapotranspiration as they were positively correlated (Table 3).

449 Analyzing year-round radio telemetry studies and year-round mark-recapture studies, we  
450 found that the season of greatest mortality varied among species but that spring/summer was  
451 often the season of greatest mortality, rather than winter (Appendix 1; Table 2). Similarly,  
452 Bergerud (1988), in a comprehensive review of northern grouse population data, concluded that  
453 density-dependent shortage of winter food and winter mortality did not create a winter  
454 population bottleneck and did not explain breeding population sizes. We recognize that the  
455 sample of species represented in our sample of year-round radio telemetry and mark-recapture  
456 studies is not a fully representative subset of terrestrial bird diversity. Year-round radio  
457 telemetry studies of survival can only be conducted with bird species large enough to carry  
458 transmitters with large and long-lived batteries (Appendix 1). We suggest these methodological  
459 considerations resulted in our review of a large number of studies of populations in the order  
460 Galliformes (quail, grouse, turkeys) that are physically large and economically important  
461 because of hunting (Appendix 1). Likewise, year-round mark-recapture studies are generally  
462 limited to populations which are easily observed or recaptured and have known distributions  
463 year-round. These methodological issues probably resulted in the prior publication and our  
464 review of a large number of studies of Anseriformes (geese, ducks) and resident populations in  
465 the order Passeriformes (songbirds; Appendix 1). Thus, our results on the seasonality of  
466 mortality represent a narrow subset of bird diversity, and we made no attempt to control for the  
467 influence of a phylogeny given this already taxonomically-biased sample. Yet, Ashmole's  
468 hypothesis was originally proposed to generally apply to all bird species across a range of  
469 ecological conditions, including marine birds and terrestrial birds (Ashmole 1961, 1963) and has

470 been broadly applied to explaining clutch size variation in a diversity of terrestrial bird species  
471 (Ricklefs 1980, Jetz, Sekercioglu & Böhning-Gaese 2008). Therefore we consider our results on  
472 the seasonality of mortality as generally falsifying the assumption of greatest mortality in winter  
473 across terrestrial bird species, although some bird populations did show greatest mortality in  
474 winter (Table 2; Appendix 1).

475         A positive relationship between seasonal survival rates and seasonal productivity is  
476 another implicit assumption of Ashmole's hypothesis (Ashmole 1961, Ricklefs 1980). Here we  
477 showed that the relationship between monthly survival rates and actual evapotranspiration, a  
478 measure of productivity, is highly variable among resident bird populations. Some populations  
479 show a positive relationship, some populations show a negative relationship, and most show no  
480 significant relationship. While this may again be the result of limited statistical power to detect  
481 such effects, even the non-significant relationships showed no hint of the positive relationship  
482 expected under Ashmole's hypothesis (Table 3). We take these results to indicate that resource  
483 abundance is not the driving factor in determining seasonal variation in survival (also see  
484 Bergerud 1988) across the diversity of high-latitude bird species represented in our meta-  
485 analysis.

486         This latter point was reinforced by our findings that the season of greatest mortality is not  
487 the lean fall/winter season, but most frequently summer, or alternatively, equivocal support for  
488 either. While inability to identify the season of greatest mortality (i.e. equivocal studies) may  
489 simply result from low statistical power to detect differences among seasons, the positive result  
490 of greatest mortality in the summer in numerous bird populations (Appendix 1; Table 2) is  
491 surprising and is quite contrary to Ashmole's hypothesis. Potential explanations for this pattern  
492 are many, and include increased predation intensity in summer, predation costs of reproduction,

493 and possible carry-over effects from resource scarcity during winter. The only potential  
494 explanation consistent with a role for population size limitation by winter resources in life  
495 history evolution is possible carry-over effects from winter (i.e., Norris et al. 2004). Yet, such  
496 carry-over effects do not represent population limitation as predicted by Ashmole's hypothesis  
497 and would require an expansion and complication of the mechanistic basis of Ashmole's  
498 hypothesis. We suggest examining all three of these potential explanations will be necessary to  
499 understand why mortality is frequently greatest in the summer.

500         We found that starvation was clearly not the leading proximate cause of mortality in the  
501 majority of wild bird populations studied, contrary to the assumptions of Ashmole's hypothesis.  
502 We instead found that predation was the most frequently reported cause of adult mortality in  
503 winter or in studies that combined adult winter mortality data with other seasons (Appendix 2;  
504 Figures 2-4). Again, Bergerud (1988) came to the same conclusions in his analysis of northern  
505 grouse populations. While our review and meta-analysis included only the subset of bird species  
506 that can carry radio-telemetry devices and are amenable to tracking using such technology, it  
507 includes a wide body size range and a wide range of phylogenetic diversity (Appendix 2). Thus  
508 we consider our results generalizable – predation is likely the leading proximate source of  
509 mortality in free-living bird populations, as it is in the sessile eggs and nestlings of altricial birds  
510 (Martin 1995). However, it is critical to consider that we only reviewed studies that assigned  
511 proximate sources of mortality. Resource availability is theoretically proposed (McNamara &  
512 Houston 1990) and empirically demonstrated (Lima & Dill 1990; Cresswell & Whitfield 2008)  
513 to affect the predation risks accepted by birds and thus resource availability may still be the  
514 ultimate cause of much mortality in the bird populations studied. Yet, again, we note that

515 mortality was not greatest during the lean season, so even if food is contributing to predation  
516 mortality, it is not working as predicted under Ashmole's hypothesis.

517         Our results suggest serious problems with the mechanistic basis of Ashmole's hypothesis,  
518 despite its broad support from more indirect methods (Table 1). The vast majority of populations  
519 included in our review and meta-analysis were temperate non-migratory species. While the  
520 majority of mortality in migrant bird populations may occur during migration (Sillett & Holmes  
521 2002), the standing assumption for resident birds is that most mortality occurs during winter due  
522 to resource limitation (Ricklefs 1980). Yet, we found in many populations that most mortality  
523 occurs during spring or summer due to predation. Reconciling these results with current theory  
524 meant to explain life history variation and population regulation will require deeper  
525 consideration of the relationship between risk-taking and breeding, the potential role of extrinsic  
526 mortality in regulating populations of birds and its role in life history evolution, carry-over  
527 effects from conditions during one season to another, and the relationship between ultimate and  
528 proximate sources of mortality. Furthermore, future empirical research on the timing and causes  
529 of mortality in wild bird populations resident at low latitudes will provide greater insights into  
530 mechanisms of population regulation and the potential validity of Ashmole's hypothesis. We  
531 also find that our results raise important questions about how and when bird populations are  
532 regulated – understanding the relative importance of food, predation, and disease in regulating  
533 population sizes and the season(s) in which regulation occurs is of paramount importance in  
534 understanding population biology, but we still know surprisingly little about these processes in  
535 birds – particularly in the tropics.

536

537

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649 Table 1. Studies reporting support for Ashmole’s hypothesis. ‘AE’ stands for actual evapotranspiration. ‘Support for Ashmole’s  
 650 hypothesis’ indicates whether results were interpreted as either positive (+) or negative support (-).

Study	Study Species	Main Result	General Method	Support for Ashmole’s Hypothesis
Ashmole 1961	<i>Parus major</i> (from Gibb 1954)	High seasonal variation in population size suggested high over-winter mortality	Seasonal survey / census	+
Blondel 1985	3 <i>Parus</i> spp.	Clutch size increased with increasing seasonality of resource availability between mainland and island	Interpopulation comparison	+
Dunn & MacInnes 1985	<i>Branta canadensis</i> in N. America	Negative relationship between clutch size and latitude; unclear relationship between clutch size and productivity	Interpopulation comparison	-
Dunn et al. 2000	<i>Tachycineta bicolor</i> (nc) in N.	Positive relationship between summer productivity and clutch size after controlling for	Interpopulation comparison	+

	America	breeding density		
Evans et al. 2005	11 species in UK and New Zealand	Clutch size of introduced species reduced in less seasonal environment (New Zealand)	Compared clutch size between native and introduced range	+
Griebeler, Caprano & Böhning-Gaese 2010; Griebeler & Böhning-Gaese 2004		Models predict increased clutch size with increased seasonality of resources assuming population size limited by winter resource abundance	ecogenetic individual-based models	+
Jetz et al. 2008	5,290 bird species worldwide	Positive relationship between temperature seasonality and clutch size after controlling for other effects	Interspecific comparison	+
Koenig 1984	<i>Colaptes auratus</i> in N. America	Clutch size negatively correlated with winter AE	Interspecific comparison	+
Koenig 1986	21 N. American species of Order	Clutch size negatively correlated with winter AE	Interspecific comparison	+

Piciformes

Lepage & Lloyd 2004	106 S. African bird species	Clutch size increased with increasing seasonality of rainfall in arid regions	Interspecific comparison	+
McNamara et al. 2008		Model predicts increased clutch size with increasing seasonality assuming population size limited by winter resource abundance	Simulation-based approach using individual-based models	+
Møller 1984	<i>Hirundo rustica</i> and <i>Delichon</i> <i>urbica</i>	Clutch size not related to winter AE or ratio between summer and winter AE	Interpopulation comparison	-
Ricklefs 1980	13 breeding bird communities worldwide	Mean clutch size negatively correlated with winter AE	Compared mean clutch size among localities	+
Yom-Tov, Christie & Iglesias 1994	177 bird species in S. America	Smaller clutch sizes in S. America related to possible reduced climatic variability in temperate S. Hemisphere	Compared pattern of clutch size increase with latitude between N. America and S. America	+

Yom-Tov 1995	Bird species from 5 regions worldwide	Clutch size increased with increasing levels of competition from migrant species	Compared clutch size across regions	+
Young 1994	<i>Troglodytes</i> <i>aedon</i>	Clutch size negatively related to winter AE and positively to AE seasonality but not after controlling for latitude	Interpopulation comparison	+

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651



652 Table 2. Season of greatest mortality and study methodology across 69 year-round studies of  
 653 mortality in wild bird populations. Spring/summer was the season of greatest mortality in radio-  
 654 telemetry studies ( $\chi^2 = 10.5$ , 2 df,  $P = 0.005$ ) but not in capture-recapture studies ( $\chi^2 = 2.0$ , 2 df,  
 655  $P = 0.37$ ). Spring/summer was the season of greatest mortality pooled across study  
 656 methodologies ( $\chi^2 = 6.0$ , 2 df,  $P = 0.050$ ). Data sources are shown in Appendix 1.

Season of greatest mortality	Radio-telemetry	Capture-recapture	Pooled
Spring / summer	21	9	30
Fall / winter	9	6	15
Equivocal (no strong seasonal pattern; migration)	6	12	18
Confounded by hunting <sup>1</sup>	5	1	6
Total	41	28	69

657 <sup>1</sup>Not included in statistical tests – shown for comparison purposes only.

658 Table 3. Relationship between monthly actual evapotranspiration (AE) and monthly survival  
659 rates in 39 wild bird populations. The  $\beta$  and P-value are from a simple linear model of the effect  
660 of log-transformed AE on arc-sine transformed survival rates. Populations in which there was a  
661 significant effect of AE on survival rates are shown in bold. Hunted populations are denoted for  
662 comparison. Data sources are shown in Appendix 1.

Species	Study	$\beta_{AE}$	P	Hunted
Anas fulvigula	Bielefeld & Cox 2006	-0.05	0.42	Y
Colinus virginianus	Burger et al. 1995	<b>0.02</b>	<b>&lt; 0.01</b>	Y
Colinus virginianus	Cox et al. 2004	<b>0.09</b>	<b>&lt; 0.01</b>	Y
Colinus virginianus	Terhune et al. 2007	<b>-0.02</b>	<b>&lt; 0.01</b>	Y
Alectoris chukar	Robinson et al. 2009	<b>-0.03</b>	<b>0.05</b>	Y
Tetrao tetrix	Angelstam 1984	-0.02	0.42	Y
Tetrao tetrix	Caizergues & Ellison 1997	0.01	0.64	Y
Bonasa bonasia	Montadert & Leonard 2003	<b>-0.06</b>	<b>&lt; 0.01</b>	Y
Bonasa umbellus	Devers et al. 2007	<b>0.02</b>	<b>0.03</b>	Y
Bonasa umbellus	Thompson & Fritzell 1989	<b>0.02</b>	<b>&lt; 0.01</b>	Y
Bonasa umbellus	Small et al. 1993	<b>0.02</b>	<b>0.21</b>	Y
Dendragapus canadensis	Herzog 1979	<b>-0.07</b>	<b>0.03</b>	Y
Lagopus lagopus	Smith & Willebrad 1999	0.00	0.68	Y
Centrocercus urophasianus	Sika 2006	<b>0.01</b>	<b>0.09</b>	Y
Tympanuchus pallidicinctus	Hagen et al. 2007	0.01	0.60	Y
Tympanuchus pallidicinctus	Wolfe et al. 2007	-0.01	0.79	
Tympanuchus pallidicinctus	Wolfe et al. 2007	-0.01	0.73	

Meleagris gallopavo	Palmer et al. 1993	-0.01	0.71	Y
Meleagris gallopavo	Wright et al. 1996	<b>-0.01</b>	<b>0.08</b>	Y
Meleagris gallopavo	Lehman et al. 2005	-0.01	0.80	Y
Meleagris gallopavo	Hubbard et al. 1999	-0.01	0.30	Y
Meleagris gallopavo	Kurzejeski et al. 1987	0.00	0.60	Y
Meleagris gallopavo	Nguyen et al. 2003	-0.01	0.46	Y
Meleagris gallopavo	Roberts et al. 1995	0.00	0.15	Y
Meleagris gallopavo	Spohr et al. 2004	-0.02	0.25	Y
Meleagris gallopavo	Vander Haegen et al. 1988	-0.01	0.49	Y
Meleagris gallopavo	Miller et al. 1998	<b>-0.05</b>	<b>&lt; 0.01</b>	Y
Melegaris gallopavo	Vangilder & Kurzejeski 1995	<b>0.00</b>	<b>0.70</b>	Y
Rostrhamus sociabilis	Bennetts & Kitchen 1999	<b>0.11</b>	<b>&lt; 0.01</b>	
Haematopus ostralegus	dit Durrell 2007	-0.04	0.17	
Dryocopus pileatus	Bull 2001	-0.02	0.12	
Petroica goodenovii	Major & Gowing 2001	<b>-0.30</b>	<b>0.08</b>	
Sitta europea	Nilsson 1982	<b>0.04</b>	<b>0.03</b>	
Parus atricapillus	Smith 1967	-0.04	0.64	
Parus atricapillus	Brittingham & Temple 1988	<b>0.04</b>	<b>&lt; 0.01</b>	
Motacilla clara	Piper 2002	<b>-0.07</b>	<b>&lt; 0.01</b>	
Turdus merula	Robinson et al. 2010	-0.02	0.56	
Sylvia boehmi	Schaefer et al. 2006	0.01	0.27	
Sylvia lugens	Schaefer et al. 2006	-0.13	0.17	
Melospiza melodia	Arcese et al. 1992	<b>0.09</b>	<b>0.01</b>	

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663 Figure 1. Graphical representation of Ashmole's hypothesis. Amplitude of seasonal variation in  
664 resource availability is thought to be greater at temperate than tropical latitudes. Adult  
665 population sizes are thought to be regulated by resource availability minima that occur during the  
666 non-breeding season (i.e. winter) which are more severe in temperate regions. Adults that  
667 survive to the summer resource availability maxima are thus thought to have more resources  
668 available for reproduction in temperate regions than tropical regions. Adapted and expanded  
669 from Ricklefs (1980).

670

671 Figure 2. Forest plot of the proportion of known mortalities due to predation in populations of  
672 adults in winter (n=40; references in Appendix 2). Each entry represents a study and the exact  
673 binomial proportion of known mortalities due to predation is shown. Box size corresponds to the  
674 magnitude of the random effects weight,  $W_i$ . The random-effects pooled estimate for  
675 conservative and highly conservative tests (see Methods) is shown at bottom.

676

677 Figure 3. Forest plot of the proportion of known mortalities due to predation in populations of  
678 mixed age classes studied year-round (n=39; references in Appendix 2). Each entry represents a  
679 study and the exact binomial proportion of known mortalities due to predation is shown. Box  
680 size corresponds to the magnitude of the random effects weight,  $W_i$ . The random-effects pooled  
681 estimate for conservative and highly conservative models (see Methods) is shown at bottom.

682

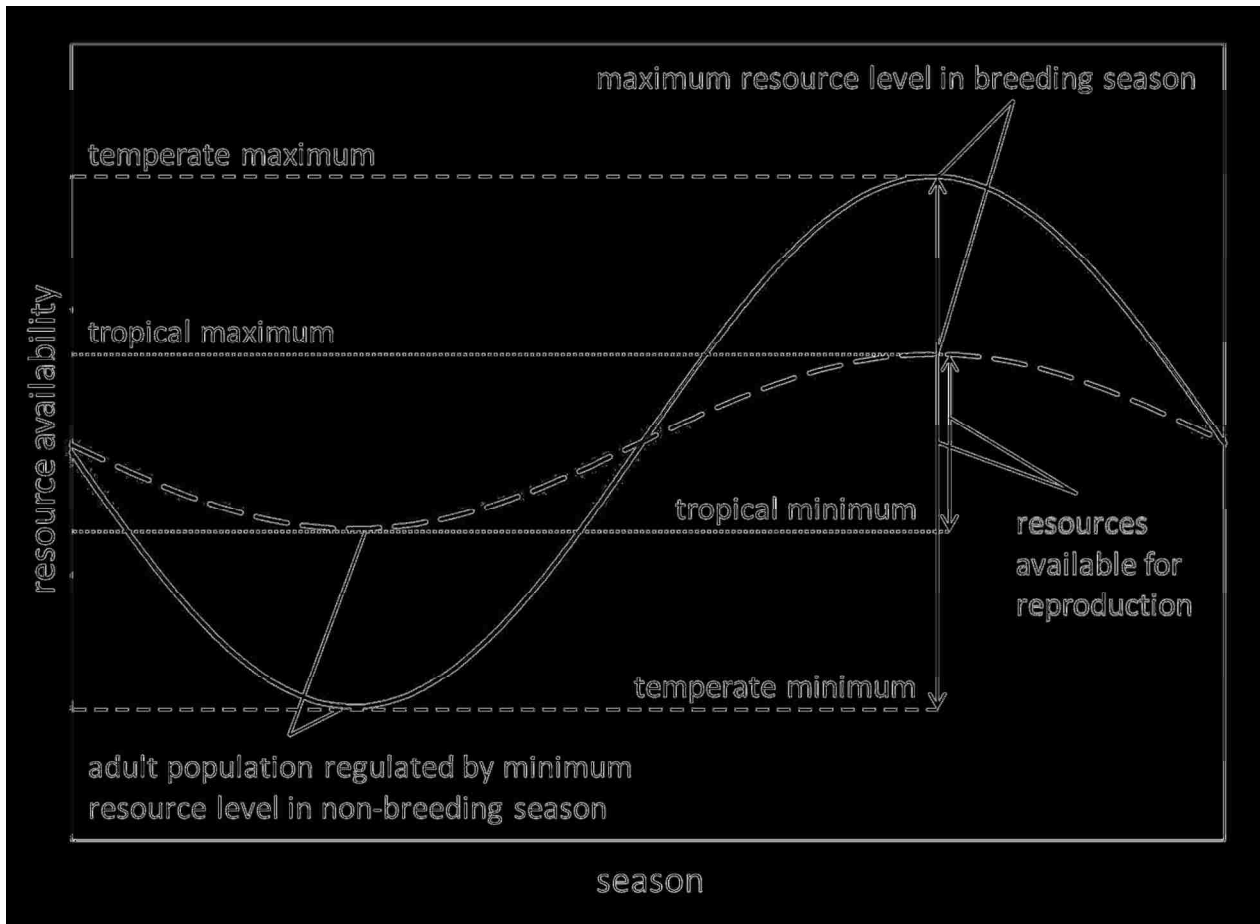
683 Figure 4. Forest plot of the proportion of known mortalities due to predation in populations that  
684 were not categorized as adults in winter, fledglings, or year-round studies (n = 31; Appendix 2).  
685 Each entry represents a study and the exact binomial proportion of known mortalities due to

686 predation is shown. Box size corresponds to the magnitude of the random effects weight,  $W_i$ .  
687 The random-effects pooled estimate for conservative and highly conservative models (see  
688 Methods) is shown at bottom.

689

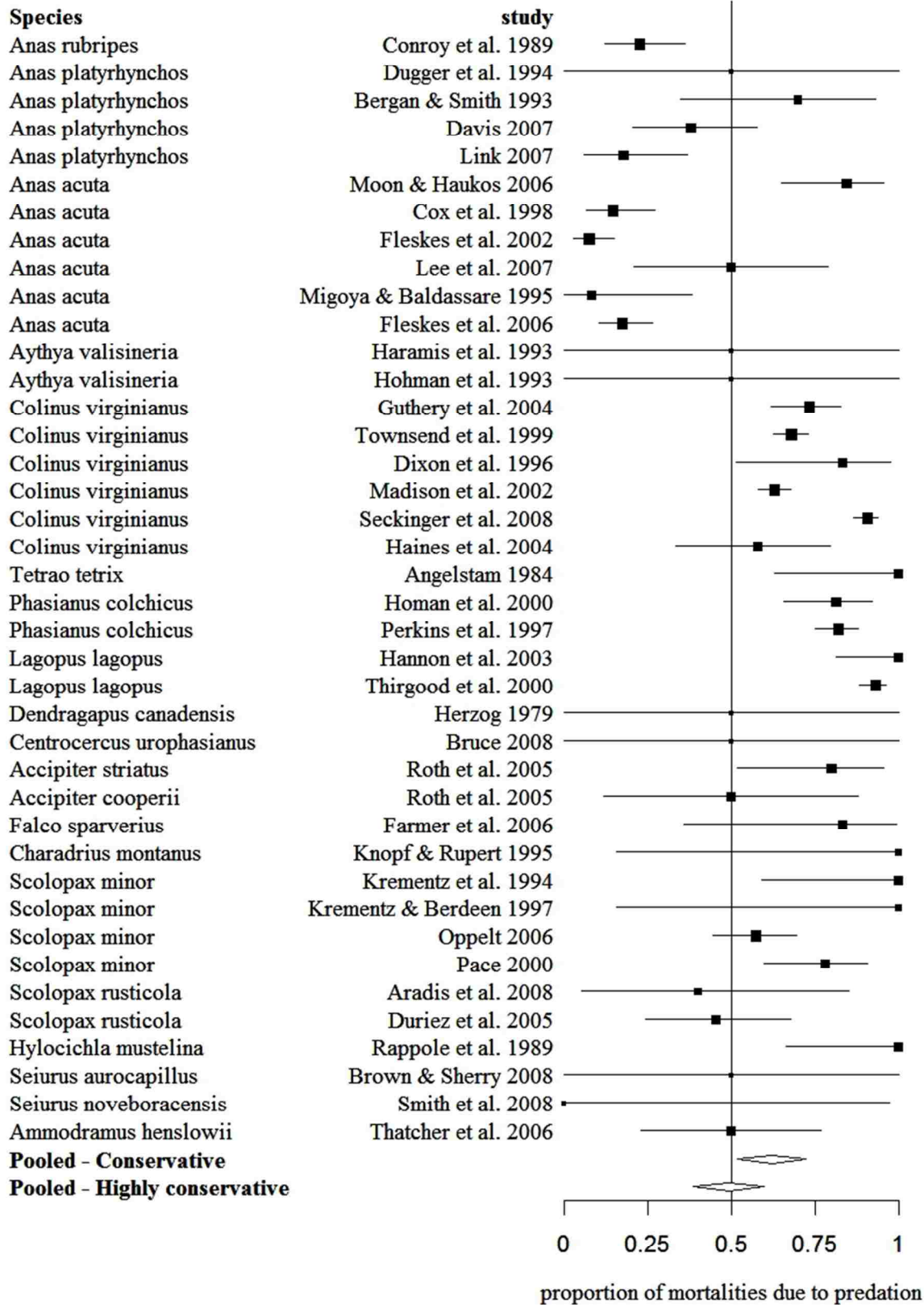
690 Figure 5. Forest plot of the proportion of known mortalities in populations studied daily (open  
691 symbols;  $n=10$ ) or less frequently (closed symbols;  $n=17$ ), shown by species (references in  
692 Appendix 2). Each entry represents a study and the exact binomial proportion of known  
693 mortalities due to predation is shown. Box size corresponds to the magnitude of the random  
694 effects weight,  $W_i$ . The random-effects pooled estimate for each group from a mixed model  
695 weighted by  $W_i$  is shown at bottom.

696



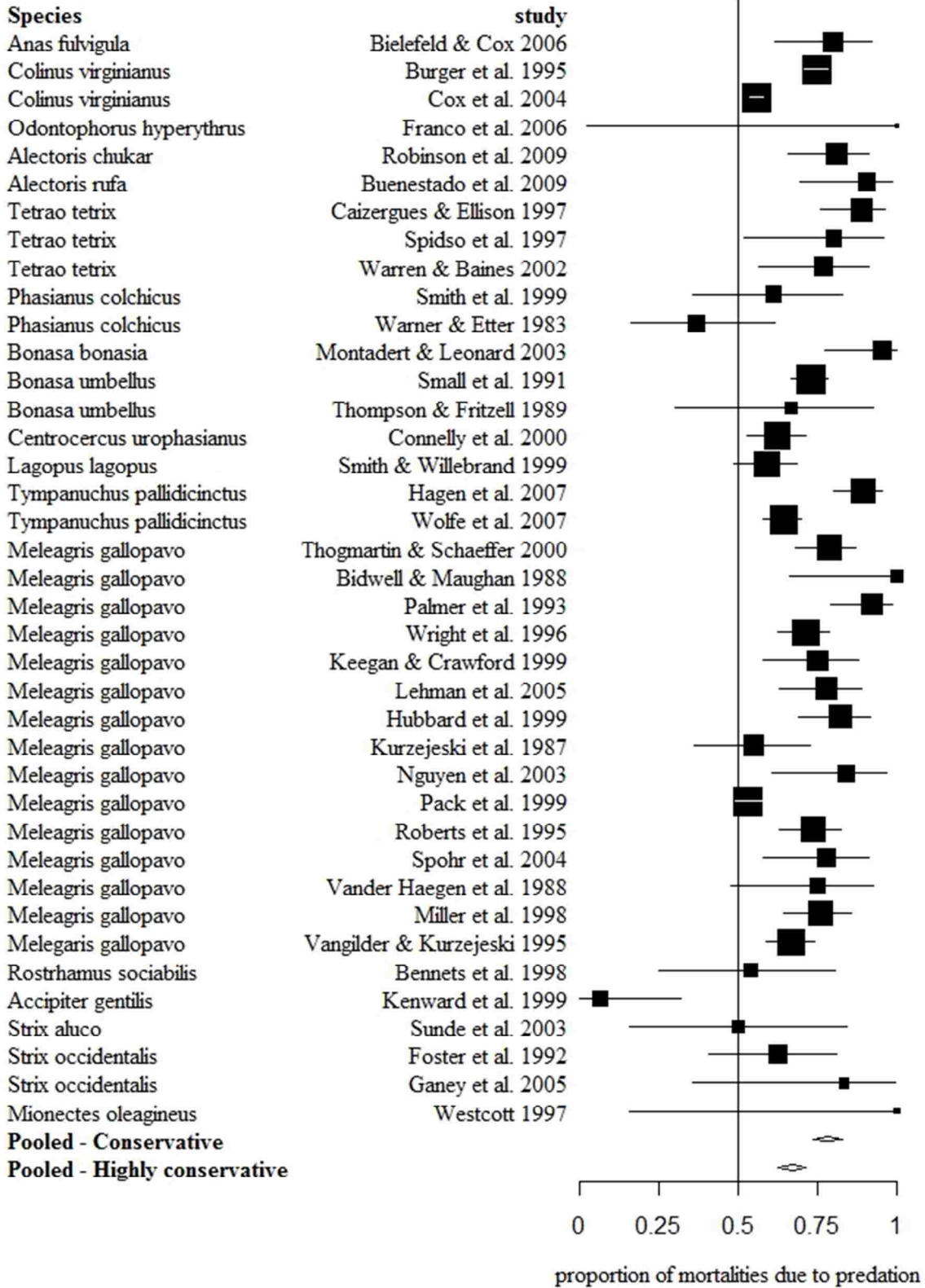
697

698 Figure 1.



699

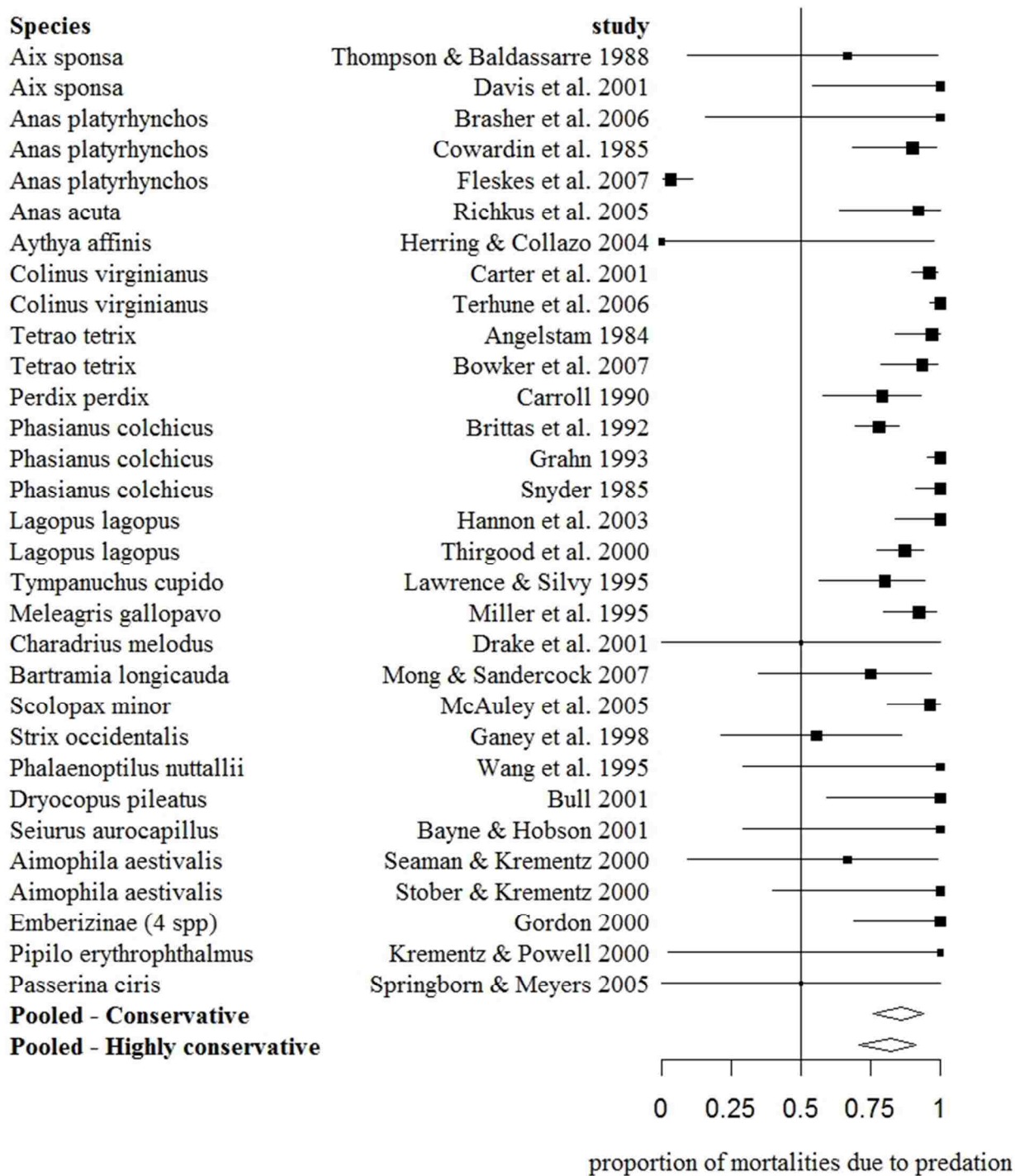
700 Figure 2.



701

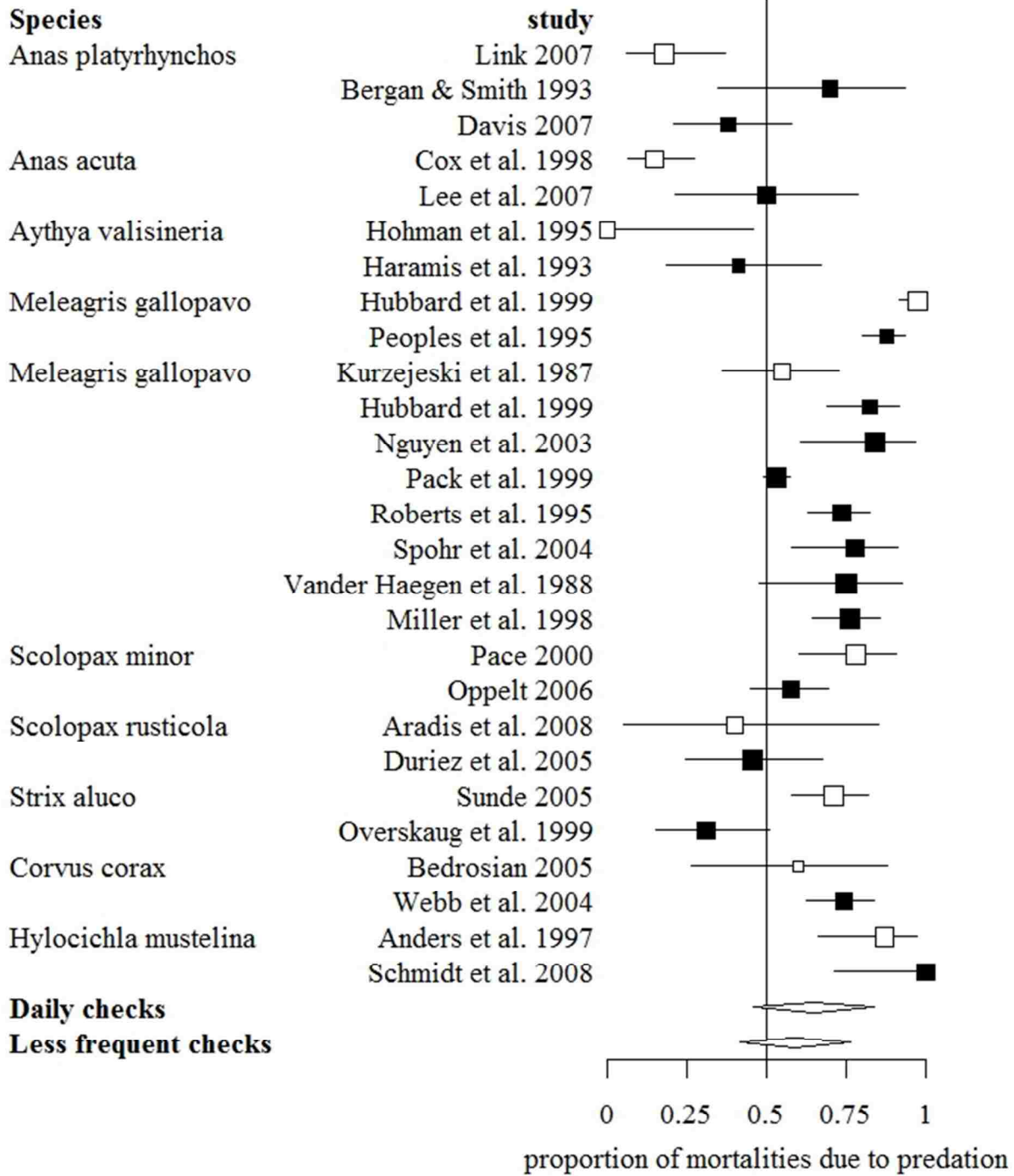
702 Figure 3.





703

704 Figure 4.



705

706 Figure 5.

707 **Chapter 3: Causes of world-wide variation in parental provisioning behavior relative to**  
708 **offspring number**

709

710 Daniel C. Barton, Penn Lloyd, and Thomas E. Martin

711

712 **Abstract**

713 Patterns of variation in reproductive strategies, such as the latitudinal increase in fecundity of  
714 terrestrial birds, are hypothesized to be caused by ecological factors such as resource limitation  
715 or age-specific extrinsic mortality. Numerous classic tests of such hypotheses examined the  
716 responses of parents to natural or experimental variation in offspring number. Yet, such studies  
717 often provided ambiguous tests because of their focus on single species and unnatural  
718 experimental designs that increased offspring number beyond the normal range of phenotypic  
719 variation. We redressed these issues by comparing the responses of songbird parents to  
720 unmanipulated and manipulated variation in offspring number within normal ranges. We  
721 examined among-species variation in the reaction norm between parental provisioning rate and  
722 variation in offspring number for a diversity of bird species on three continents. This reaction  
723 norm integrates critical life history tradeoffs and alternative hypotheses predict differing within-  
724 and among-species patterns of variation. First, variation in the slope of the reaction norm of per-  
725 offspring provisioning to unmanipulated variation in brood size among 29 bird species was  
726 largely explained by variation in adult mortality rate. However, the subset of species with high  
727 adult mortality appeared to adjust offspring number to parental provisioning capacity, as  
728 predicted by food limitation theory. Second, we experimentally reduced broods by  
729 approximately half in 9 bird species with divergent life histories and found that reduced broods  
730 showed increased per-nestling provisioning and reduced total provisioning, consistent with food

731 limitation. However, the magnitude of the decrease in total provisioning rate increased with  
732 decreasing adult mortality among species, suggesting that reproductive effort is differentially  
733 adjusted to reduced brood size based on adult mortality probability. These results suggest that  
734 ecological factors thought to cause life history variation may interact and shift in importance  
735 across species and regions.

736

737 **Key-words** life history, evolution, latitudinal gradient, reproductive effort, provisioning  
738 behavior, parental care, food limitation, nest predation, adult mortality

739

## 740 **Introduction**

741 Explaining variation in parental effort and offspring number is a fundamental goal of life history  
742 theory (Cody 1966, Roff 1992, Stearns 1992). A widely-recognized yet poorly explained pattern  
743 in life history variation is the greater fecundity and parental effort of terrestrial bird species at  
744 higher latitudes (Lack 1947, Martin et al. 2000, Jetz et al. 2008). Replication of this pattern  
745 across diverse radiations of terrestrial birds and across regions suggests an environmental factor  
746 that covaries with latitude is responsible, and almost all such explanations invoke either resource  
747 limitation (Lack 1947, 1954, Ricklefs 2010) or age-specific extrinsic mortality (Williams 1966,  
748 Law 1979, Michod 1979, Martin 2004). A long-enduring context in which many of these  
749 explanations were originally proposed or tested were studies of the response of provisioning rate  
750 to natural or experimental variation in offspring number (Nur 1984, Saether 1984, Linden and  
751 Moller 1989, Golet et al. 1998, Martin 2004).

752         Provisioning of offspring was central to such tests because it is thought to be sensitive to  
753 food availability (Martin 1995) and extrinsic mortality (Skutch 1949, Martin et al. 2000, 2011)

754 while encapsulating tradeoffs between offspring quality and number (Smith and Fretwell 1974)  
755 or between current and future reproduction (Williams 1966). Larger brood sizes require more  
756 provisioning to maintain offspring quality (Lack 1947, Linden and Moller 1989), but such  
757 increased effort may come at the expense of energy, physiological stress, and intrinsic mortality  
758 for parents (Sanz and Tinbergen 1999, Nilsson 2002, Bonier et al. 2011). Thus, the reaction  
759 norm of parental provisioning rate to brood size has long been thought to represent a key  
760 component of life histories and its shape has been proposed to reflect major alternative  
761 explanations for life history variation (Nur 1984, VanderWerf 1992, Conrad and Robertson  
762 1993). Yet, past empirical tests of such explanations have provided only ambiguous support for  
763 any alternative (VanderWerf 1992, Martin 2004). We suggest this ambiguity arises for two  
764 reasons. First, past studies focused on measuring this reaction norm in single species mostly of a  
765 limited range of life history variation, which does not allow generalization to the broader  
766 observed range of life history variation (VanderWerf 1992, Conrad and Robertson 1993, Martin  
767 2004). Second, many past studies focused on experimentally increasing brood size outside of  
768 natural ranges, which exposes parents to novel situations in which their responses are unlikely to  
769 be adaptive (Golet et al. 1998). Thus, the design of previous studies may have yielded results  
770 both difficult to interpret and generalize.

771         We attempted to redress these two issues and to provide a clear test of three major  
772 alternative explanations for latitudinal gradients in reproductive strategy: food limitation (Lack  
773 1947, 1954), nest predation risk (Skutch 1949, Martin et al. 2000), and adult mortality risk  
774 (Williams 1966, Law 1979, Michod 1979, Martin 2004). We expand predictions of previous  
775 studies on single species to a comparative context to increase our ability to discriminate among  
776 alternatives and generalize our results. We compare variation among species in the slope of the

777 reaction norm of per-nestling provisioning rate to offspring number based on unmanipulated in  
778 offspring number kept within natural limits (Fig. 1a). The slope of this reaction norm is  
779 theoretically predicted to vary among species in alternative ways under the food limitation and  
780 adult mortality hypotheses (Nur 1984). In addition, responses of provisioning rate to brood size  
781 may contrast for unmanipulated versus experimental variation. We further develop contrasting  
782 predictions for both how per-offspring provisioning rate is expected to respond to manipulated  
783 brood size under each alternative hypothesis, and for how total provisioning rate (which reflects  
784 parental effort; Nilsson 2002) is expected to respond to manipulated brood size under each  
785 alternative.

786         The food limitation hypothesis posits that, within and among species, parents adjust  
787 offspring number to available food resources (Lack 1954, Nur 1984, Pettifor et al. 1988) because  
788 natural selection favors the clutch size that maximizes the number of surviving offspring given  
789 available food resources (Lack 1947). Over natural variation in brood size, the food limitation  
790 hypothesis thus predicts proportionate increase of provisioning with brood size within species  
791 (i.e. reaction norm slopes of 0; Fig. 1a). When brood size is experimentally reduced, this  
792 hypothesis predicts parents will hold total provisioning rate relatively constant, because parents  
793 set provisioning effort to available food (Lack 1954, Pettifor et al. 1988). As a consequence, per-  
794 offspring provisioning rate would increase in reduced broods, which may increase offspring  
795 quality (Smith and Fretwell 1974).

796         The nest predation hypothesis suggests that since visually-cuing predators may be  
797 attracted to nests by parental activity, high nest predation risk constrains total provisioning rate  
798 and brood size (Skutch 1949, Martin et al. 2000, 2011). This hypothesis predicts that species  
799 with high nest predation risk should increasingly reduce per-offspring provisioning rates with

800 increasing brood size to minimize total provisioning rate and predation risk to offspring (Eggers  
801 et al. 2005). Thus, the slope of the within-species reaction norm of per-offspring provisioning  
802 rate to natural variation in offspring number should be steeper in species higher nest predation  
803 risk (Fig. 1a). Similarly, when brood size is experimentally reduced, total provisioning rate  
804 should decrease with decreased brood size and more strongly in species at higher risk of  
805 predation.

806         The adult mortality hypothesis proposes species with lower adult mortality minimize  
807 costs to self during reproduction to maximize iteroparity, while species with high mortality  
808 maximize current reproductive effort (Williams 1966, Law 1979, Michod 1979, Ghalambor and  
809 Martin 2001). Over natural variation in brood size, this hypothesis predicts that species with low  
810 adult mortality risk should show steeper negative slopes (i.e., larger decreases in per-nestling  
811 provisioning effort with increasing brood size) than species with high adult mortality risk (Fig.  
812 1a). When brood size is experimentally reduced, species with low adult mortality risk are  
813 expected to decrease total provisioning rate more than species with high adult mortality risk.  
814 These predictions reflect that longer-lived species (i.e., those with low adult mortality) should  
815 reduce effort to minimize risk to iteroparity (Williams 1966; Ghalambor and Martin 2001).  
816 Conversely, species with high adult mortality should show little change in effort because  
817 probability of future breeding is low with high adult mortality (i.e. they have little iteroparity to  
818 preserve by reducing effort). Thus, per-offspring provisioning rate is expected to increase in  
819 reduced broods of species with high adult mortality, potentially increasing offspring quality.

820         We tested the predictions of these alternative hypotheses using a comparative-  
821 experimental study of passerine birds. We measured and compared responses of parents to  
822 natural variation in brood size among a wide diversity of bird species representing a wide array

823 of life history diversity at four study sites in North and South America and Africa. We further  
824 conducted experimental reductions in brood size with a subset of bird species at two study sites  
825 in North and South America.

826

## 827 **Methods**

828 We compared responses of parents to natural variation in brood size among 29 different  
829 bird species (Fig. 2) from Arizona (34° N), South Africa (34° S), Argentina (28° S), and  
830 Venezuela (8° N), and to experimental reductions in brood size among 9 different bird species  
831 from Arizona and Venezuela (Fig. 2). These study sites and further details of the natural history  
832 and life history traits of many of the species included in this study have also been described  
833 previously (e.g. Ghalambor and Martin 2001, Martin et al. 2006, Martin et al. 2011).

834 We measured parental provisioning rate (trips/hr) and per-nestling provisioning rate  
835 (trips/hr/nstl) using videotaped observations of nests (4-8 h in length) during the nestling period  
836 in Arizona (1993-2006), South Africa (2000-2004), Argentina (1997-2000), and Venezuela  
837 (2000-2008). During each year at each site, we located, monitored, and measured parental care  
838 at nests following standardized methodology (Martin and Geupel 1993, Martin et al. 2006).  
839 Offspring number was determined by direct observation of each nest, and nestling age was  
840 determined using observation of critical developmental transitions. Provisioning rates generally  
841 increase with nestling age (Lyon et al. 1987; Martin et al. 2011), and we therefore divided  
842 nestlings into 3 age groups based on major developmental transitions: before, during, or after  
843 'pin-break', the emergence of flight feathers from feather sheaths. Because 'pin-break'  
844 represents a major developmental transition in altricial birds near completion of musculoskeletal



845 development, we standardized our reported results to this age to allow comparisons among  
846 species with different nestling periods (Martin et al. 2011).

847 We experimentally manipulated brood size in 9 bird species (Fig. 2) at the Venezuela  
848 (2005-2008) and Arizona (2007-2010) study sites. To manipulate brood size, we removed  
849 approximately half the eggs in a complete clutch and replaced them with dummy eggs. We also  
850 reduced brood size by inserting a thermocouple into a single egg in a clutch as part of another  
851 study conducted at the same study sites (e.g. Martin et al. 2007). We removed dummy eggs or  
852 the egg containing the thermocouple at the time of hatching. Experimentally reduced nests were  
853 filmed every other day following hatching until failed or fledged, and were matched with a  
854 within-season control nest with the same initial clutch size when possible.

855 We estimated the slope of the relationship between per-offspring provisioning rate and  
856 offspring number using ANCOVA with per-offspring provisioning rate as the response variable,  
857 species and nestling age as factors, and species by nestling age and species by brood size  
858 interactions. We used parameter estimates of the species by brood size interaction as species-  
859 specific slopes of the relationship between per-offspring provisioning rate and offspring number  
860 while controlling for the effect of age. We used the t-scores of the parameter estimates of the  
861 species by brood size interaction to test whether species-specific slopes differed from 0.

862 We tested the effect of experimental brood size reductions on provisioning rates using  
863 two different ANOVA approaches. First, we tested whether the natural log of total provisioning  
864 rate and the natural log of per-nestling provisioning rate at 'pin break' ( $\pm 1$  day) differed between  
865 within-season matched-pair treatment (reduced) and control nests. We used an ANOVA design  
866 containing the treatment by species interaction (our test of whether response varied among  
867 species), species, and a blocking variable for pair. Second, because not all reduced nests were

868 (or could be, in the case of odd-numbered clutch sizes) reduced by exactly half, we used an  
869 ANCOVA design containing an interaction between species and a covariate for the proportionate  
870 magnitude of treatment (range: 0-0.66), species, and a blocking variable for pair. We then used  
871 this model to predict the response of each species to an average brood size reduction of 50% to  
872 standardize the magnitude of treatment for across species comparisons.

873 We quantified offspring quality in manipulated (reduced) and control broods by  
874 measuring nestling mass, which is an important predictor of juvenile survival in altricial birds  
875 (Nur 1984, Golet et al. 1998). We measured mass of all nestlings in a nest using a digital  
876 balance at 'pin break' age ( $\pm 1$  day; see above). We tested the effect of experimental brood size  
877 reductions on nestling mass and tarsus using an ANCOVA model containing effects of treatment  
878 by species interaction, age by species interaction, species, and a within-subject error term for  
879 nest (because observations of nestlings within each nest are not independent from each other).

880 To determine nestling predation, we monitored nests following established protocols  
881 (Martin and Geupel 1993) to determine the number of days each nest was active and to assess  
882 success or failure. We estimated daily predation risk during the nestling period using the  
883 Mayfield method (Mayfield 1975, Johnson 1979). To determine adult mortality rates, we  
884 captured and individually marked adult birds using individual combinations of colored and  
885 numbered leg bands, and systematically recaptured and visually re-sighted marked adults  
886 throughout each study area each year. Adult survival rates were estimated using Cormack-Jolly-  
887 Seber or multistate mark-recapture models in Program MARK (Lebreton et al. 1992, White and  
888 Burnham 1999, Chapter 3). We tested for predicted correlations between estimated reaction  
889 norm slopes and responses to brood size manipulations using Pearson's correlation coefficient.  
890 We controlled for the effects of a phylogenetic hypothesis (Figure 2) on correlated evolution in

891 the continuous characters we measured (adult and nestling mortality rate, reaction norm slope)  
892 using phylogenetic independent contrasts (Felsenstein 1985) as implemented in the PDAP  
893 Package (Midford et al. 2009) for Program Mesquite (Maddison and Maddison 2011). We used  
894 published phylogenetic hypotheses (Jönsson and Fjeldså 2006, Davis 2008) to develop our  
895 phylogenetic hypothesis (Figure 2).

## 896 **Results**

897 The relationship between per-nestling provisioning rate and natural variation in brood  
898 size varied significantly among species (Fig. 1b). Estimates of reaction norm slopes of per-  
899 offspring provisioning rate were negative for all 29 species examined and significantly less than  
900 0 for 16 species. Slopes that did not differ from 0 could reflect low statistical power in some  
901 cases, but nevertheless, some species appear to adjust brood size proportionately to available  
902 food while others do not (Fig. 1b), lending mixed support to the food limitation hypothesis.  
903 Differences among species in reaction norm were not correlated with among-species differences  
904 in nest predation risk (Figure 3a), contrary to the prediction of the nest predation hypothesis.  
905 Differences among species in reaction norm slope were highly correlated with among-species  
906 differences in adult mortality rates (Figure 3b), supporting the prediction of the adult mortality  
907 hypothesis. Neither of these correlations was affected by the potentially confounding influence  
908 of phylogeny (Figure 2, Appendix 3).

909 We experimentally reduced brood size at 48 nests of 9 species that survived to ‘pin-  
910 break’ age (many more were reduced and failed due to predation) matched with 48 control nests  
911 within year and season. The natural log of total provisioning rate varied significantly among  
912 species, as did the response of the natural log of total provisioning rate to experimental reduction  
913 (Figure 4a). The three tropical (Venezuela) species showed a significant reduction in total

914 provisioning rate in response to the brood reduction treatment while the six temperate (Arizona)  
915 species did not (Figure 4). The natural log of per-nestling provisioning rate also varied  
916 significantly among species as did the response of natural log of per-nestling provisioning rate to  
917 experimental reduction (Figure 5a). The 6 temperate species showed a significant increase in  
918 per-nestling provisioning rates in response to reductions (t-test from ANOVA,  $P < 0.05$ ; Figure  
919 5a) while the 3 tropical species did not (t-test from ANOVA,  $P > 0.68$ ; Figure 5a). Reduced  
920 broods showed increased nestling mass at 'pin break' age in 4 of the 6 Arizona species (Figure  
921 5b) while data were not sufficient for tests of reduced brood size on nestling mass in Venezuela  
922 species.

923         Magnitude of response of the natural log of total provisioning rate to brood size reduction  
924 varied from a -2.0% change (*J. hyemalis*) to a -50.5% change (*B. tristriatus*). Because the  
925 magnitude of brood size manipulation varied among treatment-control pairs from -33% to -66%  
926 and among species from -35.2% to -57.1%, we used an ANCOVA approach to predict species  
927 responses to a standardized brood size reduction of 50%. The slope of the relationship between  
928 brood size reduction magnitude and the natural log of provisioning rate varied significantly  
929 among species ( $F_{\text{species} \times \text{magnitude}} 9, 39 = 6.9606, P < 0.001$ ) and was significantly different than 0 in  
930 three species (t-test from ANCOVA,  $P < 0.01$ ) and approached significance in two species (t-test  
931 from ANCOVA,  $0.05 < P < 0.10$ ). The predicted responses of the natural log of provisioning  
932 rate to a 50% brood size reduction varied from -1.3% (*P. chlorurs*) to -50.5% (*B. tristriatus*;  
933 Figure 4b). The predicted among-species responses to a 50% reduction from this ANCOVA  
934 approach and the estimated response of species to brood size reductions uncorrected for  
935 magnitude from ANOVA were highly correlated ( $N = 9, r = 0.99, P < 0.001$ ).

936           The three tropical Venezuelan species with the lowest adult mortality, *B. tristriatus*, *P.*  
937 *flavipes*, and *H. leucophrys*, showed the only statistically significant (Figure 4a) and the largest  
938 (Figure 4b) reductions in total provisioning rate in response to brood size reduction. Differences  
939 among species in percentage reduction of total provisioning rate were not correlated with among-  
940 species differences in nest predation risk ( $N = 9$ ,  $r < 0.2$ ,  $P > 0.5$ ), providing no support for the  
941 nest predation hypothesis. Differences among species in percentage reduction of total  
942 provisioning rate were highly correlated with among-species differences in adult mortality risk  
943 (Figure 4b), supporting the adult mortality hypothesis. This correlation was independent of  
944 phylogeny (Figure 2, Appendix 3).

945

## 946 **Discussion**

947           We found that a major prediction of food limitation theory – the apparent adjustment of  
948 brood size to parental provisioning capacity (Lack 1947, 1954) – was upheld in species with high  
949 adult mortality risk in both comparative (Figures 1b, 3b) and experimental (Figure 4) study  
950 designs. Species with high adult mortality risk apparently proportionately adjusted brood size  
951 and provisioning rate over the range of unmanipulated variation (Figure 1b; Pettifor et al. 1988).  
952 Species with high adult mortality risk also continued to provision at a high rate when brood size  
953 was experimentally reduced (i.e. were relatively unresponsive to manipulation; Figure 4) and  
954 showed increased per-nestling provisioning rate in experimentally reduced broods. Reduced  
955 broods in 4 of the 6 temperate species that showed increased per-nestling provisioning rate also  
956 showed increased nestling mass (Figure 4), suggesting benefits of increased per-offspring  
957 provisioning for offspring quality (Nur 1984; Pettifor et al. 1988).

958           However, species with low adult mortality risk showed alternative responses of  
959 provisioning rate to natural and experimental variation in brood size inconsistent with the  
960 predictions of the food limitation hypothesis. Responses of parents to natural and experimental  
961 variation in brood size varied across a gradient of adult mortality risk (Figures 1b, 3b). Species  
962 with high adult mortality appeared to show high parental effort across natural variation in brood  
963 size and maintained high parental effort when brood size was reduced. In contrast, species with  
964 low adult mortality show reduced per-nestling provisioning rate with increasing natural brood  
965 size and reduced total provisioning rate (and similar per-nestling provisioning rate) with  
966 experimentally reduced brood size. Surprisingly, we did not see evidence that nest predation risk  
967 affected the responses of parents to natural or experimental variation in brood size, despite the  
968 known effects of nest predation risk on mean provisioning rates in songbirds (Martin et al. 2000;  
969 Martin et al. 2011).

970           Food limitation may thus appear highly important in constraining the expression of life  
971 history traits in some contexts – such as within north-temperate regions, where much previous  
972 research has been conducted (Martin 2004) – but not in others, such as across regions including  
973 the tropics. Food limitation may similarly vary in importance for determining the outcome of  
974 life history microevolution (Walsh and Reznick 2008). These results may help resolve the  
975 disconnect between studies that supported food limitation as an explanation for life history  
976 variation in birds (Lack 1954, VanderWerf 1992, Jetz et al. 2008, Ricklefs 2010) and other  
977 studies that suggest food limitation is an insufficient explanation for life history variation  
978 (Owens and Bennett 1995, Ferretti et al. 2005).

979           Reviews of previous experimental brood size manipulation experiments have noted  
980 responses to brood size manipulation are highly variable among species (Linden and Moller

981 1989, VanderWerf 1992). Explanations for such wide variation in responses are likely two-fold:  
982 first, many previous studies focused on experimental increases beyond the range of natural brood  
983 size variation which expose parents to novel situations in which their responses are unlikely to be  
984 adaptive (Golet et al. 1998). Second, variation among species in responses is likely to be partly  
985 caused by differences among species in their life history traits (VanderWerf 1992). Our study  
986 confirms this idea, and we show how life history variation – expressed as adult mortality rates –  
987 may explain among-species differences in response to brood size manipulations (Figures 4, 5)  
988 and further show how such among-species differences may be used to test alternative  
989 explanations for life history variation.

990         We see three alternative explanations for variation among species in provisioning effort  
991 that is correlated with adult mortality rates, such as the results reported here. First, variation  
992 among species in reproductive traits may be a consequence of extrinsic adult mortality risk as  
993 suggested by classic theory that suggests low extrinsic mortality favors reduced reproductive  
994 effort to preserve iteroparity (Williams 1966, Law 1979, Michod 1979, Martin 2004). Second,  
995 seasonal resource availability dynamics may impose winter mortality and then provide high  
996 levels of food availability in the breeding season in temperate regions as predicted by Ashmole’s  
997 hypothesis (Ricklefs 2010). Third, adult mortality rates may be a proximate consequence of  
998 variation in provisioning effort (Stearns 1992). The second explanation is likely insufficient  
999 because terrestrial bird species do not generally show the patterns of mortality predicted by  
1000 Ashmole’s hypothesis (Chapter 1). Our results suggest that the third explanation is insufficient  
1001 because we found that longer-lived species adjust their reproductive effort, as measured by  
1002 provisioning rate, downwards in response to brood size reductions (Figure 4). This result is not  
1003 expected if variation among species in adult mortality rates is simply a consequence of

1004 reproductive effort. Thus, we suggest that variation among species in adult mortality risk may  
1005 explain variation in reproductive effort as measured by provisioning behavior or clutch size as  
1006 proposed by Williams (1966).

1007         Our results suggest that adult mortality risk plays an over-arching role in determining  
1008 how species resolve critical tradeoffs between current and future reproduction and between  
1009 number and quality of offspring. Yet, species with high adult mortality rate may resolve the  
1010 tradeoff between current and future reproduction in favor of maximizing current reproduction  
1011 and thus encounter the constraints of food limitation. We suggest that food limitation shifts in  
1012 importance as an interacting function of adult mortality risk across geographic regions.

1013

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1113 **Figure legends**

1114 Figure 1. Among-species variation in reaction norms of per-offspring provisioning rate to  
1115 natural variation in offspring number. **a**, Three potential reaction norms of per-offspring  
1116 provisioning rate to offspring number. When the slope of this reaction norm is 0, parents  
1117 increase total provisioning rate proportionally with increased offspring number, resulting in  
1118 constant per-offspring provisioning. When the slope of this reaction norm is less than 0, parents  
1119 increase total provisioning rate slower (less than proportional) with offspring number, thus per-  
1120 offspring provisioning declines with increasing offspring number. Food limitation theory  
1121 predicts slopes of 0 (e.g. line 1). The nest predation alternative predicts steeper slopes in species  
1122 with higher nest predation (e.g. line 3 vs. line 2), while the adult mortality alternative predicts  
1123 steeper slopes in species with lower adult mortality. **b**, Estimated reaction norms of per-nestling  
1124 provisioning rate to natural variation in brood size for 29 passerine bird species from 4 study  
1125 sites (N = 1644 observations). Slope varied significantly among species ( $F_{\text{species} \times \text{brood size } 29, 1529} =$   
1126 9.74,  $P < 0.001$ ). The point estimate of all slopes is negative and for 16 is significantly different  
1127 from 0 (t-test from ANCOVA,  $P < 0.05$ ).

1128

1129 Figure 2. Phylogeny, standard name, geographic site, and life history traits of 29 bird species  
1130 studied. Annual adult mortality rate was estimated using Cormack-Jolly-Seber models of mark-  
1131 resight-recapture data collected from the study sites. Nestling predation risk was estimated for  
1132 each species using the Mayfield method based on large sample sizes from each species and study  
1133 site. Reaction norm slopes are the slope of the linear relationship between per-nestling  
1134 provisioning rates and offspring number estimated by ANCOVA. Slopes of reaction norms  
1135 significantly different from 0 (t-test from ANCOVA,  $P < 0.05$ ) are shown in bold. Phylogeny

1136 branches and species data are color-coded by geographic site. Branches basal to multiple  
1137 geographic sites are shown in black. Names of 7 of the 9 species for which brood size was  
1138 manipulated are starred, while two (*Turdus migratorious*, *Poecile montanus*) are not shown  
1139 because they were only included in the experimental portion of the study. Phylogeny based on  
1140 the supertrees of Jønsson and Knud (2006) and Davis (2008).

1141

1142 Figure 3. Among-species variation in reaction norms of per-offspring provisioning rate to  
1143 natural variation in offspring number and relationship to ecological factors. **a**, Reaction norm  
1144 slope and daily nest mortality were not correlated among species (n= 29, Pearson correlation,  $r =$   
1145  $-0.03$ ,  $P = 0.86$ ) contrary to the nest predation hypothesis. **b**, Reaction norm slope (from Figure  
1146 1b) and annual adult mortality were highly correlated among species (n = 29, Pearson  
1147 correlation,  $r = 0.68$ ,  $P < 0.001$ ), supporting the adult mortality hypothesis. The reported  
1148 correlations were not strongly influenced by the potentially confounding effects of phylogeny  
1149 (Appendix 3).

1150

1151 Figure 4. Among-species variation in responses of total provisioning rate to experimental brood  
1152 size reductions and relationship to adult mortality risk. **a**, Species-specific total provisioning rate  
1153 in control and reduced broods (by about 50%) in 9 bird species in Arizona and Venezuela.  
1154 Closed symbols are control group means, and open symbols are reduction group means. Sample  
1155 sizes (N = number of treatment-control pairs): *J. hyemalis* (7), *O. celata* (5), *P. chlorurus* (9), *P.*  
1156 *montanus* (4), *T. aedon* (4), *T. migratorious* (5), *B. tristriatus* (5), *P. flavipes* (5), *H. leucophrys*  
1157 (4). Each species-specific treatment-control pair is connected and labeled by color.  
1158 Provisioning rate varied among species ( $F_{\text{species } 8, 39} = 59.03$ ,  $P < 0.001$ ) and response of

1159 provisioning rate to reduction varied among species ( $F_{\text{species} \times \text{treatment } 9, 39} = 6.82, P < 0.001$ ).

1160 Significant reductions in provisioning rate in response to reductions were observed in 3 of 9

1161 species (indicated by \*; t-test from ANOVA,  $P < 0.01$ ) and response approached significance in

1162 2 of 9 species (indicated by †; t-test from ANOVA,  $0.05 < P < 0.10$ ). **b**, Comparison of

1163 predicted percent reduction in total provisioning rate in response to a standardized 50% brood

1164 size reduction among species in relation to annual adult mortality (from ANCOVA). Percent

1165 reduction was significantly correlated with annual adult mortality rate ( $N = 9$ , Pearson

1166 correlation,  $r = 0.906, P < 0.001$ ), supporting the adult mortality hypothesis. Significant

1167 reductions are indicated by \* (t-test from ANOVA,  $P < 0.01$ ) and † (t-test from ANOVA,  $0.05 <$

1168  $P < 0.10$ ). This result was independent of phylogenetic effects (Appendix 3).

1169

1170 Figure 5. Among species variation in responses of per-nestling provisioning rate and offspring

1171 mass to experimental brood size reductions. **a**, Species-specific per-nestling provisioning rate in

1172 control and reduced broods (by about 50%) in 9 bird species in Arizona and Venezuela. Closed

1173 symbols are control group means, and open symbols are reduction group means. Sample sizes

1174 same as shown in Figure 4. Each species-specific treatment-control pair is connected and

1175 labeled by color. Response of per-nestling provisioning rate to reduction varied among species

1176 ( $F_{\text{species} \times \text{treatment } 9, 39} = 9.80, P < 0.001$ ). Significant increases in per-nestling provisioning rate in

1177 response to reductions were observed in 6 of 9 species (indicated by \*; t-test from ANOVA,  $P <$

1178  $0.01$ ). **b**, Species-specific nestling mass at ‘pin break’ in control and reduced (by about 50%)

1179 broods of 6 bird species in Arizona. Symbols as in panel **a**. Sample sizes (treatment, control

1180 nests): *J. hyemalis* (7,6), *O. celata* (5,9), *P. chlorurus* (8,10), *P. montanus* (3,4), *T. aedon* (3,6),

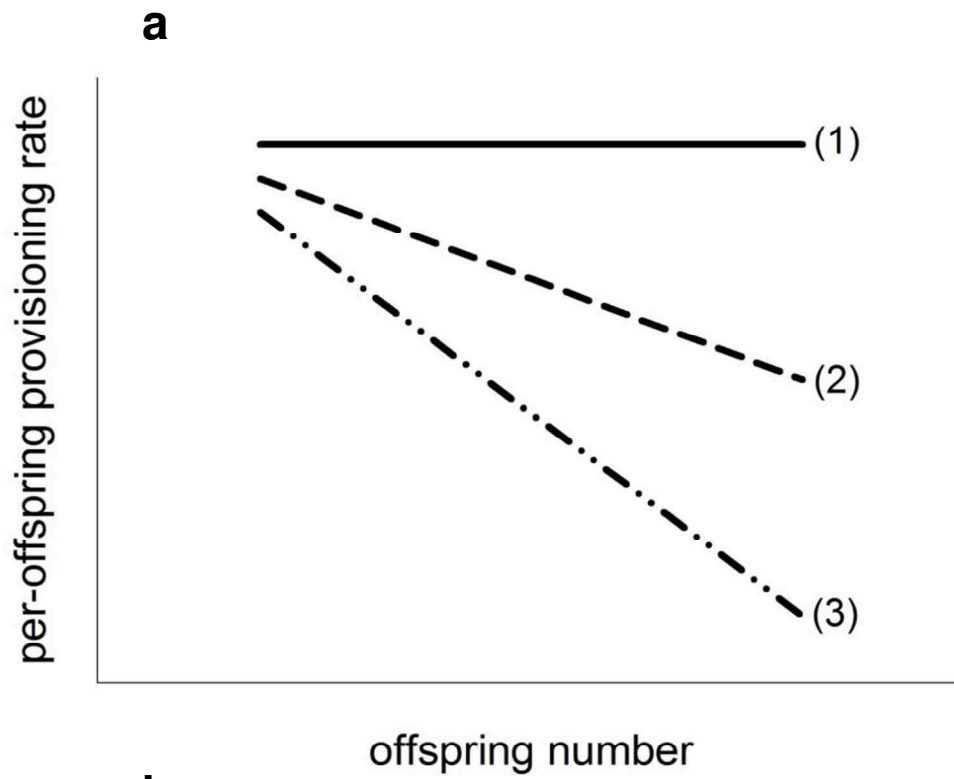
1181 *T. migratorious* (4,4). Response of nestling mass to brood size reduction varied among species

1182 ( $F_{\text{species} \times \text{treatment}} = 6.81, P < 0.001$ ). Significant increases in nestling mass were observed in 3  
1183 of 6 species (indicated by \*; t-test from ANOVA,  $P < 0.05$ ) and approached significance in 1 of 6  
1184 species (indicated by †; t-test from ANOVA,  $0.05 < P < 0.10$ ).

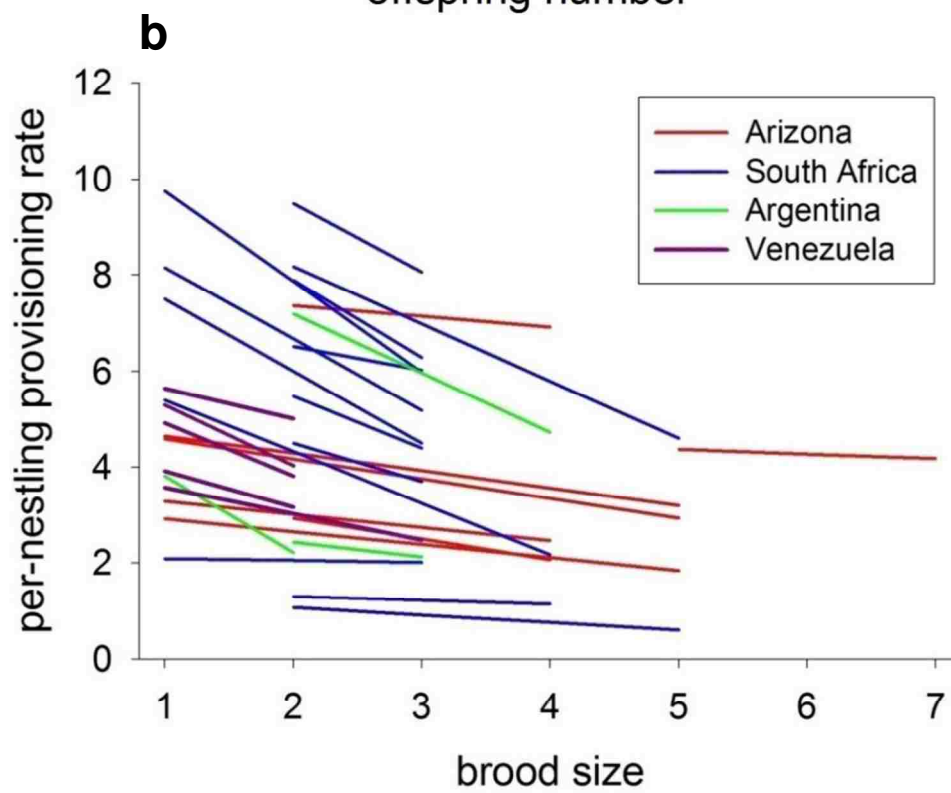
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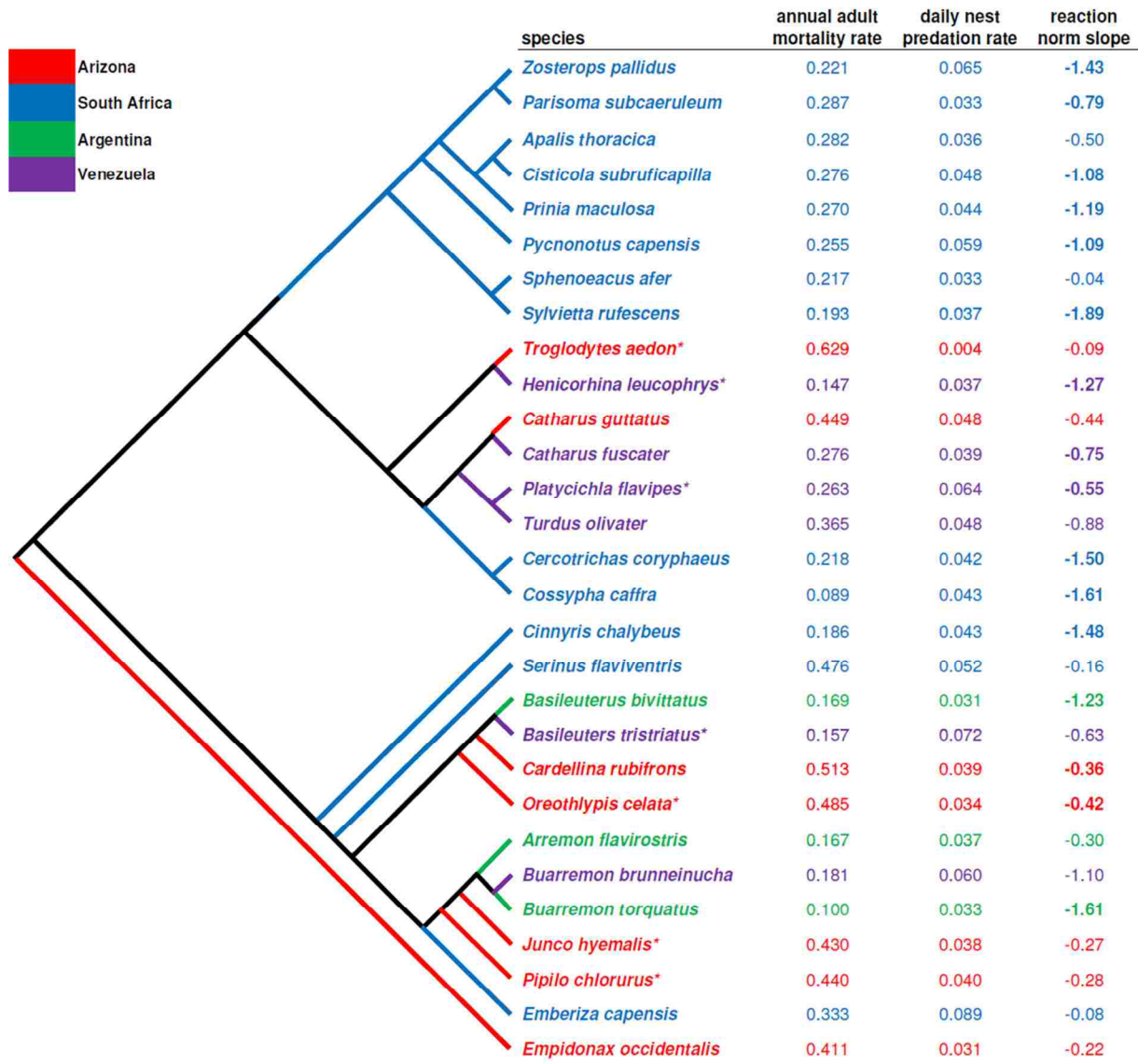


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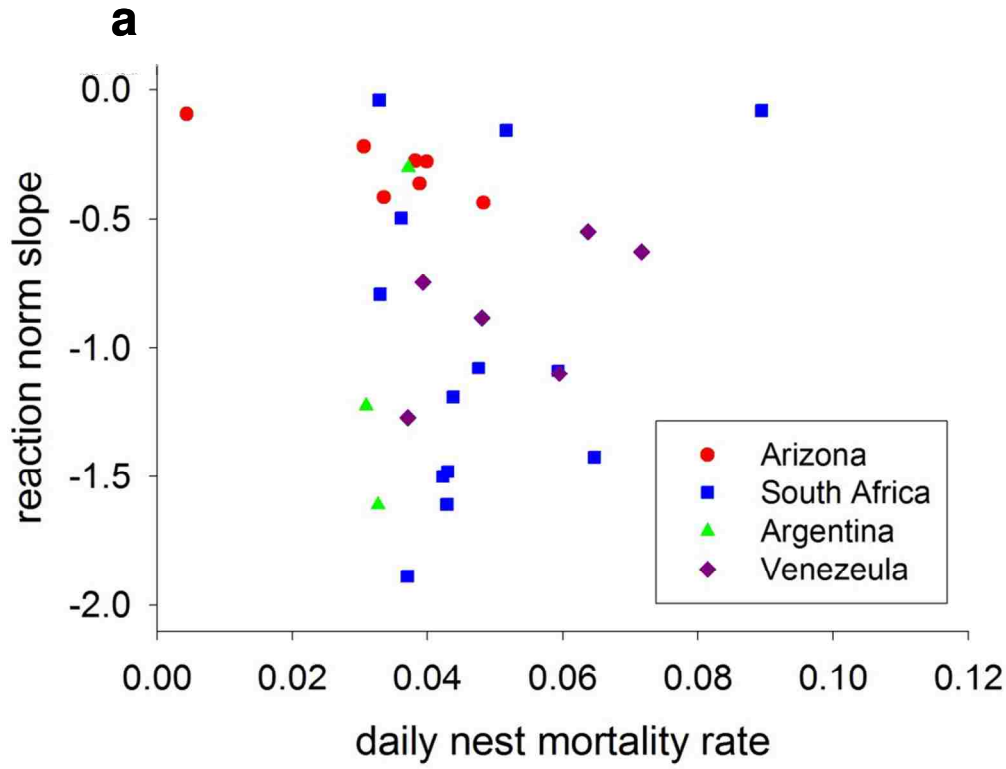
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1189 Figure 1.

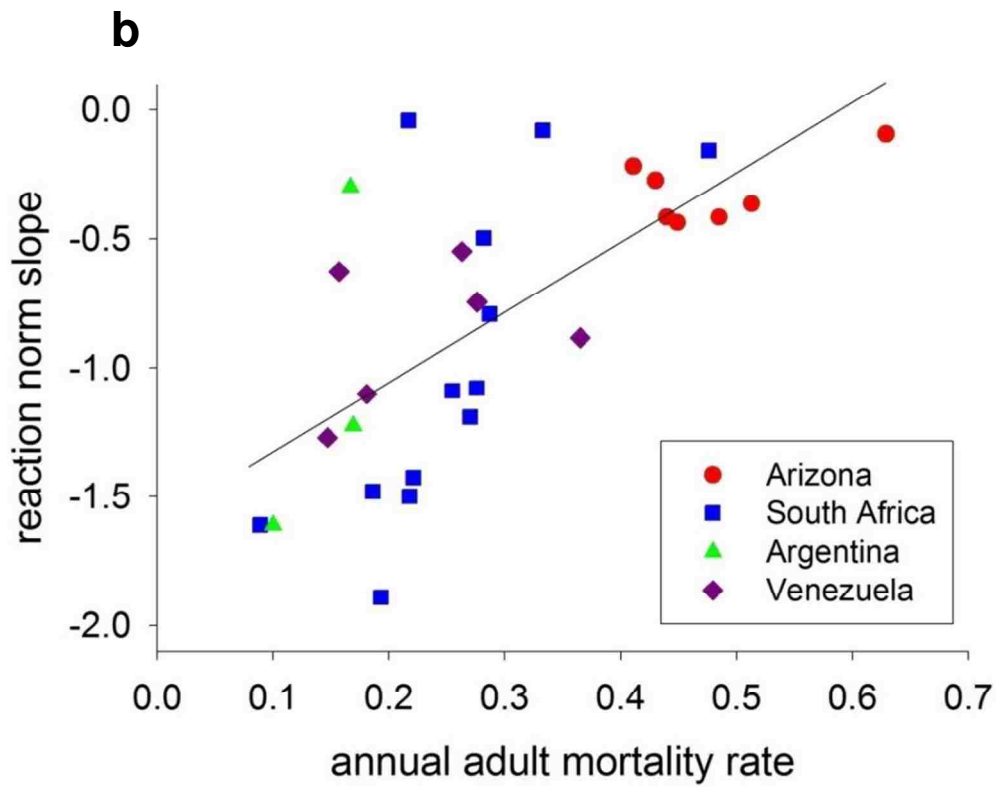


1190

1191 Figure 2.

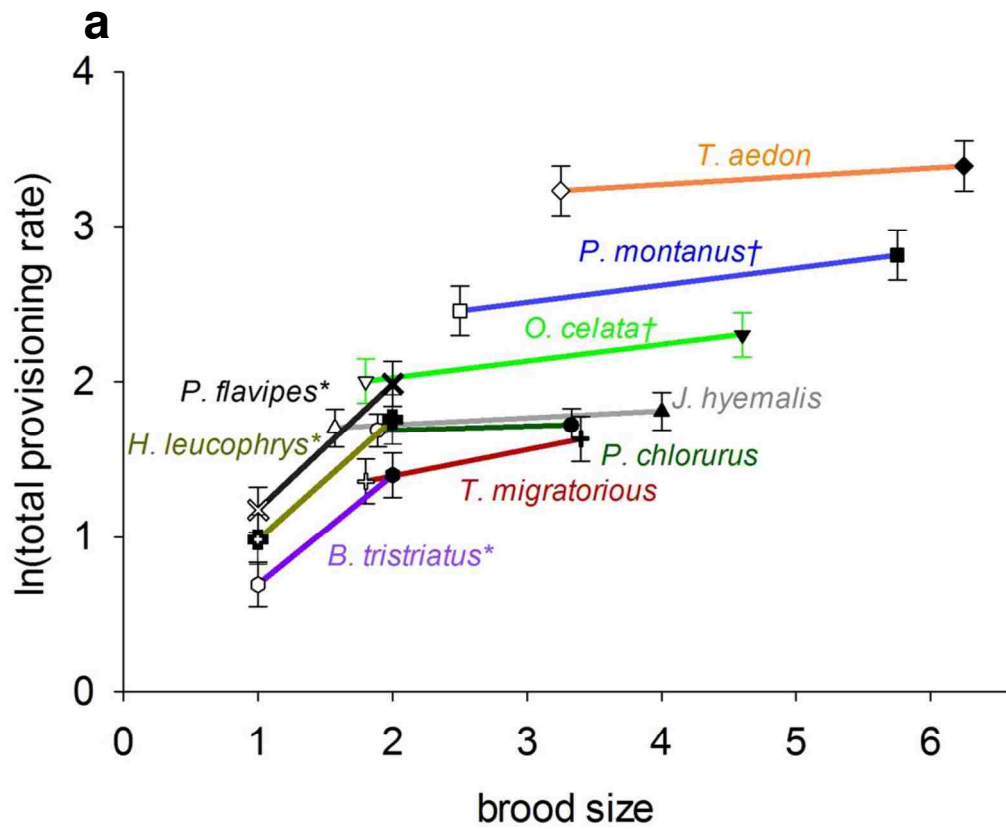


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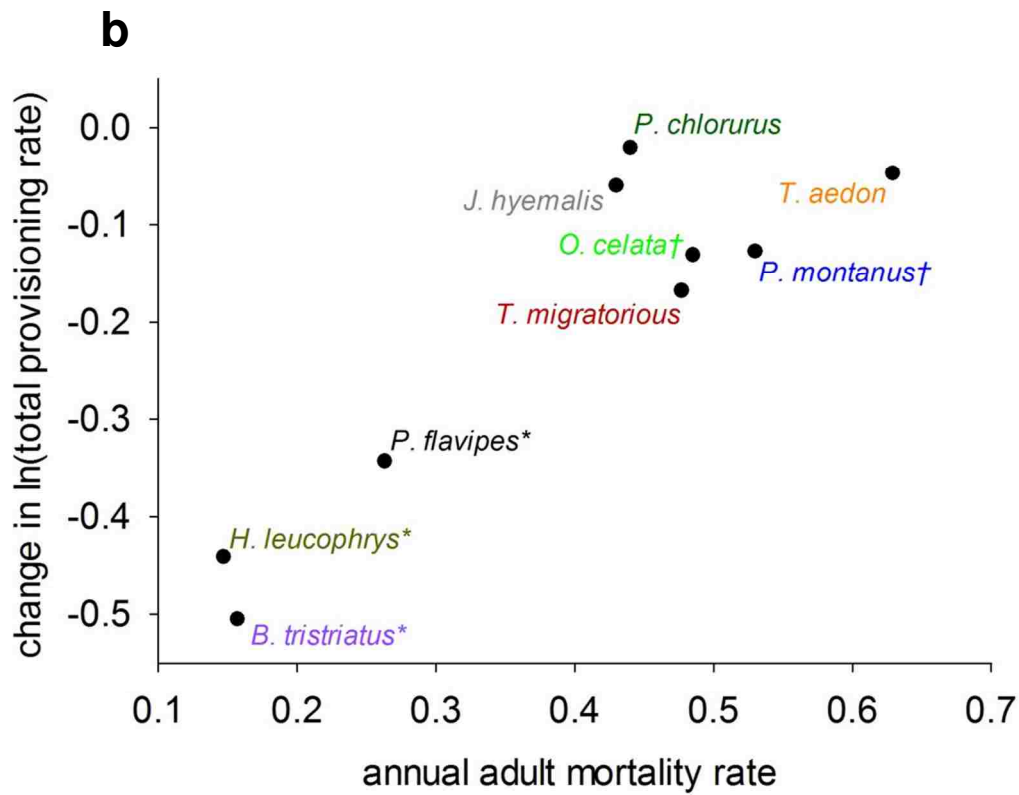


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1194 Figure 3.

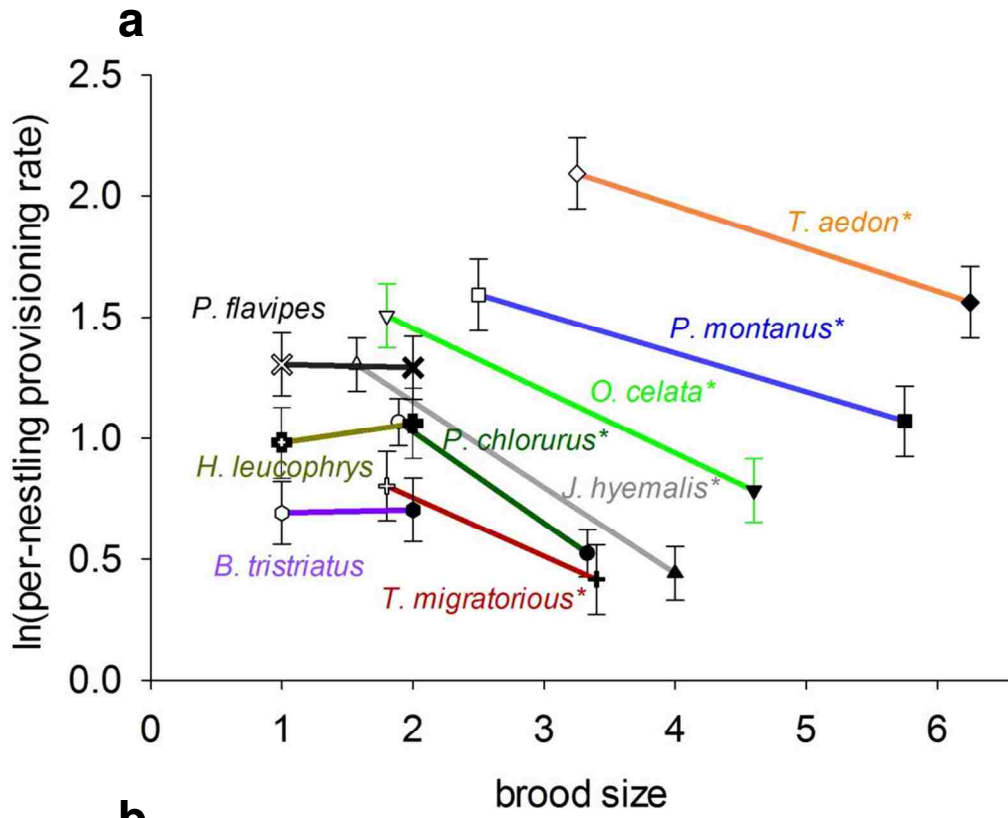


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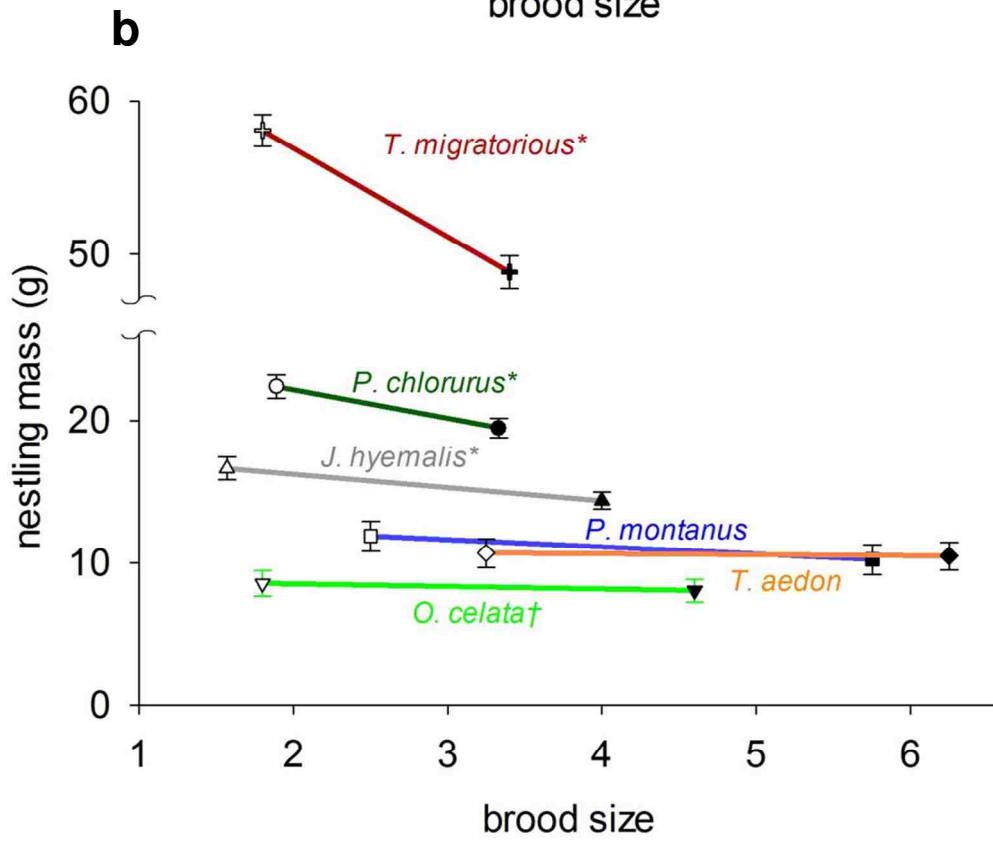


1196

1197 Figure 4.



1198



1199

1200 Figure 5.

1201 **Chapter 4: Dispersal and spatiotemporal variance in survival in three passerine bird**  
1202 **species with differing migratory behavior**

1203

1204 Daniel C. Barton and Thomas E. Martin

1205

1206 **Abstract**

1207 Variation among species in dispersal movements and seasonal migration are widely predicted to  
1208 hold major consequences for variation among species in population dynamics. Spatiotemporal  
1209 variance in survival probability is an important determinant of population growth rate in  
1210 iteroparous organisms. Spatial variance in survival is generally expected to decrease with  
1211 increasing dispersal, while seasonal migration is generally predicted to decrease temporal  
1212 variance in survival, yet tests of these predictions are lacking. Further, seasonal migration may  
1213 increase propensity for dispersal, yet tests of this prediction are also generally lacking. Using a  
1214 long-term (1993-2008) study of marked birds in Arizona and a multi-state mark-recapture  
1215 modeling framework, we describe variation in breeding dispersal movements among three  
1216 species of songbirds with differing migratory strategies. We further tested whether spatial  
1217 variance in survival decreased with increasing dispersal movements, and whether temporal  
1218 variance was lower in more migratory species, as predicted by general theory. We show that  
1219 dispersal movements were greatest in a long-distance migrant, moderate in a short-distance  
1220 migrant, and least in a resident songbird species. We find reduced spatial variance in survival  
1221 rates in species with greater dispersal movements. We find more complex differences in  
1222 temporal variation in survival rates among the three species, discuss potential explanations for  
1223 such differences, and compare with a larger published dataset of 19 additional species with

1224 differing migratory strategies. Our results suggest that spatial variance in survival rates of  
1225 songbirds is, as predicted by theory, partly mediated by breeding dispersal, while temporal  
1226 variance in survival rates is not lower in migratory species, contrary to expectations.

1227

1228 **Key words:** dispersal, spatial variance, temporal variance, survival, migration

1229 **Introduction**

1230 Dispersal is a key biological phenomenon with diverse causes and important consequences for  
1231 ecological and evolutionary processes (Clobert et al. 2001). Dispersal distances vary widely  
1232 among individuals and populations (Bowler and Benton 2005) with widely recognized  
1233 consequences for processes such as population dynamics (Kareiva et al. 1990, Daniels and  
1234 Walters 2000) and gene flow (Bohonak 1999). Dispersal may influence the degree of  
1235 spatiotemporal variance in survival rates (Pienkowski and Evans 1985) in addition to causing  
1236 biased underestimates of survival rates through permanent emigration in open populations (e.g.  
1237 Cilimburg et al. 2002). Survival rates and their variances are important deterministic  
1238 components of population growth rates (Tuljapurkar 1982, Saether and Bakke 2000, Clark and  
1239 Martin 2007) and are thus critical in the study of population ecology (Kareiva et al. 1990). Yet,  
1240 the effects of differences among species in dispersal distances on spatiotemporal variance in  
1241 survival rates or on bias in survival rate estimates remain generally unknown.

1242         The degree of spatiotemporal variance in population growth rate has important  
1243 consequences for long-term population growth rate (Tuljapurkar 1982) and probability of  
1244 population persistence in declining populations (Morris and Doak 2002). Increased  
1245 spatiotemporal variance in survival rates generally decreases long-term population growth rates  
1246 and reduces probability of population persistence, particularly in iteroparous organisms (Gaillard  
1247 et al. 2000, Sibly and Hone 2002). The two components of spatiotemporal variance, spatial  
1248 variance and temporal variance, are thought determined by alternative mechanisms. Spatial  
1249 variance in demographic parameters is expected to be low when dispersal is relatively high  
1250 (effectively creating a single population) or when a high degree of environmental correlation  
1251 exists across space (causing subpopulations to have similar rates; Harrison 1991). Conversely,



1252 spatial variance is expected to be high when dispersal is low or when a low degree of  
1253 environmental correlation exists across space. Temporal variance in survival rates is thought to  
1254 be greater in populations that experience more variable climatic conditions during an annual  
1255 cycle, and in the case of birds, this is represented by species that remain resident at high latitude  
1256 compared with long-distance migrants (Greenberg 1980, Mönkkönen 1992). Yet, tests of these  
1257 predictions are generally lacking.

1258         Dispersal also causes biased underestimates of survival in open populations, because  
1259 permanent emigration is not separable from mortality (Lebreton et al. 1992). Increased  
1260 permanent emigration from finite study areas reduces apparent survival while true survival  
1261 remains unchanged and unknown (Cilimburg et al. 2002; Zimmerman et al. 2007). Extending  
1262 this prediction to a comparative context, populations with reduced dispersal movements would  
1263 likely show reduced differences between apparent survival at a smaller spatial scale and a larger  
1264 spatial scale. Conversely, populations with greater dispersal movements would likely show  
1265 larger differences between small and large spatial scale estimates (Zimmerman et al. 2007).  
1266 While permanent emigration is widely recognized to create biased underestimates of survival in  
1267 open populations, the degree to which variation among populations in dispersal impacts the size  
1268 of this bias remains mostly unexamined (Marshall et al. 2004).

1269         We wished to test the influence of variation in dispersal behavior on spatial and temporal  
1270 variation in survival. *We a priori* selected for comparison three bird species that differed in their  
1271 migratory behavior because migratory strategy is a potential predictor of among-species  
1272 differences in dispersal behavior (Paradis et al. 1998). Bird species vary dramatically in their  
1273 seasonal migration strategies, ranging from the longest-distance movements known in animals to  
1274 resident species with limited home ranges (Alerstam 2001). Migrants generally have greater

1275 mobility than resident bird species and experience an increased probability of chance events  
1276 during migration that may influence dispersal movement (Alerstam 2001). Seasonal migrant  
1277 populations are therefore generally predicted to show greater dispersal movements than resident  
1278 populations, although tests are rare and have somewhat confounded migration with dispersal  
1279 (Paradis et al. 1998, Dawideit et al. 2009). Understanding variation among bird species in  
1280 dispersal movements and whether this variation is related to migratory strategy thus also remains  
1281 an open question. Nonetheless, migratory strategy was a reasonable basis of choosing species to  
1282 compare dispersal movements and consequences for spatiotemporal variation in survival rates.

1283         We conducted a 16-year mark-recapture-resight study of three passerine bird species with  
1284 differing migratory strategies. We tested four predicted effects of differences among species in  
1285 migratory and dispersal behavior for population dynamics and survival estimates. First, we  
1286 tested our assumption, suggested by a previous study (i.e. Paradis et al. 1998), that breeding  
1287 dispersal increased with migratory distance. We improved upon previous designs for testing the  
1288 relationship between dispersal movements and migratory behavior (Paradis et al. 1998, Dawideit  
1289 et al. 2009) by using the same methodology for comparisons across species, and by using a live  
1290 encounters study of wild populations rather than a band-recovery study which may confound  
1291 dispersal and migration. Second, we tested whether increased rates of breeding dispersal  
1292 reduced spatial variance in survival rates. Third, we tested whether increased migratory distance  
1293 reduced temporal variance in survival rates. We also compared our estimates of temporal  
1294 variance in survival rates with a published dataset of temporal variance in survival rates to test  
1295 whether migratory behavior predicted temporal variance in survival rates in a broader sample of  
1296 species. Fourth, we tested whether variation among species in dispersal movements biased

1297 survival estimates differentially by testing whether the difference between small and large spatial  
1298 scale estimates of survival increased with increasing dispersal movements.

## 1299 **Methods**

### 1300 *Study Site and Species*

1301 We conducted a long-term study of marked individual birds in Coconino County,  
1302 Arizona, USA from 1993-2008. Details on the location, plant community, and breeding bird  
1303 community of this study site have been detailed previously (Martin 1998, 2001). We marked,  
1304 recaptured, and resighted Mountain Chickadee (*Poecile gambelli*), Gray-headed Junco (*Junco*  
1305 *hyemalis dorsalis*), and Orange-crowned Warbler (*Oreothylpis celata orestera*). Mountain  
1306 Chickadee is a year-round resident or altitudinal migrant (Dixon and Gilbert 1964). Gray-headed  
1307 Junco is a short-distance migrant (Nolan et al. 1995). Orange-crowned Warbler is a neotropical  
1308 migrant that winters from northern to southern Mexico (Sogge et al. 2010). All three species are  
1309 songbirds (Order Passeriformes) of similar body size (approximately 10-20 g) that are territorial  
1310 and socially monogamous during the breeding season.

1311 Our study area was sub-divided into five strata, each containing between two and six  
1312 study plots (20 study plots total). Each stratum was a set of plots physically sub-divided from  
1313 other strata by either large canyons or a ridge and secondary road. The centroids of strata varied  
1314 from 0.9-10.7 km in distance from each other, and the five strata varied from 47-107 ha in area.  
1315 Study plots were individual snow-melt drainages of mixed coniferous and deciduous vegetation  
1316 (Martin 1998, 2001). We visited each plot to capture and mark birds from 6-10 times per season  
1317 during May-July of each year. We captured birds using mist-nets and individually marked them  
1318 using unique combinations of colored plastic leg bands and a single unique U.S. Fish and  
1319 Wildlife Service numbered aluminum band. Mist-netting effort was kept relatively constant

1320 across space and time from 1993-2008, using constant-effort mist-netting largely following the  
1321 Monitoring Avian Productivity and Survival (MAPS) protocol (DeSante et al. 1995), but also  
1322 allowing shifts of net positions among years. In particular, each plot was sampled by 10-20 mist-  
1323 nets, depending on plot size, for 6 hours per visit in each of the 6-10 visits per year. We also  
1324 included targeted trapping of birds near nests located for other research purposes at the same  
1325 study area. We concurrently conducted standardized nest-searching and monitoring at the same  
1326 study area. This afforded an opportunity to resight individuals by one or more observers  
1327 approximately every other day during May-July in every year.

1328         We created individual encounter histories representing the years and strata in which  
1329 individuals were encountered from our mist-netting and resighting data. While some individuals  
1330 included in the study were originally marked as juveniles, our encounter histories only included  
1331 data collected from the age of first breeding on, and thus our dispersal and survival estimates are  
1332 limited to breeding adults. It was occasionally impossible to assign the identity of a resighted  
1333 marked individual with certainty due to imperfect reading of band combinations in the field, and  
1334 such observations were excluded. We split encounter histories into groups for analysis by  
1335 species and sex, except for Mountain Chickadee, which could not always be sexed reliably using  
1336 in-hand criteria during parts of the breeding season. Our design thus consisted of a multi-state  
1337 mark-recapture model with five spatial strata and five species-sex groups (i.e. Orange-crowned  
1338 Warbler males, Orange-crowned Warbler females, Gray-headed Junco males, Gray-headed  
1339 Junco females, and Mountain Chickadee).

#### 1340 *Estimating Breeding Dispersal*

1341         We used multistate mark-recapture models (White et al. 2006) to estimate annual  
1342 probabilities of survival ( $S$ ), breeding dispersal characterized as transition probability between

1343 strata ( $\psi$ ), and encounter probability ( $p$ , the probability of encountering animals known to be  
1344 alive and present in the study area). We were specifically interested in testing whether  $\psi$  varied  
1345 among our five species-sex groups and thus developed our model set with this goal in mind. We  
1346 modeled  $S$  and  $p$  as functions of strata and group, and modeled  $\psi$  as a function of group and a  
1347 linear function of distance between strata. Our simplest model was thus one where  $S$ ,  $\psi$ , and  $p$   
1348 were constant among strata and groups ( $K$  or number of parameters = 3), while our most general  
1349 model contained interacting effects of strata and species and a within-species additive effect of  
1350 sex on  $S$  and  $p$ , and interacting effects of linear distance between strata and species and a within-  
1351 species additive effect of sex on  $\psi$  ( $K = 42$ ). We were unable to evaluate temporal and spatial  
1352 variance in  $S$ ,  $\psi$ , and  $p$  simultaneously due to a large number of inestimable parameters in a fully  
1353 time-varying model.

1354         We used an information-theoretic approach to compare the relative support for alternative  
1355 models because we were interested in comparing the explanatory value of non-nested models  
1356 and specifically in comparing alternative models of  $\psi$  among groups (Burnham and Anderson  
1357 2002). Parameters, likelihoods, and a sample size- and overdispersion-corrected derivation of  
1358 Akaike's Information Criteria (QAICc) were estimated for alternative models using Program  
1359 MARK (White and Burnham 1999). We estimated the overdispersion parameter,  $\hat{c}$  (the ratio of  
1360 observed variance to predicted model variance), using the median  $\hat{c}$  approach implemented in  
1361 Program MARK. Because multistate models frequently exhibit likelihood functions with  
1362 multiple maxima, we further used Markov-chain Monte Carlo (MCMC) estimation in Program  
1363 MARK to generate posterior probability distributions of parameter estimates and evaluate  
1364 whether our models had converged upon global maximum-likelihood (ML) estimates. We then

1365 used QAICc to compare relative support for alternative models in the context of our biological  
1366 question of interest, whether  $\psi$  varied among species-sex groups.

1367         Multistate mark-recapture models as implemented in Program MARK make two  
1368 assumptions germane to our study. First, our analysis assumes that survival from time  $i$  to time  
1369  $i+1$  does not depend upon the strata occupied by an individual in time  $i+1$ , and that mortality  
1370 occurs before movement. Second, our analysis assumes that individuals move among strata at  
1371 the same time, which in this case is between breeding seasons. We were unable to evaluate the  
1372 first assumption using our design, while the second assumption appears unlikely to be violated  
1373 because our marking and resighting was limited to the breeding season (May-July). Almost all  
1374 movement among strata apparently occurred outside of the sampling period because observed  
1375 movements among strata within a breeding season were very rare.

#### 1376 *Dispersal and Apparent Survival Differences Between Small and Large Spatial Scales*

1377 Our design involved an open population and a finite study area, and thus permanent emigration  
1378 from the entire study area was not separable from mortality (Lebreton et al. 1992). However, we  
1379 were able to compare amongst groups the extent to which differential dispersal may bias  
1380 estimates of survival by comparing estimated survival at a smaller spatial scale (a single stratum)  
1381 with survival at a larger spatial scale (the entire study area; Marshall et al. 2004). If we define  $S^r$   
1382 as the apparent survival probability of individuals in stratum  $r$  and  $\psi^{rr}$  as the probability of an  
1383 individual transitioning from stratum  $r$  to stratum  $r$ , apparent survival ( $\phi$ ) within stratum  $r$ ,  $\phi^r$ , is  
1384 the product  $S^r \psi^{rr}$ . The parameter  $\phi^r$  thus excludes individuals that transition to another strata. In  
1385 multistate models where strata are spatially based, such as ours,  $\phi^r$  is the apparent survival at a  
1386 smaller spatial scale than  $S^r$ . We will refer to this difference as  $S^r - \phi^r$ . We quantified the  
1387 magnitude of that difference to test the degree to which dispersal may differentially bias

1388 estimates of apparent survival in species with increased dispersal. We used estimates of  $S$  and  $\psi$   
1389 from the global multistate mark-recapture (described above;  $S_{sp(sex) \cdot st}$   $P_{sp(sex) \cdot strata}$   $\Psi_{sp(sex) \cdot dist}$ ) to  
1390 calculate the difference between  $\phi^r$  and  $S^r$  because this model allowed  $S$  and  $\psi$  to vary freely  
1391 among strata. We calculated  $S^r - \phi^r$  for each stratum and species-sex grouping and then  
1392 calculated the arithmetic mean of  $S^r - \phi^r$  within each species-sex grouping for comparison.

### 1393 *Estimating Spatial and Temporal Variance in Survival Rates*

1394 We estimated spatial and temporal variance in annual survival rates using a Bayesian  
1395 hierarchical modeling approach implemented in Program MARK version 6.1 (White et al. 2009)  
1396 that separates process and sampling variance in estimates of survival rates (Gould and Nichols  
1397 1988; Lukacs et al. 2008). This modeling approach assumes that the logit-transformed survival  
1398 rate of each group  $g$  in stratum  $i$  or year  $t$  is a realization of a normally distributed random  
1399 variable (a ‘hyperdistribution’) with mean  $\mu$  and standard deviation  $\sigma$ . We estimated the  
1400 posterior probability distribution of parameters  $\mu$  and  $\sigma$  using a Metropolis-Hastings Markov  
1401 chain Monte Carlo algorithm in Program MARK. We used uninformative prior distributions for  
1402  $\mu$  (Normal (0,100)) and  $1/\sigma^2$  (Gamma(0.001,0.001)). We determined the number of tuning, burn-  
1403 in, and posterior sampling iterations based upon inspection of preliminary chains of varying  
1404 length for signs of lack of convergence to an equilibrium distribution (McCarthy 2007). We then  
1405 visually examined the posterior probability distribution for signs of non-convergence and used  
1406 the Gelman-Rubin statistic to quantitatively assess whether independent chains with alternative  
1407 starting values converged upon similar posterior distributions (Gelman and Rubin 1992). To  
1408 estimate the posterior probability distribution of the spatial variance in survival rates, we  
1409 generated 25,000 samples of the posterior probability distribution after 5,000 ‘tuning’ samples  
1410 and 10,000 ‘burn-in’ samples following inspection of multiple preliminary chains. To estimate

1411 the posterior probability distribution of temporal variance in survival rates, we generated 250,000  
1412 samples of the posterior probability distribution after 25,000 ‘tuning’ samples and 25,000 ‘burn-  
1413 in’ samples.

1414 We estimated spatial and temporal variance in survival separately due to the large  
1415 number of inestimable parameters in a fully time-dependent multistate model. To estimate  
1416 spatial variance in survival rates, we used a multi-state mark-recapture model where  $S$  and  $p$   
1417 varied among species and strata with an additive effect of sex, and where  $\psi$  varied as a function  
1418 of the interaction between species-sex group and linear distance between strata. Thus, logit  $S$  of  
1419 strata  $i$  ( $N = 5$  strata) was modeled as normally distributed with mean  $\mu_S$  and variance  $\sigma_S$ , and the  
1420 distribution of logit  $p$  of strata  $i$  was modeled in the same fashion as  $\mu_p$  and variance  $\sigma_p$ . To  
1421 estimate temporal variance in survival rates, we combined all strata and used a time-dependent  
1422 Cormack-Jolly-Seber model where  $\phi$  and  $p$  varied among species with an additive effect of sex.  
1423 Thus, logit  $\phi$  and  $p$  of interval  $t$  ( $N = 15$  intervals) were modeled as normally distributed with  
1424 mean  $\mu_\phi$  or  $\mu_p$  and variance  $\sigma_\phi$  or  $\sigma_p$ . Instead of explicitly testing whether parameters vary among  
1425 groups or strata (see Estimating Breeding Dispersal, above) this approach assumes that there is  
1426 variance and focuses on estimating the magnitude of spatial or temporal process variance (the  
1427 hyperdistribution parameter  $\sigma$ ). We report the posterior probability distributions of  $\mu$  and  $\sigma$  of  
1428 these hyperdistributions and compare them among species-sex groups to test the prediction that  
1429 spatial variance and temporal variance decrease with increasing dispersal movement and  
1430 migratory distance.

## 1431 **Results**

### 1432 *Breeding Dispersal*



1433 We marked 604 adult Orange-crowned Warbler (291 male, 313 female), 1066 adult  
1434 Gray-headed Junco (539 male, 527 female), and 455 adult Mountain Chickadee during the  
1435 course of the study, and this sample of individuals was used in all reported analyses. Multi-state  
1436 mark-recapture models that included differences among species and sex in breeding dispersal  
1437 probability,  $\psi$ , were strongly supported by QAICc (Table 1). Comparisons of QAICc among  
1438 models showed model selection uncertainty among models including an effect of species and sex  
1439 on the intercept and slope of the relationship between  $\psi$  and distance ( $\psi_{\text{species}\cdot\text{sex}\cdot\text{distance}}$  and  
1440  $\psi_{\text{species}(\text{sex})\cdot\text{distance}}$ ), a model including only an effect of species and sex on the intercept  
1441 ( $\psi_{\text{species}(\text{sex})+\text{distance}}$ ), and a model including an effect of species but not sex on the intercept  
1442 ( $\psi_{\text{species}+\text{distance}}$ ). Median  $\hat{c}$  was estimated from the global model as 1.186, which suggested  
1443 appropriate model fit and a lack of significant overdispersion. We graphically present model-  
1444 averaged estimates from these top 4 models, which represented 98.1% of the QAICc weight  
1445 (Figures 1 and 2). Estimates of the intercept of the  $\psi$ -distance function, averaged across sexes,  
1446 were greatest for Orange-crowned Warbler and least for Mountain Chickadee, with Gray-headed  
1447 Junco intermediate, and slopes differed among species (Figure 1). Estimates of the intercept of  
1448 the  $\psi$ -distance function were greater in females than males in both Orange-crowned Warbler and  
1449 Gray-headed Junco (Figure 2).

#### 1450 *Apparent Survival Differences Between Small and Large Spatial Scales*

1451 Survival in stratum  $r$  ( $S^r$ ) is greater than stratum-specific survival ignoring individuals  
1452 that moved to other strata ( $\phi^r$ ) if dispersal among strata ( $\psi^{rx}$ ) is greater than 0 by definition, and  
1453 we estimated  $\psi^{rx} > 0$  for all three species in this study (Figures 1, 2). This difference,  $S^r - \phi^r$ ,  
1454 thus varied among species-sex groups as a result of variation in  $\psi^{rx}$  among species-sex groups.  
1455 Orange-crowned Warblers showed the largest values of  $S^r - \phi^r$  (males: average 0.040, range

1456 0.038-0.042; females: average 0.108, range 0.102-0.111), with Gray-headed Juncos intermediate  
1457 (males: average 0.038, range 0.036-0.045; females: average 0.067, range 0.063-0.079) and  
1458 Mountain Chickadees the least (average 0.023, range 0.018-0.030).

1459 *Spatiotemporal Variance in Survival Rates*

1460           Posterior probability distributions of mean ( $\mu$ ) and spatial process variance ( $\sigma$ ) of strata  
1461 survival rates ( $S$ ) from our Bayesian multistate modeling approach differed by sex and species  
1462 (Table 2). Mean survival rate ( $S$ ) was greater in males than females in Orange-crowned Warbler  
1463 and Gray-headed Junco, although the difference between sexes was greater in the warbler than  
1464 the junco (Table 2). Mean survival rate, averaged across sexes, was relatively greatest in  
1465 Orange-crowned Warbler, similar in Gray-headed Junco, and least in Mountain Chickadee  
1466 (Table 2). Spatial process variance did not vary in the same order; it was relatively greatest in  
1467 Mountain Chickadee, moderate in Orange-crowned Warbler, and least in Gray-headed Junco, but  
1468 the magnitude of these differences was small (Table 2).

1469           Posterior probability distributions of mean ( $\mu$ ) and temporal process variance ( $\sigma$ ) of  
1470 annual survival rates ( $\phi$ ) from our Cormack-Jolly-Seber modeling approach differed by sex and  
1471 species (Table 3). Mean survival rate ( $\phi$ ) was greater in male than female Orange-crowned  
1472 Warblers and Gray-headed Juncos (Table 3). Mean survival rate, averaged across sexes, was  
1473 greatest in Orange-crowned Warbler, moderate in Gray-headed Junco, and least in Mountain  
1474 Chickadee (Table 3). Temporal variance, averaged by species, was greatest in Mountain  
1475 Chickadee, moderate in Orange-crowned Warbler, and least in Gray-headed Junco (Table 3).  
1476 Temporal process variance was marginally greater in female than male Orange-crowned  
1477 Warblers (Table 3).

1478

1479 **Discussion**

1480           The suggestion that more highly migratory species are generally more mobile and thus  
1481 may be more dispersal-prone (Paradis et al. 1998, Alerstam 2001), which served as our initial  
1482 assumption for choosing species, was supported. As predicted, Orange-crowned Warbler, the  
1483 most highly migratory of the three species examined, showed the greatest probability of breeding  
1484 dispersal, while Mountain Chickadee, the year-round resident, showed the lowest probability of  
1485 dispersal, with Gray-headed Junco intermediate (Figure 1). Females showed greater probability  
1486 of breeding dispersal in both Orange-crowned Warbler and Gray-headed Junco, and indeed,  
1487 differences among species in dispersal were partly driven by higher probabilities of breeding  
1488 dispersal in females (Figure 2). Female-biased breeding dispersal is already widely recognized  
1489 as the dominant pattern in birds (Greenwood 1980, Greenwood and Harvey 1982, Clobert et al.  
1490 2001; but see Öst et al. 2011). Given our small sample size of three species, we do not make a  
1491 strong case for a causal connection between migratory behavior and dispersal, except to note that  
1492 the correlation exists, and variation among species in dispersal allowed us to test the effects of  
1493 interspecific variation in dispersal on spatiotemporal variance in survival.

1494           We explored the extent to which the observed differences among species in dispersal ( $\psi$ )  
1495 affected the differences between apparent survival estimates of each strata including individuals  
1496 that dispersed to other strata ( $S^r$ ) and strata-specific rates excluding such dispersal ( $\phi^r$ ). An  
1497 increase in the difference  $S^r - \phi^r$  with increasing  $\psi$  is unremarkable because it is expected by  
1498 definition. However, the magnitude of the differences in  $S^r - \phi^r$  is indicative of the relative  
1499 degree of bias in  $S$ . We found that average  $S^r - \phi^r$  was small in Mountain Chickadee (0.023),  
1500 which showed the lowest dispersal movements (Fig. 1) while average  $S^r - \phi^r$  was relatively large  
1501 in Orange-crowned Warbler females (0.108), which showed the greatest dispersal (Figs. 1, 2).

1502 Thus, we suggest that the difference between true survival and apparent survival may diverge  
1503 quite rapidly with increasing dispersal movements, particularly in migrant species. Our strata  
1504 were all > 47 ha in size, yet we still observed somewhat large  $S^r - \phi^r$  in female Orange-crowned  
1505 Warbler and Gray-headed Junco. This further confirms small study areas underestimate survival  
1506 in migrant species that display high breeding dispersal (Cilimburg et al. 2002), and that this bias  
1507 may be larger in females (Marshall et al. 2004) which generally show greater dispersal in birds  
1508 (Greenwood 1980). Our results further show how the relative degree of this bias may vary  
1509 among species with alternative dispersal behavior.

1510         Variation in dispersal movement can affect the way populations respond to a dynamic  
1511 environment, through the effects of dispersal on spatial variance in demographic parameters  
1512 including population size, reproduction, and survival rates (Clobert et al. 2001). We found small  
1513 differences among species in the spatial variance of survival rates (Table 2), which is jointly  
1514 determined by dispersal movements and environmental correlation across space (Morris and  
1515 Doak 2002). In interpreting our results, we make the explicit assumption that the degree of  
1516 environmental correlation is very high and similar for the three species examined, which is not  
1517 unreasonable as the three species were examined at the same sites in the same years in a small  
1518 geographic area (<12 km at its widest point). Thus, differences observed among species in  
1519 spatial variance are caused by differences in dispersal movement. In the case of the species  
1520 studied here, spatial variance in survival rates decreased with increasing breeding dispersal, as  
1521 predicted by general metapopulation theory (Bowler and Benton 2005). The observed  
1522 differences in spatial variance of survival rates were small relative to mean survival rates.  
1523 Nonetheless, such differences in the variance of demographic parameters may have large effects

1524 on population growth rates (e.g. Schorcht et al. 2009) depending upon the elasticity of the  
1525 demographic parameter in question (Saether and Bakke 2000, Clark and Martin 2007).

1526         We found fairly large temporal variation in survival rates, as well as large differences  
1527 among species in temporal variation in survival rates (Table 3). Temporal process variance in  
1528 demographic parameters is theoretically predicted to decrease with increasing parameter  
1529 elasticity because natural selection is thought to canalize variance in traits with large effects on  
1530 fitness (Pfister 1998; Gaillard et al. 2000). This prediction has been increasingly supported  
1531 across a broad range of taxa including birds (Schmutz 2009), but temporal variance is also  
1532 predicted to increase with increasing environmental variability (Greenberg 1980, Monkkonen  
1533 1992). Species resident at higher latitudes in temperate regions are thought to be exposed to  
1534 higher degrees of environmental variability than migratory species that retreat to more equatorial  
1535 latitudes during the winter, and thus the high-latitude residents may show greater temporal  
1536 variation in survival rates. We found, instead, that Orange-crowned Warbler, the longest-  
1537 distance migrant, showed the greatest temporal variance, with the high-latitude resident  
1538 Mountain Chickadee intermediate, and Gray-headed Junco showing the least temporal variance.  
1539 Temporal variance in survival rates can be caused by numerous stochastic and deterministic  
1540 ecological processes, such as random weather events (Jonzen et al. 2002) or deterministic  
1541 density-dependence in survival (Ekman 1984). In the case of migratory species, stochastic  
1542 weather-driven processes occurring on the wintering grounds or in stopover habitat may have  
1543 significant effects on demographic parameters including survival rates (Sillett et al. 2000, Sillett  
1544 and Holmes 2002). Further, our study measured apparent survival, as is the case with all studies  
1545 of open marked populations, and thus temporal variance in apparent survival could also reflect

1546 temporally variable long-distance dispersal not measurable by our study. We are unable to test  
1547 this alternative using our dataset of three species.

1548         We compared our results *post-hoc* with 19 previously published studies of songbird  
1549 (Order Passeriformes) temporal survival variance reviewed by Schmutz (2009). We tested  
1550 whether temporal variance in survival rates differed between 14 migrant and 8 resident species  
1551 and found no difference (Figure 3). Thus, while migratory strategy may be predicted to play a  
1552 role in determining the degree of temporal variation in survival rates, we are unable to support  
1553 this prediction with either our results or by combining our results with previously published  
1554 results.

1555         We tested three predictions regarding the relationship between dispersal, migratory  
1556 strategy, and population dynamics. Given the importance of dispersal behavior in numerous  
1557 ecological and evolutionary processes, including population responses to anthropogenic  
1558 disturbances such as climate change (Thomas et al. 2004) and fragmentation (Cushman 2006),  
1559 our results suggest that population response to disturbances may depend on migratory strategy  
1560 assuming that dispersal increases with migratory distance (Paradis et al. 1998, Fig. 1). Further,  
1561 increasing dispersal appeared to decrease spatial variance in survival rates among species.  
1562 Spatial variance in survival rates is an important component in many metapopulation models  
1563 (Harrison 1991; Morris and Doak 2002), and this again suggest that intrinsic differences among  
1564 populations in migratory behavior may strongly impact population biology.

1565         We found that temporal variance in survival rates differed among species, yet was not  
1566 predicted by differences among species in migratory behavior. Examining a larger dataset of  
1567 temporal variance estimates of songbirds assembled by Schmutz (2009) suggested that temporal  
1568 variance in apparent survival rates is not predicted by differences among populations in

1569 migratory behavior. Overall, our results thus suggest that differences among populations in  
1570 dispersal behavior may be linked to migratory behavior and strongly impact key processes such  
1571 as spatial variance in survival rates, yet may not be linked to temporal variance in survival rates  
1572 as previously proposed. Dispersal and spatiotemporal variance in survival rates are critical  
1573 elements of population dynamics, and thus consideration of a species' migratory strategy may be  
1574 important in understanding and predicting population responses to natural and anthropogenic  
1575 disturbances.

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1712 sampling estimates of spotted owl survival and population trends. Journal of Applied  
1713 Ecology 44:963-971.

1714 Table 1. Model selection criteria for alternative multi-state mark-recapture models of the effects  
 1715 of species (sp), sex, and strata (st) on survival ( $S$ ) and recapture probability ( $p$ ) and of the effects  
 1716 of species, sex, and distance between strata (dist) on movement probability ( $\psi$ ). Models are  
 1717 shown in descending order by QAICc. The 10 models with lowest QAICc are shown with the  
 1718 most general global model and most reduced model for comparison (in italics). The top four  
 1719 models, representing 98.1% of QAICc weight, are in bold.

Model			QAICc <sup>1</sup>	$\Delta$ QAICc <sup>2</sup>	Weight <sup>3</sup>	$\ell^4$	$K^5$	Deviance
<b><math>S_{sp(sex)}</math></b>	<b><math>P_{sp(sex)+st}</math></b>	<b><math>\psi_{sp(sex)+dist}</math></b>	<b>3483.30</b>	<b>0</b>	<b>0.550</b>	<b>1</b>	<b>20</b>	<b>1883.5</b>
$S_{sp(sex)}$	$P_{sp(sex)+st}$	$\psi_{sp(sex)\cdot dist}$	3484.44	1.134	0.312	0.567	22	1880.5
$S_{sp(sex)}$	$P_{sp(sex)+st}$	$\psi_{sp\cdot sex\cdot dist}$	3487.50	4.20	0.067	0.123	24	1879.5
$S_{sp(sex)}$	$P_{sp(sex)+st}$	$\psi_{sp+dist}$	3487.99	4.70	0.053	0.096	18	1892.2
$S_{sp(sex)}$	$P_{sp(sex)\cdot strata}$	$\psi_{sp(sex)\cdot dist}$	3491.28	7.98	0.010	0.019	30	1870.9
$S_{sp(sex)+st}$	$P_{sp(sex)+st}$	$\psi_{sp(sex)\cdot dist}$	3491.79	8.49	0.008	0.014	26	1879.7
$S_{sp(sex)+st}$	$P_{sp(sex)\cdot strata}$	$\psi_{sp(sex)\cdot dist}$	3498.64	15.34	0	0.001	34	1870.0
$S_{sp(sex)}$	$P_{sp(sex)+st}$	$\psi_{sp\cdot dist}$	3498.80	15.50	0	0	19	1901.0
$S_{sp(sex)+st}$	$P_{sp(sex)}$	$\psi_{sp(sex)\cdot dist}$	3499.9	16.62	0	0	22	1896.0
$S_{sp(sex)}$	$P_{sp(sex)}$	$\psi_{sp(sex)+dist}$	3501.00	17.70	0	0	16	1909.3
<i><math>S_{sp(sex)\cdot st}</math></i>	<i><math>P_{sp(sex)\cdot strata}</math></i>	<i><math>\psi_{sp(sex)\cdot dist}</math></i>	<i>3510.86</i>	<i>27.56</i>	<i>0</i>	<i>0</i>	<i>42</i>	<i>1865.7</i>
<i><math>S.</math></i>	<i><math>p.</math></i>	<i><math>\psi.</math></i>	<i>3585.99</i>	<i>102.69</i>	<i>0</i>	<i>0</i>	<i>3</i>	<i>2020.6</i>

1720

1721 <sup>1</sup>QAICc =  $-2 \cdot \log \text{likelihood} / c\text{-hat} + 2K + 2K(k+1) / (n\text{-ess} - K - 1)$  where n-ess is effective sample  
 1722 size.

1723 <sup>2</sup>  $\Delta$ QAICc = QAICc of model – QAICc of lowest model

1724  $e^{(-1/2 * \Delta QAICc \text{ of model } i)} / \sum (-1/2 \Delta QAICc)$

1725 <sup>4</sup>Model likelihood

1726 <sup>5</sup>Number of parameters

1727



1728 Table 2. Posterior probability distribution mean and standard deviation (SD) for annual survival  
 1729 ( $\mu$ ) and spatial process variation in annual survival ( $\sigma$ ) for three species of birds, estimated from  
 1730 a hierarchical Bayes multi-state mark-recapture model. Orange-crowned Warbler (ocwa) and  
 1731 Gray-headed Junco (ghju) are shown split by sex and overall, while Mountain Chickadee (moch)  
 1732 was not separated by sex.

	$\mu$		$\sigma$	
	mean	SD	mean	SD
ocwa	0.5667	0.0360	0.0279	0.0247
ocwa m	0.6023	0.0353	0.0278	0.0245
ocwa f	0.5302	0.0364	0.0281	0.0246
ghju	0.5556	0.0300	0.0325	0.0317
ghju m	0.5657	0.0299	0.0317	0.0316
ghju f	0.5454	0.0302	0.0329	0.0316
moch	0.5106	0.0467	0.0440	0.0402

1733

1734 Table 3. Posterior probability distribution mean and standard deviation (SD) for annual survival  
 1735 ( $\mu$ ) and temporal process variation in annual survival ( $\sigma$ ) for three species of birds, estimated  
 1736 from a hierarchical Bayes Cormack-Jolly-Seber model. Orange-crowned Warbler (ocwa) and  
 1737 Gray-headed Junco (ghju) are shown split by sex and averaged, while Mountain Chickadee  
 1738 (moch) was not separated by sex.

	$\mu$		$\sigma$	
	mean	SD	mean	SD
ocwa	0.5945	0.0580	0.1728	0.0535
ocwa m	0.6406	0.0553	0.1650	0.0514
ocwa f	0.5467	0.0597	0.1778	0.0551
ghju	0.5589	0.0349	0.0580	0.0444
ghju m	0.5764	0.0345	0.0575	0.0439
ghju f	0.5411	0.0352	0.0585	0.0449
moch	0.5440	0.0682	0.1399	0.1011

1739

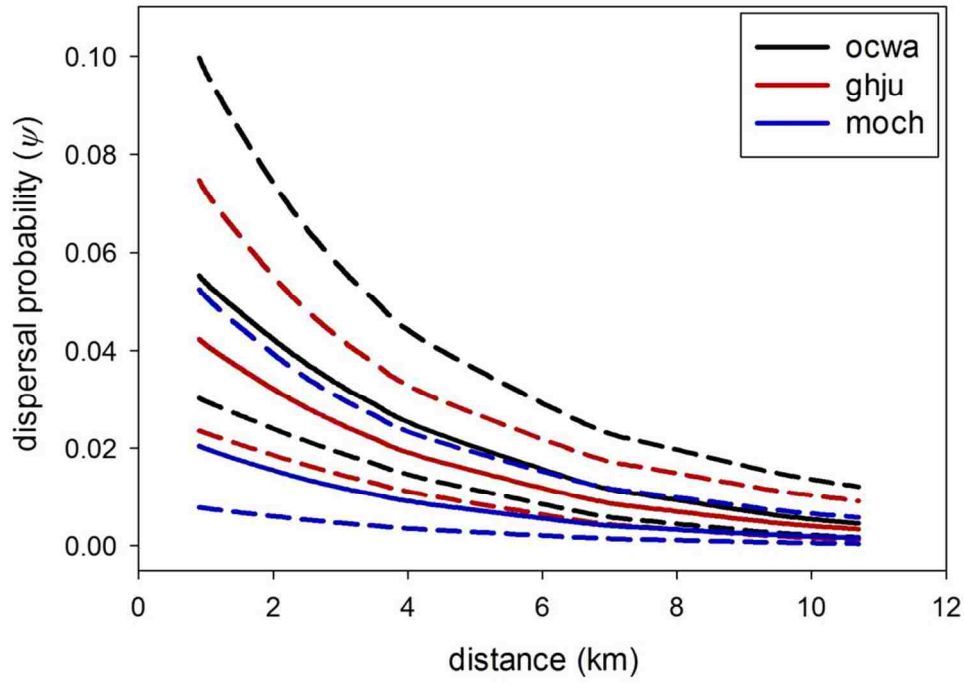
1740 Figure 1. Model-averaged maximum likelihood estimates and 95% confidence intervals of the  
1741 relationship between breeding dispersal probability ( $\psi$ ) and distance between strata for three  
1742 species of passerine bird: Orange-crowned Warbler (ocwa), Gray-headed Junco (ghju), and  
1743 Mountain Chickadee (moch).

1744

1745 Figure 2. Model-averaged maximum likelihood estimates and 95% confidence intervals of the  
1746 relationship between breeding dispersal probability ( $\psi$ ) and distance between strata by sex for  
1747 two species of passerine bird: Orange-crowned Warbler (ocwa, panel A), Gray-headed Junco  
1748 (ghju, panel B).

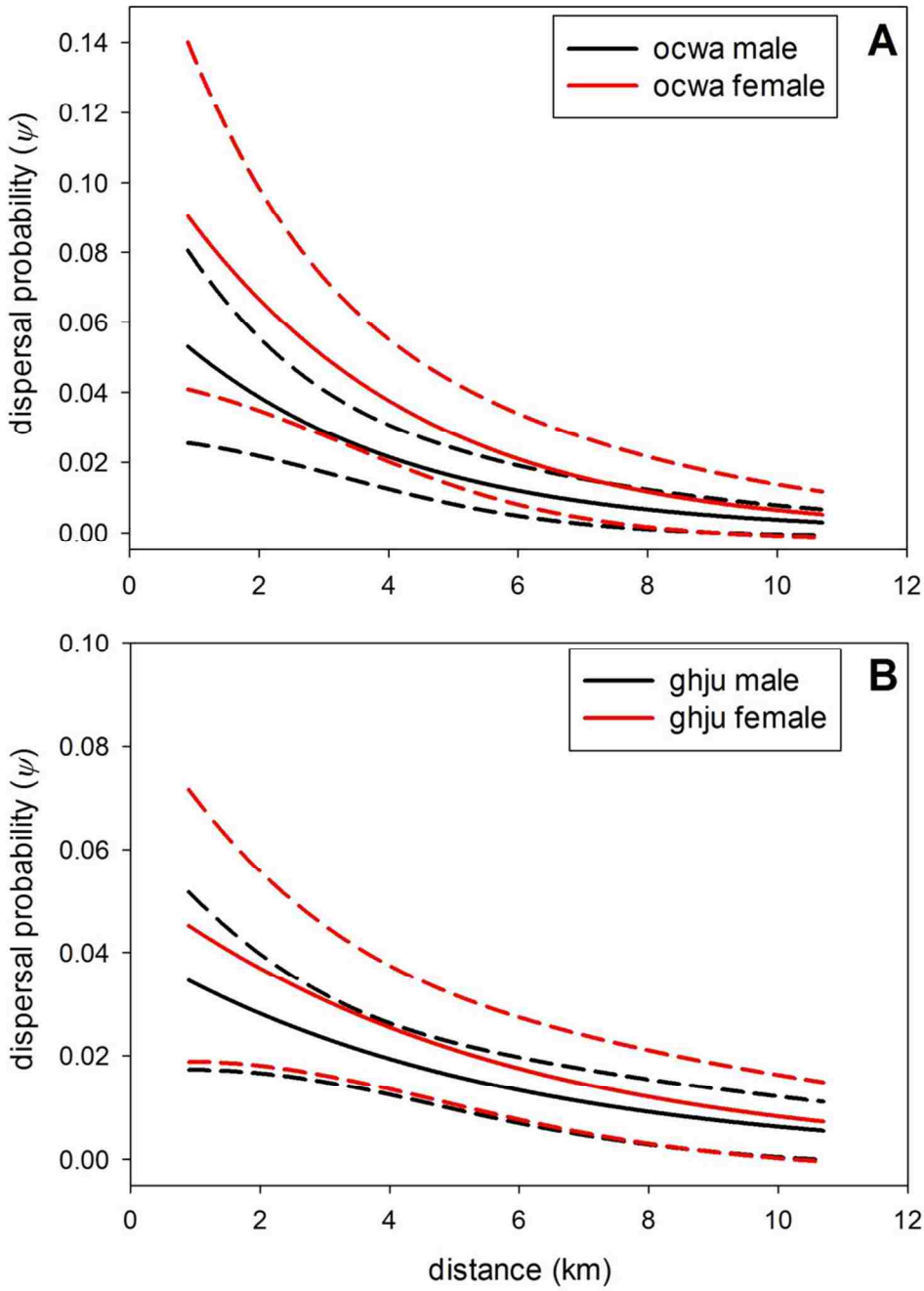
1749

1750 Figure 3. Comparison of estimates of the temporal process coefficient of variation (CV) in  
1751 annual survival rates of 22 populations of migrant and resident passerine birds. Sample size  
1752 (number of studies) is shown above each box. Data from this study (Table 3) and Schmutz  
1753 (2009).



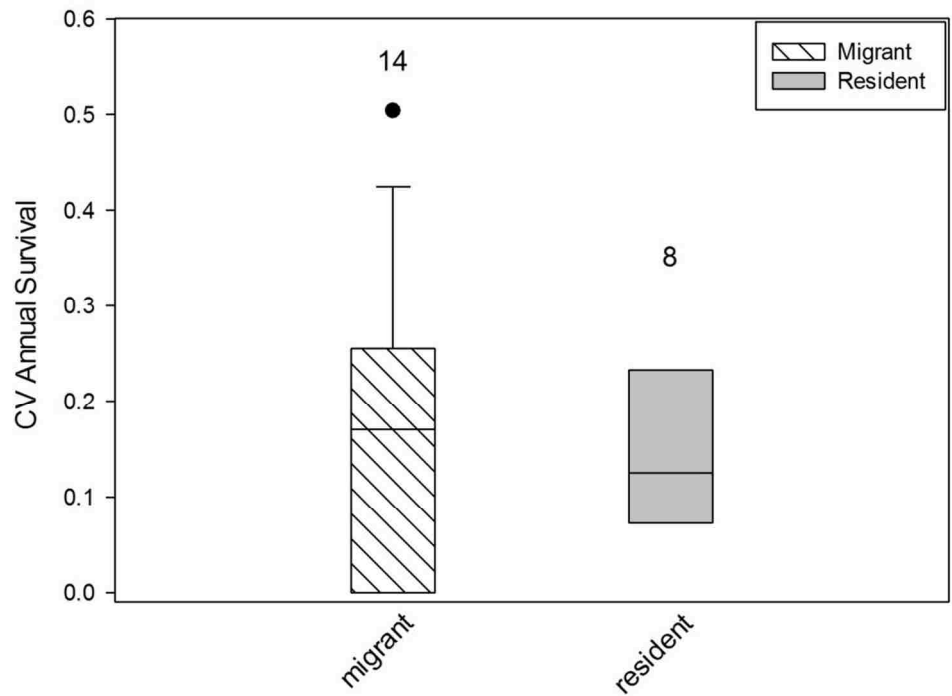
1754

1755 Figure 1.



1756

1757 Figure 2.



1758

1759 Figure 3.

1760 **Appendix 1.** Species, season of greatest mortality and data sources used in meta-analyses of  
 1761 seasonal mortality rates (Chapter 2, Tables 2 and 3).  
 1762

species	season of greatest mortality (Table 2)	study	monthly mortality rate analysis (Table 3)	method
<i>Tetrao tetrax</i>	summer	Angelstam 1984	y	T
<i>Melospiza melodia</i>	winter	Arcese et al. 1992	y	MR
<i>Rostrhamus sociabilis</i>	winter	Bennetts & Kitchen 1999	y	T
<i>Anas fulvigula</i>	winter	Bielefeld and Cox 2006	y	T
<i>Parus atricapillus</i>	winter	Brittingham & Temple 1988	y	MR
<i>Alectoris rufa</i>	summer	Buenestado et al. 2009		T
<i>Dryocopus pileatus</i>	equivocal confounded by	Bull 2001	y	T
<i>Colinus virginianus</i>	hunting	Burger et al. 1995 Caizergues & Ellison	y	T
<i>Tetrao tetrax</i>	summer	1997	y	T
<i>Protonotaria citrea</i>	equivocal*	Calvert et al. 2010		MR
<i>Seiurus noveboracensis</i>	equivocal*	Calvert et al. 2010		MR

<i>Branta hrota</i>	summer*	Clausen et al. 2001		MR
<i>Centrocercus</i>				
<i>urophasianus</i>	summer	Connelly et al. 2000		T
<i>Serinus serinus</i>	winter	Conroy et al. 2002		MR
	confounded by			
<i>Colinus virginianus</i>	hunting	Cox et al. 2004	y	T
	confounded by			
<i>Bonasa umbellus</i>	hunting	Devers et al. 2007	y	T
<i>Haematopus</i>				
<i>ostralegus</i>	summer	dit Durrell 2007	y	MR
<i>Parus montanus</i>	equivocal	Ekman & Askenmo 1986		MR
<i>Parus cristatus</i>	equivocal	Ekman & Askenmo 1986		MR
<i>Strix occidentalis</i>	winter	Ganey et al. 2005		T
<i>Chen caerulescens</i>				
<i>atlantica</i>	equivocal*	Gauthier et al. 2001		MR
<i>Tympanuchus</i>				
<i>pallidicinctus</i>	summer	Hagen et al. 2007	y	T
<i>Dendragapus</i>				
<i>canadensis franklinii</i>	summer	Herzog 1979	y	T
<i>Meleagris gallopavo</i>				
<i>silvestris</i>	summer	Hubbard et al. 1999	y	T
<i>Chen canagica</i>	summer*	Hupp et al. 2007		T
<i>Branta canadensis</i>	equivocal*	Hupp et al. 2010		T



<i>Dendroica cerulea</i>	equivocal*	Jones et al. 2004		MR
<i>Meleagris gallopavo</i>		Keegan & Crawford		
<i>intermedia</i>	equivocal	1999		T
<i>Meleagris gallopavo</i>				
<i>silvestris</i>	summer	Kurzejeski et al. 1987	y	T
<i>Parus montanus</i>	equivocal	Lahti et al. 1998		MR
<i>Meleagris gallopavo</i>				
<i>merriami</i>	summer	Lehman et al. 2005	y	T
<i>Anser</i>	confounded by			
<i>brachyrhynchus</i>	hunting*	Madsen et al. 2002		MR
<i>Petroica goodenovii</i>	summer	Major & Gowing 2001	y	MR
<i>Meleagris gallopavo</i>				
<i>silvestris</i>	summer	Miller et al. 1998	y	T
		Montadert & Leonard		
<i>Bonasa bonasia</i>	winter	2003	y	T
<i>Pyrrhula murina</i>	equivocal	Monticelli et al. 2009		MR
<i>Centrocercus</i>				
<i>urophasianus</i>	equivocal	Moynahan et al. 2006		T
<i>Meleagris gallopavo</i>				
<i>silvestris</i>	summer	Nguyen et al. 2003	y	T
<i>Sitta europea</i>	winter	Nilsson 1982	y	MR
<i>Somateria spectabilis</i>	equivocal	Oppel & Powell 2010		T
<i>Meleagris gallopavo</i>	confounded by	Pack et al. 1999		T

<i>silvestris</i>	hunting			
<i>Meleagris gallopavo</i>	summer	Palmer et al. 1993	y	T
<i>Motacilla clara</i>	summer	Piper 2002	y	MR
<i>Anser anser</i>	equivocal*	Pistorius et al. 2006		MR
<i>Meleagris gallopavo</i>				
<i>silvestris</i>	summer	Roberts et al. 1995	y	T
<i>Alectoris chukar</i>	winter	Robinson et al. 2009	y	T
<i>Turdus merula</i>	summer	Robinson et al. 2010	y	MR
		Sankamethawee et al.		
<i>Alophoixus pallidus</i>	summer	2011		MR
<i>Sylvia boehmi</i>	summer	Schaefer et al. 2006	y	MR
<i>Sylvia lugens</i>	winter	Schaefer et al. 2006	y	MR
<i>Anser albifrons</i>				
<i>frontalis</i>	equivocal*	Schmutz & Ely 1999		MR
<i>Centrocercus</i>				
<i>urophasianus</i>	summer	Sika 2006	y	T
<i>Dendroica</i>				
<i>caerulescens</i>	equivocal*	Sillett & Holmes 2002		MR
<i>Lichenostomus</i>				
<i>melanops</i>	equivocal	Smales et al. 2009		MR
	confounded by			
<i>Bonasa umbellus</i>	hunting	Small et al. 1993		T
<i>Bonasa umbellus</i>	summer	Small et al. 1993		MR

			Smith & Willebrand		
<i>Lagopus lagopus</i>	winter	1999	y	T	
<i>Parus atricapillus</i>	winter	Smith 1967	y	MR	
<i>Tetrao tetrix</i>	winter	Spidso et al. 1997		T	
<i>Meleagris gallopavo</i>					
<i>silvestris</i>	summer	Spohr et al. 2004	y	T	
<i>Strix aluco</i>	equivocal	Sunde et al. 2003		T	
<i>Colinus virginianus</i>	winter	Terhune et al. 2007	y	T	
		Thompson & Fritzell			
<i>Bonasa umbellus</i>	winter	1989	y	T	
<i>Meleagris gallopavo</i>		Vander Haegen et al.			
<i>silvestris</i>	summer	1988	y	T	
		Vangilder & Kurzejeski			
<i>Melegaris gallopavo</i>	summer	1995	y	T	
<i>Branta bernicula</i>					
<i>nigricans</i>	summer*	Ward et al. 1997		MR	
<i>Tympanuchus</i>					
<i>pallidicinctus</i>	summer	Wolfe et al. 2007	y	T	
<i>Tympanuchus</i>					
<i>pallidicinctus</i>	summer	Wolfe et al. 2007	y	T	
<i>Meleagris gallopavo</i>	summer	Wright et al. 1996	y	T	

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1763 \*Migrant species excluded from some analyses (see Methods and Results).

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2202 **Appendix 3.** Phylogenetic analyses of the relationships between ecological factors and reaction  
2203 norms of feeding rate to offspring number (Chapter 3).

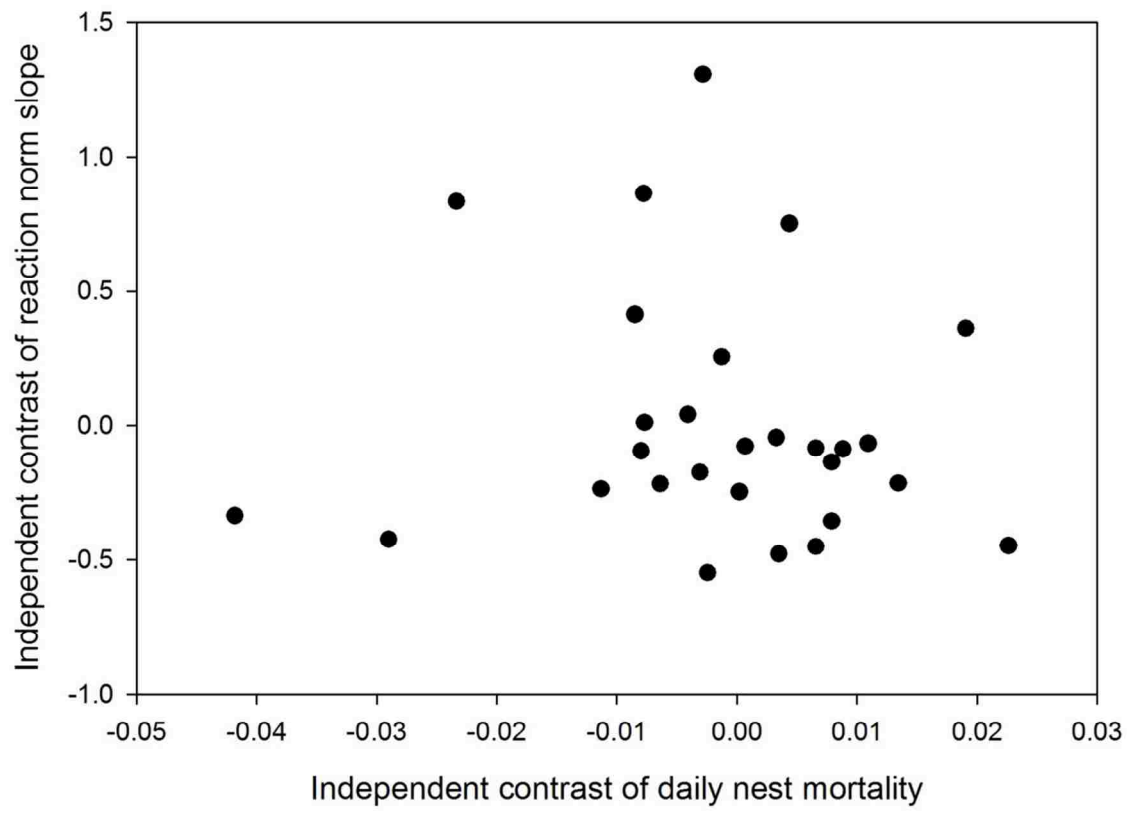
2204 Figure 1. Phylogenetic independent contrasts of daily nest mortality and reaction norm slope  
2205 were not significantly correlated (Pearson correlation,  $r = -0.085$ ,  $P = 0.67$ ).

2206

2207 Figure 2. Phylogenetic independent contrasts of annual adult mortality and reaction norm slope  
2208 were significantly correlated (Pearson correlation,  $r = 0.617$ ,  $P < 0.001$ ). The line shown is the  
2209 regression through the origin.

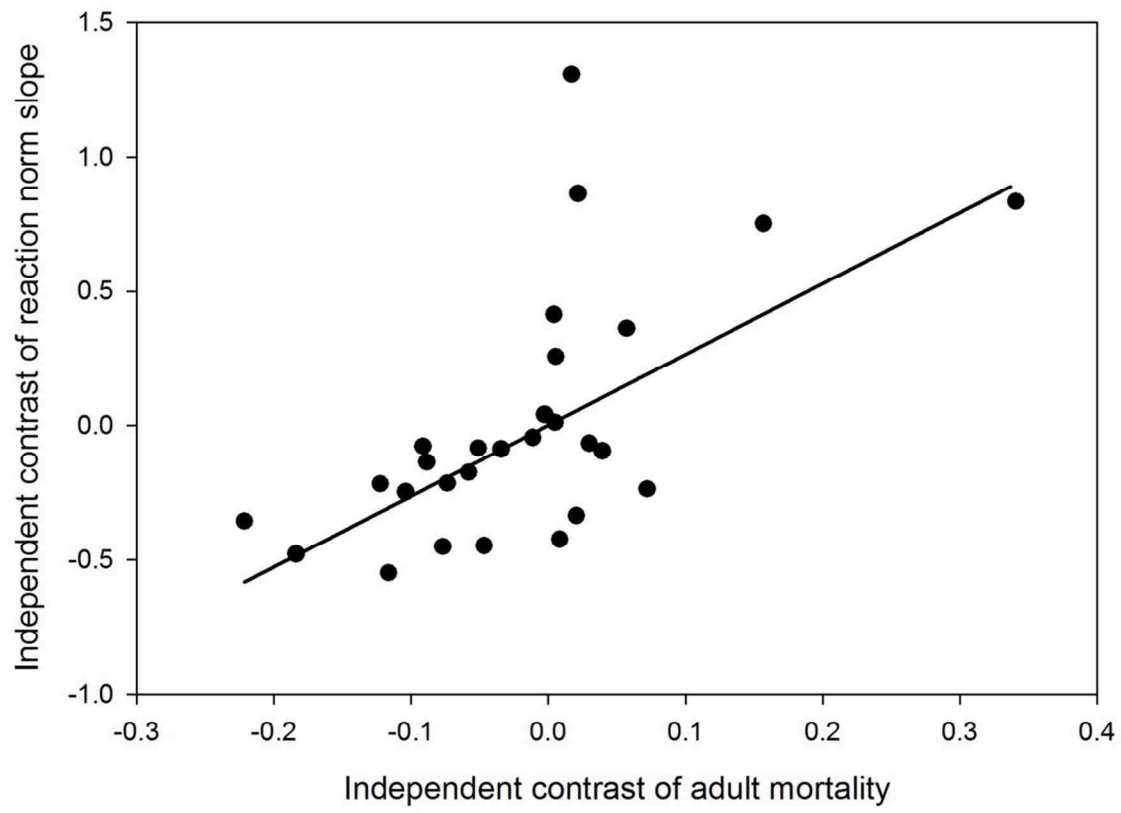
2210

2211 Figure 3. Phylogenetic independent contrasts of annual adult mortality and percentage change in  
2212 provisioning rate in response to brood size reduction were significantly correlated (Pearson  
2213 correlation,  $r = 0.938$ ,  $P < 0.001$ ). The line shown is the regression through the origin.



2214

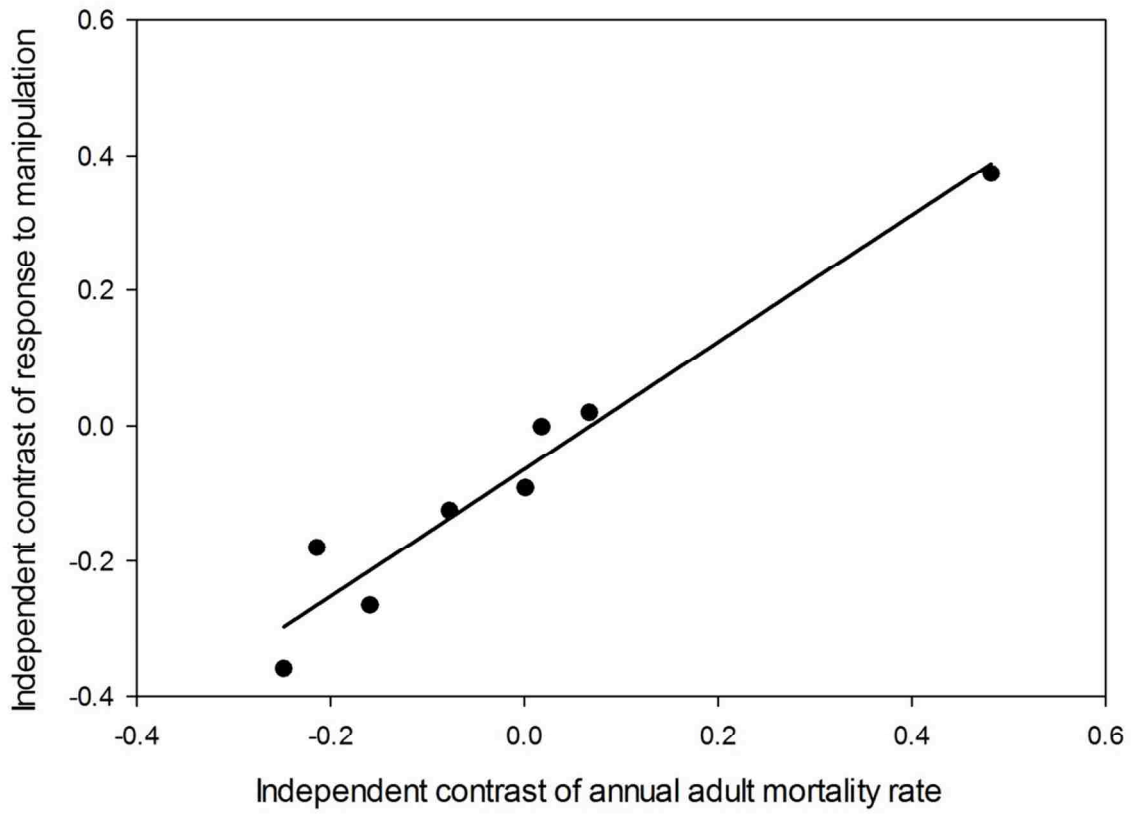
2215 Figure 1.



2216

2217 Figure 2.

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2220 Figure 3.

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